## **Determinants of the Acute-Phase Protein C-Reactive Protein in Myocardial Infarction Survivors: The Role of Comorbidities and Environmental Factors**

Regina Rückerl,<sup>1</sup> Annette Peters,<sup>1,2</sup> Natalie Khuseyinova,<sup>3</sup> Mariarita Andreani,<sup>4</sup> Wolfgang Koenig,<sup>3\*</sup> Christa Meisinger,<sup>1,5</sup> Konstantina Dimakopoulou,<sup>6</sup> Jordi Sunyer,<sup>7,8,9,10</sup> Timo Lanki,<sup>11</sup> Fredrik Nyberg,<sup>12,13</sup> and Alexandra Schneider<sup>1</sup>

**BACKGROUND:** C-reactive protein (CRP), a sensitive marker of the acute-phase response, has been associated with future cardiovascular endpoints independently of other risk factors. A joint analysis of the role of risk factors in predicting mean concentrations and variation of high-sensitivity CRP (hsCRP) in serum has not been carried out previously.

**METHODS:** We used data from 1003 myocardial infarction (MI) survivors who had hsCRP measured monthly up to 8 times and multivariate mixed effects statistical models to study the role of time-variant and -invariant factors on the geometric mean of and the intraindividual variation in hsCRP concentrations.

**RESULTS:** Patients with  $\geq 6.5\%$  glycosylated hemoglobin (HbA1c) had 26.2% higher hsCRP concentrations (95% CI, 7.2%–48.6%) and 20.7% greater variation in hsCRP values ( $P = 0.0034$ ) than patients with lower baseline Hb  $A_{1c}$  values (<6.5%). Similar but less pronounced differences were seen in patients with a selfreported diagnosis of type 2 diabetes. hsCRP concentrations showed less variation in patients who reported angina pectoris, congestive heart failure, or emphysema  $(-11.0\%, -24.9\%, \text{ and } -41.6\%, \text{ respectively}, \text{vs}$ patients without these conditions) but greater variation in males and smokers  $(+24.8\%$  and  $+27.3\%$ , respectively, vs females and nonsmokers). Exposures in the 24 h before blood sampling, including exposure to environmental tobacco smoke, alcohol consumption, and extreme stress, did not have a major impact.

**CONCLUSIONS:** One or 2 hsCRP measurements may not be sufficient to adequately characterize different patient groups after MI with similar precisions. We found hsCRP concentrations to be especially variable in males, smokers, and patients with increased Hb  $A<sub>1c</sub>$ values.

© 2008 American Association for Clinical Chemistry

C-reactive protein  $(CRP)$ ,<sup>14</sup> a sensitive marker of the acute-phase response, has attracted increasing attention in recent years because many epidemiologic studies have shown consistent positive associations between high-sensitivity CRP (hsCRP) concentrations in the peripheral circulation and the risk of future cardiovascular events, independently of established risk factors. Associations have been found with angina pectoris*(1 )* and "hard" coronary and cerebrovascular events in men and women *(2 )*. Koenig et al. *(3 )* reported an almost 3-fold increase in the risk of a first major coronary event for individuals in the highest quintile of the hsCRP distribution in a random sample of initially healthy men from the general population. These findings have led to an ongoing discussion on whether hsCRP should be measured routinely in individuals at risk of cardiovascular disease *(4 )*. The CDC and the American Heart Association recently recommended that hsCRP be measured in individuals at intermediate risk (as defined by the Framingham Risk Score), with the assays to be performed on 2 samples from each

<sup>12</sup> Institute of Environmental Medicine, Karolinska Institute, Stockholm, Sweden: <sup>13</sup> AstraZeneca R&D, Mölndal, Sweden.

<sup>&</sup>lt;sup>1</sup> Helmholtz Zentrum München, German Research Center for Environmental Health, Institute of Epidemiology, Munich, Germany; <sup>2</sup> Helmholtz Zentrum München, German Research Center for Environmental Health, Focus-Network Nanoparticles and Health (NanoHealth), Munich, Germany; <sup>3</sup> Department of Internal Medicine II, Cardiology, University of Ulm Medical Center, Ulm, Germany; <sup>4</sup> Local Health Authority RM E, Department of Epidemiology ASL RME, Rome, Italy; <sup>5</sup> Central Hospital of Augsburg, MONICA/KORA Myocardial Infarction Registry, Augsburg, Germany; <sup>6</sup> Department of Hygiene and Epidemiology, University of Athens Medical School, Athens, Greece; 7 Center for Research in Environmental Epidemiology (CREAL), <sup>8</sup> Municipal Institute of Medical Research (IMIM-Hospital del Mar), <sup>9</sup> CIBER Epidemiologìa y Salud Pública (CIBERESP), and <sup>10</sup> Universitat Pompeu Fabra (UPF), Barcelona, Spain; <sup>11</sup> Environmental Epidemiology Unit, National Public Health Institute (KTL), Kuopio, Finland;

<sup>\*</sup> Address correspondence to this author at: Department of Internal Medicine II—Cardiology, University of Ulm Medical Center, Robert-Koch Str. 8, D-89081 Ulm, Germany. Fax 49-731-500-45021; e-mail wolfgang.koenig@uniklinikulm.de.

Received July 16, 2008; accepted November 6, 2008.

Previously published online at DOI: 10.1373/clinchem.2008.112334

<sup>&</sup>lt;sup>14</sup> Nonstandard abbreviations: CRP, C-reactive protein; hsCRP, high-sensitivity CRP; BMI, body mass index; MI, myocardial infarction; ETS, environmental tobacco smoke; P-spline, penalized splines; Hb  $A_{1c}$ , glycosylated hemoglobin; CHF, congestive heart failure.

person, fasting or nonfasting, taken approximately 2 weeks apart. In the case of an hsCRP measurement  $>$  10 mg/L, indicating an acute inflammatory process, the measurement should be discarded and repeated 2 weeks later *(5 )*. For routine screening, knowledge of basic determinants of hsCRP concentrations is essential. Several determinants have been studied intensively in the past, including nutrition *(6 )*, medication *(2, 7– 10 )*, smoking *(11, 12 )*, body mass index (BMI), and physical activity *(13–15 )*; however, most of these studies relied on only 1 or 2 measurements per patient. Only a few studies have examined factors acutely affecting hsCRP concentrations *(13, 14, 16 )* or the degree of within-patient variation in hsCRP concentration. We used data from a large European study of myocardial infarction (MI) survivors who had hsCRP measured up to 8 times in an attempt to conduct, for the first time, a joint analysis of the role of risk factors in predicting the mean hsCRP concentration and the intraindividual variation in hsCRP. Given that clinical practice may consider preventive measures based on a single hsCRP measurement, this study may contribute important additional information.

## **Materials and Methods**

#### **STUDY POPULATION**

The AIRGENE study, a prospective longitudinal study of post-MI patients, was performed in 6 European cities—Athens (Greece), Augsburg (Germany), Barcelona (Spain), Helsinki (Finland), Rome (Italy), and Stockholm (Sweden). Candidates for the study were identified from population registries of MI patients [Augsburg—Cooperative Health Research in the Augsburg Region (KORA) *(17 )*; Barcelona; Stockholm] or from administrative databases of hospital admissions (Athens, Helsinki, Rome). MI was defined according to the Joint European Society of Cardiology/ American College of Cardiology Committee for the Redefinition of Myocardial Infarction *(18 )*; the study design has been described in detail elsewhere *(19 )*. In brief, the study recruited patients 35–80 years of age who had experienced an MI between 4 months and 6 years before the start of the study. Patients who had undergone interventional procedures  $\leq$ 3 months before the beginning of the study or who had chronic inflammatory diseases were not included. Because AIRGENE initially was a study of the health effects of air pollution, the recruitment of current nonsmokers was preferred, but the inclusion of smokers in some of the centers was unavoidable. All study partners had the study protocol approved by their local human-studies committees, and written informed consent was obtained from all patients. All methods used in the study

centers were conducted according to common standard operating procedures.

#### **CLINICAL MEASUREMENTS**

Patients were invited to participate in 6 to 8 clinical visits at approximately monthly intervals between May 2003 and July 2004. At the first visit, the patient completed a baseline questionnaire regarding comorbidities, regular exercise, smoking history, exposure to environmental tobacco smoke (ETS), socioeconomic status, and alcohol intake. Data recorded regarding medication intake included brand names, doses, and intake pattern. Clinical measurements included blood pressure and BMI, and a serum sample was taken to assess baseline serum lipids, glycosylated hemoglobin (Hb  $A_{1c}$ ) (an indicator of glucose control), and Nterminal pro-B-type natriuretic peptide (a marker of hemodynamic stress).

Each clinical visit was scheduled at the same time of the day and on the same day of the week to minimize the impact of circadian and weekly variation. If patients had acute infections such as a cold or influenza during the 3 days preceding the scheduled visit, examinations were postponed or the blood samples were excluded from analyses.

The patient was asked to recall medication intake for the previous 7 days at each clinical visit and to complete a short questionnaire about time-varying variables in the previous 24 h, such as active and passive smoking, physical activity, perception of extreme stress or anger, consumption of alcohol and black or green tea, and the time of the latest meal before blood draw.

Venous blood samples for preparing EDTAplasma for hsCRP measurement were drawn while the patient was sitting. Samples were cooled and stored at 4 °C for further processing within a maximum of 4 hours. EDTA-containing blood was centrifuged for 20 min at 2500*g* in a centrifuge precooled to 4 °C. Plasma aliquots were shipped on dry ice to the central laboratory in Ulm, Germany, and were stored at  $-80$  °C until analysis. Blood samples were analyzed for hsCRP by latex-enhanced immunonephelometry on a BNII analyzer (Siemens). The interassay CVs for hsCRP were 4.3%, 6.2% and 4.5% at hsCRP concentrations of 1.17 mg/L, 2.38 mg/L, and 13.5 mg/L, respectively.

#### **STATISTICAL ANALYSES**

All statistical analyses were performed with the Statistical Analysis System (SAS) software package (Version 9.1 for Windows; SAS Institute).

We calculated hsCRP CVs as described by Bland and Altman *(20 )* and Fraser and Harris *(21 )*. We used the SAS MIXED procedure to compute estimates of between- and within-individual variances, assuming nested normal random-effects models. These components of variation were then transformed into corresponding CVs, which were calculated as the square root of the respective variance-component estimates divided by the overall mean and then expressed as percentages.

*Determinants of mean hsCRP concentrations.* hsCRP data required log-transformation to fulfill the model assumption of residual normality; therefore, concentration results are given as the geometric mean. To estimate the effect of various determinants on the geometric means of hsCRP concentrations, we used mixed-effects models with random patient effects accounting for repeated measures. Because the half-life of hsCRP is 19 h *(22 )* and therefore much shorter than the intervals between visits, we assumed a compound symmetry structure for the covariance matrix to model the correlation between repeated measures in each patient. Penalized splines (P-splines) in the additive mixedmodel framework allowed for nonparametric exposure–response functions *(23 )*.

We first built a confounder model (base model), which included preselected time-invariant patient characteristics, to permit the assumption of a normally distributed random patient intercept. We tested a wide range of variables known from the literature to have a possible influence on hsCRP, such as city, age, sex, and BMI. Linear variables were added linearly to the model. The decision on whether a specific factor remained in the model was based on the goodness-of-fit according to Akaike's information criterion.

In a second step, additional time-invariant variables not initially considered for the base model (such as reported diseases, regular medication intake, and smoking history) as well as time-varying variables, such as physical activity or alcohol consumption in the 24 h before the blood draw, were added to the base model, always one at a time. To avoid overcontrol, we removed pack-years of smoking from the base model when we analyzed smoking status, and we removed Hb  $A_{1c}$  for the analysis of diabetes. Variables that described a time difference, such as the time of the last meal before the blood sampling, were categorized into 4 intervals of 6 h each: 0-5 h, 6-11 h, 12-17 h, and 18-23 h before sampling. Results are given as the percent change in the geometric mean of the hsCRP concentration.

*Determinants of hsCRP variation.* To calculate differences in variation, we used the MIXED procedure in SAS with the "repeated/group=" statement, which calculates the within-patient variation, and a "random/  $group = " statement, which allows for different inter$ cepts in the defined groups, representing the betweenpatient variation. A likelihood-ratio test was used to determine if the differences between the groups were statistically significant. Linear variables were categorized beforehand, usually with interquartile ranges. Results are given as variance estimates of log-transformed hsCRP concentrations, with between-individual and within-individual results presented separately (Fig. 1), and as the relative difference (in percent) in withinindividual variation compared with the reference group (see tables).

To account for the large number of statistical tests, we corrected the significance level of the P value to 0.00125, which equals a Bonferroni correction for 40 variables.

*Sensitivity analyses.* We conducted sensitivity analyses for comorbidities that might be associated with the intake of certain medications and used a  $\chi^2$  test to evaluate possible associations between comorbidities and medication intake. If we found an association ( $P \leq$ 0.05), we adjusted the multivariable model for the respective medication to investigate whether the comorbidity effect was altered by including medication in the model. Moreover, we calculated a model that included most of the presented variables to identify those variables that led to the greatest increase in variation.

## **Results**

## **STUDY POPULATION**

In total, 1003 patients with at least 2 valid blood samples participated in the study. Of the 6068 collected samples, 255 had to be excluded because of acute infections or surgical procedures that occurred shortly before the clinic visit. Overall, 5813 plasma samples remained for analysis (see Table 1 in the Data Supplement that accompanies the online version of this article at http://www.clinchem.org/content/vol55/issue2).

Table 2 in the online Data Supplement summarizes the patient characteristics by center, and Table 3 in the online Data Supplement presents the patient characteristics according to sex. Mean hsCRP concentrations were highest in Barcelona and lowest in Helsinki; however, hsCRP concentrations were not exceptionally high on average. In 75 samples, the hsCRP concentration was lower than 0.16 mg/L, and these values were set at 0.16 mg/L. More details are given elsewhere *(19 )*. The CV was 107% of the overall mean within individuals and 139% between individuals.

#### **ASSOCIATION BETWEEN TIME-INVARIANT VARIABLES AND hsCRP**

*Base model.* Table 1 shows the associations of patient characteristics with the geometric mean of the hsCRP concentration and its variation, as estimated jointly from the base model. Male participants had signifi-



**Fig. 1. Variation in mean hsCRP concentration between (upper panel) and within (lower panel) individuals for separate hsCRP measurements made over time.**

Variance component values are presented according to patient characteristics (sex, CHF diagnosis, HDL cholesterol concentration, body weight, and chronic bronchitis). Error bars represent 95% confidence limits.



# **Table 1. Association of time-invariant variables with the geometric mean of and the variation in hsCRP**

<sup>a</sup> GM, geometric mean; Ref, reference; NT-proBNP, N-terminal pro-B-type natriuretic peptide.

<sup>b</sup> Statistically significant after adjusting for multiple testing ( $\alpha = 0.00125$ ).

<sup>c</sup> Measured at baseline.

<sup>d</sup> BMI classification according to the WHO (2000).

<sup>e</sup> Base model including the linear variable.

f Categories correspond to interquartile ranges.



cantly lower hsCRP concentrations than female participants but had greater variation over time in log hs-CRP concentration (Fig. 1). This difference was less pronounced after controlling for the intake of hormone-replacement medications in women. We found a U-shaped relationship for age with the lowest hsCRP concentration in the age group of 50 –59 years (Fig. 2), whereas most other associations were linear. A separate analysis showed that this effect was mainly driven by the results for men; women had a positive linear association between hsCRP concentration and age (data not shown). In contrast, hsCRP variation was greatest in the oldest patient group. Overweight and obese patients *(24 )* had higher hsCRP concentrations than participants with normal weights, but the concentrations in these patients were less variable (Fig. 1). Hb  $A<sub>1c</sub>$  concentrations  $>6.5\%$  were positively associated with the geometric mean of and the variation in hsCRP concentration, whereas a diagnosis of type 2 diabetes was positively associated with the variation but not with the geometric mean (Table 2). hsCRP concentrations were also positively associated with higher concentrations of N-terminal pro–B-type natriuretic peptide and total cholesterol.

*Additional time-invariant variables and hsCRP.* Tables 2 and 3 summarize the associations of hsCRP concentration with disease history, lifestyle, and medication intake. A family history of MI was associated with slightly higher hsCRP concentrations. On the other

hand, hsCRP concentrations showed less variation in patients who reported angina pectoris, congestive heart failure (CHF), emphysema, or a family history of MI (Fig. 1), and these results remained statistically significant after adjusting for multiple testing. Time since last MI did not show any association with the geometric mean of or the variation in hsCRP concentration (Table 2).

Habitual physical activity did not influence hs-CRP concentrations; however, the variation in hsCRP concentration seemed to be higher in inactive people and lower in those who were partially active, compared with regularly active study participants. HDL cholesterol was inversely related to the geometric mean of the hsCRP concentration; greater variation in hsCRP concentration was noted in patients with increased HDL cholesterol concentrations (Table 3).

Patients reporting the intake of statins or other lipid-lowering drugs had lower hsCRP concentrations and less variation. On the other hand, patients taking angiotensin-converting enzyme inhibitors had greater variation in hsCRP concentrations, whereas the geometric mean was negatively associated with medication intake (Table 3). Use of acetylsalicylic acid or  $Ca^{2+}$ channel blockers did not affect the geometric mean of or the variation in hsCRP concentration.

Table 4 summarizes the results for different smoking-related variables. Twenty-five pack-years of smoking produced an increase of approximately 16% in the



**Table 2. Association of disease history with the geometric mean of and the variation in hsCRP concentration,**

geometric mean of the hsCRP concentration, and inclusion of smoking status in the model had little effect on this result. Including pack-years of smoking, however, removes the borderline effect for ex-smokers that we found in the model that does not include pack-years of smoking. Examination of the effects of smoking and ETS exposure revealed a heterogeneous picture. Current regular smokers and nonsmokers who reported regular ETS exposure had higher hsCRP concentrations, whereas occasional smokers seemed to have lower hsCRP



<sup>o</sup> Statistically significant<br><sup>c</sup> Measured at baseline.



## **Table 4. Association of smoking and ETS exposure with the geometric mean of and the variation in hsCRP concentration, adjusted for the variables of the base model.**

concentrations than nonsmokers not regularly exposed to cigarette smoke. The results were not statistically significant, however, especiallywhen pack-years of smokingwas included in the model. The numbers of participants were low in several of the groups (Table 4).

**ASSOCIATION BETWEEN TIME-VARYING VARIABLES AND hsCRP** Time-varying variables had either no or a small influence on hsCRP concentration (Fig. 3A). Recent alcohol consumption and extreme stress or anger were associated with lower geometric-mean hsCRP



limits.

concentrations, but the results were not statistically significant. Whereas physical activity over the previous 24 h showed no association with hsCRP concentration (Fig. 3A), physical activity between 6 and 11 hours before blood draw was associated with increased hsCRP concentrations (Fig. 3B). For the other time-varying variables, no such time-specific effects were seen.

## **SENSITIVITY ANALYSES**

Additional adjustment for medication did not change the results for comorbidities (data not shown).

With very few exceptions, the results for the variation model that included all variables at the same time did not differ much from those described in the presented tables. The results revealed the largest increases in hsCRP variation for patients with HDL cholesterol  $concentrations  $> 0.91 \text{ mmol/L}$ , an older age (especially$ 70 years), male sex, a log B-type natriuretic peptide concentration of  $\geq$ 5.98 (ng/L), and intake of antiinflammatory medication (data not shown).

## **Discussion**

We investigated repeated measurements of hsCRP in a population of MI survivors and found that the variation in hsCRP concentration within patients over time was only slightly less than the variation between patients. Moreover, our data revealed that certain subgroups had higher geometric-mean hsCRP concentrations and/or greater variation in the hsCRP concentration, but higher geometric means and greater variation did not necessarily occur together. Obese and overweight patients and certain age groups had higher hsCRP concentrations but less variation in concentration. We also found that patients who reported angina pectoris, emphysema, or CHF had less variation in hsCRP concentration, whereas the geometric-mean concentration did not seem to be affected. On the other hand, for patients with impaired glucose control, as indicated by increased baseline Hb  $A_{1c}$  concentrations  $(\geq 6.5\%)$ , we found a higher hsCRP concentration and greater hsCRP variation. We saw similar but less pronounced differences for the diagnosis of type 2 diabetes. Short-term exposures in the 24 h preceding blood draw, such as ETS exposure, alcohol consumption, or extreme stress or anger, did not have a major impact on hsCRP concentration. This study examined MI patients only, and therefore the results may not be entirely generalizable to a population without cardiovascular disease.

## **PATIENT CHARACTERISTICS THAT AFFECT hsCRP CONCENTRATION AND ITS VARIATION**

A variety of studies have examined determinants of hsCRP concentrations. Although some investigators did not report any sex differences *(25 )*, others found lower concentrations in men *(6, 7, 11, 26 )*, in line with our results. Hutchinson et al. *(26 )* hypothesized that the sex difference might be due to estrogen intake in women, and a study of diabetic women has shown significantly higher hsCRP concentrations in patients receiving hormone-replacement therapy *(7 )*. Our data revealed that intake of hormone-replacement medications had a slightly positive but nonsignificant association with hsCRP concentration (Table 3), a result that is consistent with this hypothesis.

As for the influence of age on hsCRP concentrations, some authors have found a positive linear relationship *(26 )*, but a lack of an association has also been reported *(12 )*. As far as we know, a U-shaped function, as seen in our data, has not previously been reported. This observation could be due to the way the relationships were modeled and/or to the fact that our data were based on MI survivors, whereas most studies have been conducted with participants from the general population.

Consistent with our results, others have reported positive associations of hsCRP concentration with increased BMI and obesity *(12, 15, 25 )*, for smokers compared with nonsmokers *(11, 25 )*, and for individuals with low HDL cholesterol concentrations*(12, 25 )*. Several studies have shown that statin therapy *(9, 10 )* and treatment with angiotensin-converting enzyme inhibitors*(27 )*reduce circulating hsCRP concentrations, results that are in line with our findings. Moreover, hsCRP–lowering effects have also been seen with acetylsalicylic acid *(8 )*. Although our findings were consistent with a small reduction in hsCRP concentration due to acetylsalicylic acid, these associations were not statistically significant.

To our knowledge, none of the previously published studies examined variation in hsCRP concentration over time among different subgroups or with respect to possible determinants. Interestingly, we found that an increase of and greater variation in hsCRP concentration were not necessarily related. Individuals who reported angina pectoris, CHF, or emphysema had less variation in hsCRP concentration compared with participants who did not report any of these disorders. These findings remained stable for CHF and emphysema, even after adjustment for multiple testing and associated medication intake. Emphysema is often caused by smoking *(28 )*, and our data showed that 80% of the emphysema patients were past or current smokers. Because emphysema and early-stage CHF do not necessarily include an inflammatory component, it is also conceivable that the lower variation in hsCRP concentration in these patients is merely a marker for a different mechanism, such as an underlying genetic component. Studies of twins have demonstrated a substantial genetic contribution to baseline hsCRP concentrations*(29 )*, and genetic analyses of the AIRGENE data set revealed that minor alleles of several variants of selected candidate genes were significantly associated with intraindividual variation in hsCRP concentration *(30 )*.

Whether different factors affect each other and, if so, how they do remain speculative. It is possible that a

combination of variables amplifies hsCRP variation, although it is also conceivable that certain combinations of factors can reduce such variation. Additionally, factors that are associated with high variation could just be indicators for a different mechanism. For example, the increase in variation associated with medication intake seen in our data might be a direct effect of the medication itself; however, it is more likely that the high variation is due to the underlying disease that led to the prescription of the drug.

#### **RESPONSE TO ENVIRONMENTAL FACTORS**

It is still unclear why some patients develop cardiovascular disease or experience an MI due to certain triggers, whereas others do not. Heavy physical exertion *(31 )* and extreme anger *(32 )* have been reported as causes for an acute MI. In addition, environmental stimuli such as tobacco smoke *(33 )* and air pollution *(34, 35 )* are associated with an increased risk for adverse cardiovascular events. It is conceivable that individuals with special characteristics react in a more pronounced way to environmental factors than others. A generally higher concentration of inflammation markers, and/or greater variation in inflammation might offer one possible explanation.

We found that patients with increased Hb  $A_{1c}$  concentrations and patients with self-reported type 2 diabetes have greater variation in hsCRP concentration, even in this relatively homogeneous population of MI survivors. It is plausible, but quite speculative, that these subgroups also had a stronger reaction (e.g., a more pronounced inflammatory response) to environmental factors. Studies of diabetic patients *(7 )* have shown considerably higher mean hsCRP concentrations than our population, which consisted of only about 20% diabetic individuals. Persistently increased hsCRP concentrations as well as acute changes in concentrations of inflammation markers have been associated in cohort studies with an increased risk of cardiovascular events *(2, 3 )*. This observation might represent a possible link for the reported associations of air pollution and passive smoking with adverse cardiovascular outcomes, because particle-induced systemic inflammation is one of the hypothesized pathways*(33, 36 )*. Individuals with certain diseases, such as diabetes and MI, have been demonstrated to have an enhanced susceptibility for air pollution–related conditions, possibly due to a disease-induced increased inflammatory burden *(37 )*. We did not see higher hsCRP concentrations in diabetic patients, but we did observe greater variation in hsCRP concentration compared with nondiabetic patients. Furthermore, patients with increased Hb  $A_{1c}$  concentrations ( $\geq 6.5\%$ ) had higher hsCRP concentrations and greater hsCRP variation. High Hb  $A_{1c}$  concentrations seem to reflect uncontrolled rather than undiagnosed diabetes, because 89% of the participants with Hb  $A<sub>1c</sub>$  values  $> 6.5\%$  reported a diagnosis of diabetes. On the other hand, only half of the AIRGENE population with diagnosed diabetes met the Hb  $A_{1c}$  criterion of  $\geq 6.5\%$ . This finding might indicate that metabolically stable diabetic patients are at less risk compared with patients with unstable diabetes.

Our study is in line with others *(10 )* in showing a clear negative association between statin intake and hsCRP concentration. We hypothesize that the intake of statins attenuates the impact of environmental variables, and therefore statin therapy in addition to following recommended guidelines might be beneficial in certain particularly susceptible subgroups to avoid adverse cardiovascular effects of environmental stimuli. More research in this area is clearly needed, however.

## **SHORT-TERM INFLUENCES ON hsCRP**

Several studies have demonstrated that regular moderate to vigorous exercise leads to a decrease in hsCRP concentrations, although the results are conflicting and some authors have attributed the detected negative association to a lower BMI in the individuals who exercise rather than to a direct effect of physical activity on inflammation markers *(38 )*. Short-term effects, however, have been studied only in individuals whose activities must be considered extreme, even for professional athletes *(16 )*. Although our population of MI survivors were expected to perform in only light sporting activities, we found a transient increase in hsCRP concentration 6 to 11 hours after physical activity that quickly returned to baseline concentrations. A study of postmenopausal women did not observe any increase in hsCRP concentration at 1 h or 24 h after exercise, compared with baseline concentrations *(13 )*. In addition, hsCRP concentrations measured immediately and 48 h after a 7-km hill race did not differ from baseline concentrations *(14 )*. These conflicting results might be explained by different time frames and differences in exercise intensities. A study of the time course of hsCRP concentration after surgical procedures showed a rapid increase starting 6 to 8 hours after the operation, with the highest peak at about 48 h and the concentration returning to baseline between 72 h and 144 h after the surgical intervention *(39 )*.

Regarding tea and alcohol intake, no publication has addressed the effects on hsCRP within 24 h after consumption. We found a slight decrease in hsCRP in association with tea and alcohol intake; however, whether this result reflects regular consumption or an immediate reaction is difficult to determine. A decrease in hsCRP concentration after regular consumption of black tea *(40 )* and moderate amounts of alcohol *(6 )* has been shown.

## **Conclusion**

This study is the first to measure within-patient variation in hsCRP concentration in a large study population. We confirmed and extended published results on the association of patient characteristics and intake of medications with hsCRP concentrations in male and female MI survivors. Short-term influences, however, did not seem to impact hsCRP concentrations. Males, elderly individuals, smokers, and patients with increased Hb  $A_{1c}$  concentrations had greater intraindividual variation in repeated measurements of hsCRP. In patients with manifest cardiovascular disease, in particular after MI, several hsCRP measurements may be necessary to adequately characterize their risk, especially in defined subgroups. Whether this variation also makes these patients more susceptible to adverse environmental variables needs further investigation.

**Author Contributions:***All authors confirmed they have contributed to the intellectual content of this paper and have met the following 3 requirements: (a) significant contributions to the conception and design,* *acquisition of data, or analysis and interpretation of data; (b) drafting or revising the article for intellectual content; and (c) final approval of the published article.*

**Authors' Disclosures of Potential Conflicts of Interest:** *Upon manuscript submission, all authors completed the Disclosures of Potential Conflict of Interest form. Potential conflicts of interest:*

Employment or Leadership: F. Nyberg, AstraZeneca R&D, Mölndal, Sweden.

**Consultant or Advisory Role:** None declared.

**Stock Ownership:** F. Nyberg, AstraZeneca.

**Honoraria:** None declared.

**Research Funding:** The AIRGENE study was funded as part of the European Union's 5th Framework Program, key action number 4: "Environment and Health," contract number QLRT-2002-02236. **Expert Testimony:** None declared.

**Role of Sponsor:** The funding organizations played no role in the design of study, choice of enrolled patients, review and interpretation of data, or preparation or approval of manuscript.

**Acknowledgments:** The authors are indebted to the AIRGENE study group (http://www.gsf.de/epi/de/index\_ag\_epi\_luftsch.htm). We also thank Gerlinde Trischler, University of Ulm Medical Center, for excellent technical assistance and Prof. Peter Kern, University of Ulm Medical Center, for providing access to a BN II analyzer. Finally, we express our appreciation to all study participants.

#### **References**

**1.** Haverkate F, Thompson SG, Pyke SDM, Gallimore JR, Pepys MB. Production of C-reactive protein and risk of coronary events in stable and unstable angina. Lancet 1997;349:462–6.

**2.** Ridker PM, Cushman M, Stampfer MJ, Tracy RP, Hennekens CH. Inflammation, aspirin, and the risk of cardiovascular disease in apparently healthy men. N Engl J Med 1997;336:973–9.

- **3.** Koenig W, Sund M, Frohlich M, Fischer HG, Lowel H, Doring A, et al. C-reactive protein, a sensitive marker of inflammation, predicts future risk of coronary heart disease in initially healthy middleaged men: results from the MONICA (Monitoring Trends and Determinants in Cardiovascular Disease) Augsburg Cohort Study, 1984 to 1992. Circulation 1999;99:237–42.
- **4.** Yeh ET, Willerson JT. Coming of age of C-reactive protein: using inflammation markers in cardiology. Circulation 2003;107:370–1.
- **5.** Pearson TA, Mensah GA, Alexander RW, Anderson JL, Cannon RO, Criqui M, et al. Markers of inflammation and cardiovascular disease application to clinical and public health practice—a statement for healthcare professionals from the Centers for Disease Control and Prevention and the American Heart Association. Circulation 2003;107:499–511.
- **6.** Imhof A, Woodward M, Doering A, Helbecque N, Loewel H, Amouyel P, et al. Overall alcohol intake, beer, wine, and systemic markers of inflammation in western Europe: results from three MONICA samples (Augsburg, Glasgow, Lille). Eur Heart J 2004;25:2092–100.
- **7.** Bowden DW, Lohman K, Hsu FC, Langefeld CD, Carr JJ, Lenchik L, et al. Hormone replacement therapy is associated with increased C-reactive protein in women with type 2 diabetes in the Diabetes Heart Study. Diabet Med 2006;23:

763–7.

- **8.** Ikonomidis I, Andreotti F, Economou E, Stefanadis C, Toutouzas P, Nihoyannopoulos P. Increased proinflammatory cytokines in patients with chronic stable angina and their reduction by aspirin. Circulation 1999;100:793–8.
- **9.** Rosenson RS, Tangney CC, Schaefer EJ. Comparative study of HMG-CoA reductase inhibitors on fibrinogen. Atherosclerosis 2001;155:463–6.
- **10.** Ridker PM, Rifai N, Lowenthal SP. Rapid reduction in C-reactive protein with cerivastatin among 785 patients with primary hypercholesterolemia. Circulation 2001;103:1191–3.
- **11.** Frohlich M, Sund M, Lowel H, Imhof A, Hoffmeister A, Koenig W. Independent association of various smoking characteristics with markers of systemic inflammation in men. Results from a representative sample of the general population (MONICA Augsburg Survey 1994/95). Eur Heart J 2003;24:1365–72.
- **12.** Greenfield JR, Samaras K, Jenkins AB, Kelly PJ, Spector TD, Gallimore JR, et al. Obesity is an important determinant of baseline serum C-reactive protein concentration in monozygotic twins, independent of genetic influences. Circulation 2004;109:3022–8.
- **13.** Davis J, Murphy M, Trinick T, Duly E, Nevill A, Davison G. Acute effects of walking on inflammatory and cardiovascular risk in sedentary postmenopausal women. J Sports Sci 2007;26:303–9.
- **14.** Simpson RJ, Wilson MR, Black JR, Ross JA, Whyte GP, Guy K, Florida-James GD. Immune alterations, lipid peroxidation, and muscle damage following a hill race. Can J Appl Physiol 2005;30:196–211.
- **15.** Thorand B, Baumert J, Doring A, Herder C, Kolb H, Rathmann W, et al. Sex differences in the relation of body composition to markers of inflammation. Atherosclerosis 2006;184:216–24.
- **16.** Margeli A, Skenderi K, Tsironi M, Hantzi E, Matalas AL, Vrettou C, et al. Dramatic elevations of interleukin-6 and acute-phase reactants in athletes participating in the ultradistance foot race Spartathlon: severe systemic inflammation and lipid and lipoprotein changes in protracted exercise. J Clin Endocrinol Metab 2005;90:3914–8.
- 17. Löwel H, Meisinger C, Heier M, Hormann A. The population-based acute myocardial infarction (AMI) registry of the MONICA/KORA study region of Augsburg. Gesundheitswesen 2005;67 Suppl 1:S31–7.
- **18.** Joint European Society of Cardiology/American College of Cardiology Committee for the Redefinition of Myocardial Infarction. Myocardial infarction redefined: a consensus document of The Joint European Society of Cardiology/American College of Cardiology Committee for the Redefinition of Myocardial Infarction. Eur Heart J 2000; 21:1502–13.
- **19.** Peters A, Schneider A, Greven S, Bellander T, Forastiere F, Ibald-Mulli A, et al. Air pollution and inflammatory response in myocardial infarction survivors: gene-environment interactions in a high-risk group. Inhal Toxicol 2007;19 Suppl 1:161–75.
- **20.** Bland JM, Altman DG. Measuring agreement in method comparison studies. Stat Methods Med Res 1999;8:135–60.
- **21.** Fraser CG, Harris EK. Generation and application of data on biological variation in clinical chemistry. Crit Rev Clin Lab Sci 1989;27:409–37.
- 22. Koenig W, Sund M, Frohlich M, Löwel H, Hutchinson WL, Pepys MB. Refinement of the association of serum C-reactive protein concentration and coronary heart disease risk by correction for within-subject variation over time. Am J Epidemiol 2003;158:357–64.
- 23. Greven S, Küchenhoff H, Peters A. Additive mixed models with P-splines. In: Hinde J, Einbeck J, Newell J, eds. Proceedings of the 21st International Workshop on Statistical Modelling; 2006 Jul 3–7; Galway, Ireland. Amsterdam: Statistical Modelling Society; 2006. p 201–7.
- **24.** World Health Organization. Obesity: preventing and managing the global epidemic. Report of a WHO consultation. World Health Organ Tech Rep Ser 2000;894:i–253.
- 25. García-Lorda P, Bulló M, Balanzà R, Salas-Salvadó J. C-reactive protein, adiposity and cardiovascular risk factors in a Mediterranean population. Int J Obes (Lond) 2006;30:468–74.
- **26.** Hutchinson WL, Koenig W, Frohlich M, Sund M, Lowe GD, Pepys MB. Immunoradiometric assay of circulating C-reactive protein: age-related values in the adult general population. Clin Chem 2000; 46:934–8.
- 27. Soriano S, González L, Martín-Malo A, Rodríguez M, Aljama P. C-reactive protein and low albumin are predictors of morbidity and cardiovascular events in chronic kidney disease (CKD) 3–5 patients. Clin Nephrol 2007;67:352–7.
- **28.** Petty TL. COPD in perspective. Chest 2002;121: 116S–20S.
- **29.** Pepys MB, Hirschfield GM. C-reactive protein: a critical update. J Clin Invest 2003;111:1805–12.
- **30.** Kolz M, Koenig W, Muller M, Andreani M, Greven S, Illig T, et al. DNA variants, plasma levels and variability of C-reactive protein in myocardial infarction survivors: results from the AIRGENE study. Eur Heart J 2007;29:1250–8.
- **31.** Albert CM, Mittleman MA, Chae CU, Lee IM, Hennekens CH, Manson JE. Triggering of sudden death from cardiac causes by vigorous exertion. N Engl J Med 2000;343:1355–61.
- **32.** Mittleman MA, Maclure M, Sherwood JB, Mulry RP, Tofler GH, Jacobs SC, et al., for the Determinants of Myocardial Infarction Onset Study Investigators. Triggering of acute myocardial infarction onset by episodes of anger. Circulation 1995;92: 1720–5.
- **33.** Barnoya J, Glantz SA. Cardiovascular effects of secondhand smoke: nearly as large as smoking. Circulation 2005;111:2684–98.
- **34.** Lanki T, Pekkanen J, Aalto P, Elosua R, Berglind N, D'Ippoliti D, et al. Associations of traffic related air pollutants with hospitalisation for first acute myocardial infarction: the HEAPSS study. Occup Environ Med 2006;63:844–51.
- **35.** Peters A, Dockery DW, Muller JE, Mittleman MA.

Increased particulate air pollution and the triggering of myocardial infarction. Circulation 2001; 103:2810–5.

- 36. Peters A, Döring A, Wichmann HE, Koenig W. Increased plasma viscosity during air pollution episode: a link to mortality? Lancet 1997;349: 1582–7.
- **37.** Bateson TF, Schwartz J. Who is sensitive to the effects of particulate air pollution on mortality? A case-crossover analysis of effect modifiers. Epidemiology 2004;15:143–9.
- **38.** Elosua R, Bartali B, Ordovas JM, Corsi AM, Lauretani F, Ferrucci L. Association between physical activity, physical performance, and inflammatory biomarkers in an elderly population: the InCHI-ANTI study. J Gerontol A Biol Sci Med Sci 2005; 60:760–7.
- **39.** Colley CM, Fleck A, Goode AW, Muller BR, Myers MA. Early time course of the acute phase protein response in man. J Clin Pathol 1983;36:203–7.
- **40.** De Bacquer D, Clays E, Delanghe J, De Backer G. Epidemiological evidence for an association between habitual tea consumption and markers of chronic inflammation. Atherosclerosis 2006;189: 428–35.