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Abstract *Aims:*
 The onset of clinical type 1 diabetes (T1D) is preceded by the occurrence of disease-specific autoantibodies. The level of autoantibody titers is known to be associated with progression time from the first emergence of autoantibodies to the onset of clinical symptoms, but detailed analyses of this complex relationship are lacking. We aimed to fill this gap by applying advanced statistical models.
Methods:
 We investigated data of 613 children from the prospective TEDDY study who were persistent positive for IAA, GADA and/or IA2A autoantibodies. We used a novel approach of Bayesian joint modeling of longitudinal and survival data to assess the potentially time- and covariate-dependent association between the longitudinal autoantibody titers and progression time to T1D.
Results:
 For all autoantibodies we observed a positive association between the titers and the T1D progression risk. This association was estimated as time-constant for IA2A, but decreased over time for IAA and GADA. For example the hazard ratio [95% credibility interval] for IAA (per transformed unit) was 3.38 [2.66, 4.38] at 6 months after seroconversion, and 2.02 [1.55, 2.68] at 36 months after seroconversion.
Conclusions:
 These findings indicate that T1D progression risk stratification based on autoantibody titers should focus on time points early after seroconversion. Joint modeling techniques allow for new insights into these associations.

Keywords (separated by '-') Autoantibodies - Joint modeling - Type 1 diabetes

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 Meike Köhler and Andreas Beyerlein: Shared first authorship.
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2 Joint modeling of longitudinal autoantibody patterns 3 and progression to type 1 diabetes: results from the TEDDY study

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9 Abstract

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14 bodies to the onset of clinical symptoms, but detailed anal-
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18 IAA, GADA and/or IA2A autoantibodies. We used a novel
19 approach of Bayesian joint modeling of longitudinal and
20 survival data to assess the potentially time- and covariate-
21 dependent association between the longitudinal autoanti-
22 body titers and progression time to T1D.
23

Results For all autoantibodies we observed a positive
24 association between the titers and the T1D progression
25 risk. This association was estimated as time-constant for
26 IA2A, but decreased over time for IAA and GADA. For
27 example the hazard ratio [95% credibility interval] for IAA
28 (per transformed unit) was 3.38 [2.66, 4.38] at 6 months
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32 **Conclusions** These findings indicate that T1D progres-
 33 sion risk stratification based on autoantibody titers should
 34 focus on time points early after seroconversion. Joint
 35 modeling techniques allow for new insights into these
 36 associations.

37 **Keywords** Autoantibodies · Joint modeling · Type 1
 38 diabetes

39 Introduction

40 Type 1 diabetes (T1D) is one of the most common chronic
 41 diseases in childhood, with worldwide increasing incidence
 42 [1]. The disease is preceded by a preclinical period of islet
 43 autoimmunity, which most commonly develops in early
 44 infancy [2, 3]. The presence of islet autoantibodies is asso-
 45 ciated with the progression to clinical diabetes [4]. How-
 46 ever, the time from the first emergence of autoantibodies,
 47 called seroconversion, to the onset of clinical symptoms
 48 varies considerably between individuals, ranging from
 49 weeks to decades [4].

50 It is also known that the combination of different autoan-
 51 tibodies as well as the autoantibody titer is associated with
 52 progression time [5]. For insulin autoantibodies (IAA),
 53 both their titers around seroconversion and their mean lev-
 54 els over time have been found to be associated with pro-
 55 gression to T1D [2, 6], and similar findings have been
 56 recently reported for other islet autoantibodies [7–9]. Nev-
 57 ertheless, detailed analyses of autoantibody titers over time
 58 are lacking.

59 Here, we investigated data of more than 600 islet-
 60 autoantibody-positive children followed up within the pro-
 61 spective The Environmental Determinants of Diabetes in
 62 the Young (TEDDY) study [10, 11]. In contrast to previous
 63 analyses, we used joint models of longitudinal and survival
 64 data. This class of models has the advantage to avoid poten-
 65 tial bias due to characteristics of the longitudinal markers
 66 (here autoantibodies), such as random biological fluctua-
 67 tions, informative censoring and discrete measurement time
 68 points [12]. By applying a novel approach of joint mod-
 69 eling, we gained further insights into the potentially com-
 70 plex relationship between longitudinal islet-autoantibody
 71 measures and the time to T1D progression, particularly
 72 with respect to time-varying associations of both.

73 Methods

74 TEDDY is an ongoing prospective cohort study funded by
 75 the National Institutes of Health with the primary goal to
 76 identify environmental causes of T1D. The TEDDY study
 77 enrolled 8676 children with increased genetic risk for T1D

78 who were recruited in six clinical research centers located
 79 in the USA, Finland, Germany and Sweden between 2004
 80 and 2010 shortly after birth. Detailed information on
 81 study design, eligibility and methods has been previously
 82 published [11, 13, 14]. Written informed consents were
 83 obtained for all participants from a parent or primary care-
 84 taker, separately, for genetic screening and for participa-
 85 tion in prospective follow-up. The study was approved by
 86 local Institutional Review Boards and is monitored by the
 87 External Advisory Board formed by the National Institutes
 88 of Health. For this analysis, we used the data of all children
 89 who had developed one or more persistent islet autoanti-
 90 bodies by the time of our data access (December 31, 2014).
 91 At that time point the median age of the children analyzed
 92 at their last visit was 6.5 years with a range from 0.75 to
 93 10.2 years.

94 Definition of islet autoimmunity

95 Development of persistent islet autoimmunity was assessed
 96 every 3 months and defined by the presence of at least one
 97 islet autoantibody among autoantibodies to insulin (IAA),
 98 glutamic acid decarboxylase (GADA) and insulinoma-
 99 associated protein 2 (IA2A) on two or more consecutive
 100 visits confirmed by two laboratories. Date of persistent
 101 autoimmunity to an autoantibody was defined as the draw
 102 date of the first sample of the two consecutive samples
 103 which deemed the child persistent confirmed positive for
 104 this autoantibody. As described in more detail elsewhere
 105 [7], the respective autoantibody titers were standardized
 106 to be comparable across study laboratories (University of
 107 Bristol, UK; and University of Colorado, Denver, USA) by
 108 subtracting the laboratory- and antibody-specific threshold
 109 and dividing by the laboratory- and autoantibody-specific
 110 standard deviation and were log-transformed afterward.

111 Study outcome

112 The main outcome of this analysis was the time to devel-
 113 opment of T1D after seroconversion in months. T1D diag-
 114 nosis was based on American Diabetes Association criteria
 115 [15].

116 Statistical analyses

117 Of the 8676 children enrolled, 613 had developed one or
 118 more autoantibodies at the time of our data access. We
 119 created three subsets of the data where we restricted the
 120 data to children who had seroconverted to IAA ($n = 442$),
 121 GADA ($n = 466$) or IA2A ($n = 288$), respectively. These
 122 subsets were not mutually exclusive, as children had poten-
 123 tially seroconverted to multiple autoantibodies. Children
 124 were assigned to each subset irrespectively of whether the

125 specific autoantibody was among the first islet autoantibod- 178
 126 ies to appear or appeared at a later time during follow-up. 179
 127 For example, if a child developed autoantibodies to IAA 180
 128 first and autoantibodies to GADA later, the child would be 181
 129 assigned to both the IAA and GADA subset. 182

130 We used a novel shared parameter joint model approach 183
 131 to assess the association between the longitudinal autoanti- 184
 132 body titers from seroconversion with the time to T1D. Joint 185
 133 models allow the incorporation of longitudinal titers as 186
 134 time-varying covariates into the survival model of progres- 187
 135 sion to T1D by estimating a longitudinal model and a pro- 188
 136 portional hazards model, using a joint likelihood for both 189
 137 submodels [16]. We further extended this model to a more 190
 138 flexible joint model, where we were able to assess hetero- 191
 139 geneous and nonlinear individual biomarker trajectories 192
 140 and to explore complex associations between the biomark- 193
 141 ers and the time to event [17]. We refer to Appendix for 194
 142 further details. Using this novel approach we specified the 195
 143 autoantibody titers over time as smooth, nonlinear, subject- 196
 144 specific trajectories in the longitudinal model. Furthermore 197
 145 we allowed the association between the modeled trajec- 198
 146 tories and the time to T1D to be time-varying in our main 199
 147 analysis. In additional explorative analyses we allowed the 200
 148 association to differ between subjects with different charac- 201
 149 teristics, and to differ over time between subjects with dif- 202
 150 ferent characteristics. 203

151 We fitted these models for each of the three autoantibod- 204
 152 ies IAA, GADA and IA2A, separately, within each autoan- 205
 153 tibody-specific subset. In the longitudinal submodels, we 206
 154 assessed the associations of each autoantibody titer with 207
 155 (a) age at seroconversion of the respective autoantibody, (b) 208
 156 a binary variable indicating whether the autoantibody was 209
 157 among the first autoantibodies to appear, and (c) two time- 210
 158 varying binary variables indicating which of the other two 211
 159 autoantibodies were present at each observed time point. 212
 160 In each proportional hazards submodel, we assessed the 213
 161 associations of the smooth subject-specific autoantibody 214
 162 trajectories from the longitudinal model with progression 215
 163 time from seroconversion of the respective autoantibody 216
 164 to T1D. Baseline covariates were (a) the age at serocon- 217
 165 version of the respective autoantibody and (b) whether the 218
 166 autoantibody was among the first autoantibodies to appear. 219
 167 We further assessed whether the association between the 220
 168 autoantibody trajectories and the time to T1D differed over 221
 169 time between subjects with and without a first-degree rela- 222
 170 tive with T1D or between girls and boys. Additionally we 223
 171 checked for differences in the association between HLA 224
 172 genotypes. Due to the limited size of certain HLA sub- 225
 173 groups we modeled this association as time-constant. 226

174 All models were estimated within a Bayesian frame- 227
 175 work using the R-package *bamlss* [18]. Weakly informative 228
 176 normal priors were used for all coefficients. We report the 229
 177 posterior mean estimates/hazard ratios and 95% credibility

intervals (CI) for all modeled parameters. Bayesian CIs 178
 can be interpreted as the interval in which the popula- 179
 tion parameter lies with a given probability (here 95%). 180
 We assessed convergence of the Markov chains by visual 181
 inspection of traceplots and conducted sensitivity analyses 182
 with regard to prior specification. All calculations were 183
 carried out with R version 3.2.5 [19]. 184

185 Results

186 Table 1 shows the study characteristics in each subset, i.e., 187
 188 the subsets of children who developed IAA, GADA or 188
 189 IA2A autoantibodies, respectively, at any time during fol- 189
 190 low-up. In most cases, either IAA, GADA or both were pre- 190
 191 sent at the time of the first seroconversion, whereas IA2A 191
 192 occurred at a later time point. The children seroconverted 192
 193 to the different autoantibodies at different median ages 193
 194 ($p < 0.001$, Kruskal–Wallis Test) with IAA seroconversion 194
 195 taking place at a lower median age. Apart from that, chil- 195
 196 dren with different autoantibodies were similar regarding 196
 the progression time to T1D and other variables. 197

197 The individual autoantibody patterns over time after 197
 198 seroconversion were heterogeneous, but on average IAA 198
 199 titers declined after an initial increase, and GADA and 199
 200 IA2A titers increased shortly after seroconversion and 200
 201 remained relatively stable thereafter (Supplementary 201
 202 Fig. 1). 202

203 In the joint modeling of autoantibody titers over time 203
 204 and the time to T1D, we observed for all autoantibod- 204
 205 ies a positive association between the titer and the risk of 205
 206 progression to T1D. Titers over time were lower for sub- 206
 207 jects who seroconverted at an older age for the respective 207
 208 autoantibody, and higher if the respective autoantibody 208
 209 appeared at the initial seroconversion, and if other autoan- 209
 210 tibodies were present (Table 2). For each autoantibody, a 210
 211 higher age at the respective seroconversion was also asso- 211
 212 ciated with lower risk of progression to clinical T1D. For 212
 213 example, children had a hazard ratio [95% CI] of 0.84 213
 214 [0.72, 0.98] if they seroconverted one year later for IAA. 214

215 We further investigated whether the association between 215
 216 the estimated trajectories of autoantibodies and the progres- 216
 217 sion to T1D was time-varying or constant. By using our 217
 218 approach, we observed that the association was time-vary- 218
 219 ing for IAA and GADA with the association being highest 219
 220 early after seroconversion and decreasing over time (Fig. 1) 220
 221 and stronger for IAA than GADA: The hazard ratio for IAA 221
 222 (per transformed unit) was 3.38 [2.66; 4.38] at 6 months 222
 223 after seroconversion, 3.02 [2.44, 3.81] at 12 months after 223
 224 seroconversion and 2.02 [1.55, 2.68] at 36 months after 224
 225 seroconversion (Table 3) with an average decrease in the 225
 226 hazard ratio of 10% [95% CI 2, 18%] every 6 months. The 226
 227 hazard ratio for GADA (per transformed unit) was 1.63 227

Table 1 Description of the study population by type of persistent autoantibody

Variable	Total	Type of persistent autoantibody		
		IAA	GADA	IA2A
Total number of children	613	442	466	288
Age at respective seroconversion (years)	2.2 (1.2, 3.8)	2.0 (1.1, 3.5)	2.7 (1.6, 4.2)	2.8 (1.9, 4.5)
Girls	268 (44%)	200 (45%)	212 (45%)	112 (39%)
Country				
US	206 (34%)	136 (31%)	166 (36%)	94 (33%)
Finland	153 (25%)	125 (28%)	109 (23%)	85 (30%)
Germany	47 (8%)	40 (9%)	32 (7%)	22 (8%)
Sweden	207 (34%)	141 (32%)	159 (34%)	87 (30%)
Child having a first-degree relative with T1D	128 (21%)	105 (24%)	97 (21%)	71 (25%)
HLA-DR genotype				
DR3/4	311 (51%)	241 (55%)	251 (54%)	148 (51%)
DR4/4	106 (17%)	74 (17%)	81 (17%)	64 (22%)
DR4/8	92 (15%)	71 (16%)	51 (11%)	46 (16%)
DR3/3	76 (12%)	30 (7%)	64 (14%)	16 (6%)
Other	28 (5%)	26 (6%)	19 (4%)	14 (5%)
Additionally autoantibody positive for				
IAA			302 (65%)	252 (88%)
GADA		302 (68%)		237 (83%)
IA2A		252 (57%)	237 (51%)	
Autoantibody present at first seroconversion		353 (80%)	344 (74%)	40 (14%)
Number of children who developed T1D	175 (29%)	162 (37%)	134 (29%)	127 (44%)

Values are reported as *n* (% of non-missing observations) for categorical variables and median (interquartile range) for continuous variables

Table 2 Posterior mean estimates of coefficients (β) and hazard ratios with corresponding 95% credibility intervals from joint models of autoantibody trajectories (IAA, GADA and IA2A, as estimated in longitudinal submodels) and progression to T1D (survival submodels)

Autoantibodies	Covariate	Longitudinal models		Survival models	
		β	95% CI	HR	95% CI
IAA	IAA present at first seroconversion	0.35	0.20, 0.51	0.66	0.42, 1.02
	IAA seroconversion age (years)	-0.10	-0.13, -0.07	0.84	0.72, 0.98
	GADA positive (time-varying variable)	0.31	0.24, 0.37	^a	^a
	IA2A positive (time-varying variable)	0.17	0.12, 0.22	^a	^a
GADA	GADA present at first seroconversion	0.34	0.19, 0.47	0.72	0.51, 1.03
	GADA seroconversion age (years)	-0.06	-0.09, -0.03	0.61	0.52, 0.72
	IAA positive (time-varying variable)	0.25	0.20, 0.32	^a	^a
	IA2A positive (time-varying variable)	0.06	0.02, 0.11	^a	^a
IA2A	IA2A present at first seroconversion	0.31	0.11, 0.50	1.07	0.61, 1.80
	IA2A seroconversion age (years)	-0.05	-0.08, -0.01	0.66	0.56, 0.78
	IAA positive (time-varying variable)	0.22	0.09, 0.36	^a	^a
	GADA positive (time-varying variable)	0.29	0.18, 0.41	^a	^a

Bold font indicates that the 95% CI does not include 0 (for β) or 1 (for HR)

CI credibility interval, HR hazard ratio

^a Covariate only included in the longitudinal submodel

228 [1.20, 2.30] at 6 months after seroconversion, 1.40 [1.07,
229 1.85] at 12 months after seroconversion and 0.85 [0.61,
230 1.17] at 36 months after seroconversion with an average
231 decrease of 9% [1, 15%] every 6 months. For IA2A, the
232 positive association between autoantibody titer and T1D

233 progression was estimated as time-constant: The hazard
234 ratios for IA2A (per transformed unit) were 1.56 [1.04,
235 2.42] at 6 months after seroconversion, 1.53 [1.10, 2.16] at
236 12 months after seroconversion, and 1.44 [1.005, 2.16] at
237 36 months after seroconversion with a negligible average

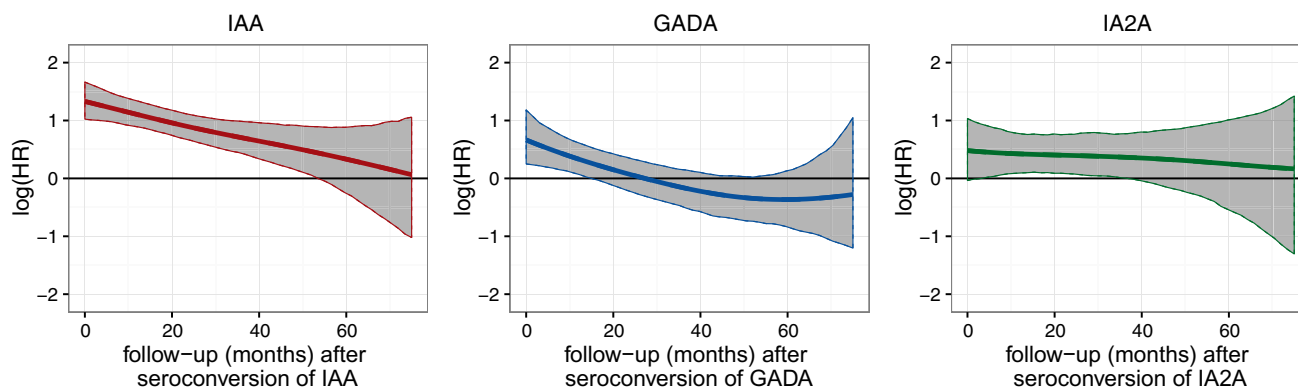


Fig. 1 Posterior mean estimates (lines) and 95% credibility intervals (shaded areas) of $\eta_a(t)$, the time-varying log hazard ratio (HR) of the association between longitudinal autoantibody trajectories and type 1 diabetes

Table 3 Posterior mean hazard ratios (HR) with corresponding 95% credibility intervals (CI) at different time points after seroconversion of each autoantibody for the association between autoantibody trajectories from the longitudinal model and progression time to type 1 diabetes

Autoantibodies	Time point (months)	HR	95% CI
IAA	0	3.78	2.78, 5.28
	6	3.38	2.66, 4.38
	12	3.02	2.44, 3.81
	24	2.43	1.94, 3.02
	36	2.02	1.55, 2.68
	48	1.69	1.17, 2.50
	60	1.39	0.77, 2.42
GADA	0	1.94	1.28, 3.25
	6	1.63	1.20, 2.30
	12	1.40	1.07, 1.85
	24	1.07	0.80, 1.41
	36	0.85	0.61, 1.17
	48	0.73	0.50, 1.04
	60	0.69	0.43, 1.14
IA2A	0	1.62	0.96, 2.81
	6	1.56	1.04, 2.42
	12	1.53	1.10, 2.16
	24	1.49	1.08, 2.13
	36	1.44	1.005, 2.16
	48	1.37	0.82, 2.33
	60	1.28	0.58, 2.74

Bold font indicates that the 95% CI of the respective association does not include the 1

238 decrease of 2% [−8, 13%] every 6 months. As indicated by
 239 the credibility intervals in Fig. 1, positive associations with
 240 T1D progression were observed for IAA up to 54 months
 241 after seroconversion, for GADA up to 18 months after sero-
 242 conversion and for IA2A between 6 and 36 months after

243 seroconversion. The traceplots indicated satisfactory conver-
 244 gence of the Markov chains (Supplementary Figs. 4–6),
 245 and sensitivity analyses showed robustness against different
 246 prior specifications (Supplementary Fig. 7).

247 We further observed differences in the time-varying
 248 association of autoantibodies with progression to T1D
 249 between children with and without a first-degree relative
 250 with T1D. For all autoantibodies the associations were
 251 higher among children with a first-degree relative at early
 252 time points and decreased more strongly within this group
 253 (Fig. 2, upper panel). For IAA, the associations between
 254 the two groups differed from seroconversion until about
 255 15 months thereafter, as indicated by the credibility bands
 256 of the differences (Fig. 2, lower panel), but only from 3 to
 257 7 months after seroconversion for GADA and from 3 to
 258 13 months after seroconversion for IA2A. For all autoan-
 259 tibodies HLA subgroups were similar in the association
 260 between autoantibody trajectories and the time to T1D
 261 (Fig. 3). An exception was a higher association for sub-
 262 jects with IAA autoantibodies and the DR3/3 genotype, a
 263 genotype which is less prevalent among IAA positive chil-
 264 dren ($n = 30, 7\%$). In accordance with the difference in
 265 the hazard, the mean titer levels between progressors and
 266 non-progressors differed more strongly within the small
 267 subgroup of DR3/3 than within other HLA genotypes with
 268 non-progressors showing an especially low level (Supple-
 269 mentary Fig. 2). We did not observe consistent differences
 270 in the association over time between girls and boys (Sup-
 271plementary Fig. 3).

272 Discussion

273 In the present study the complex relationship between lon-
 274 gitudinally measured autoantibodies and the risk of pro-
 275 gression to T1D diabetes was explored using a novel joint
 276 modeling approach. We observed potentially time-varying

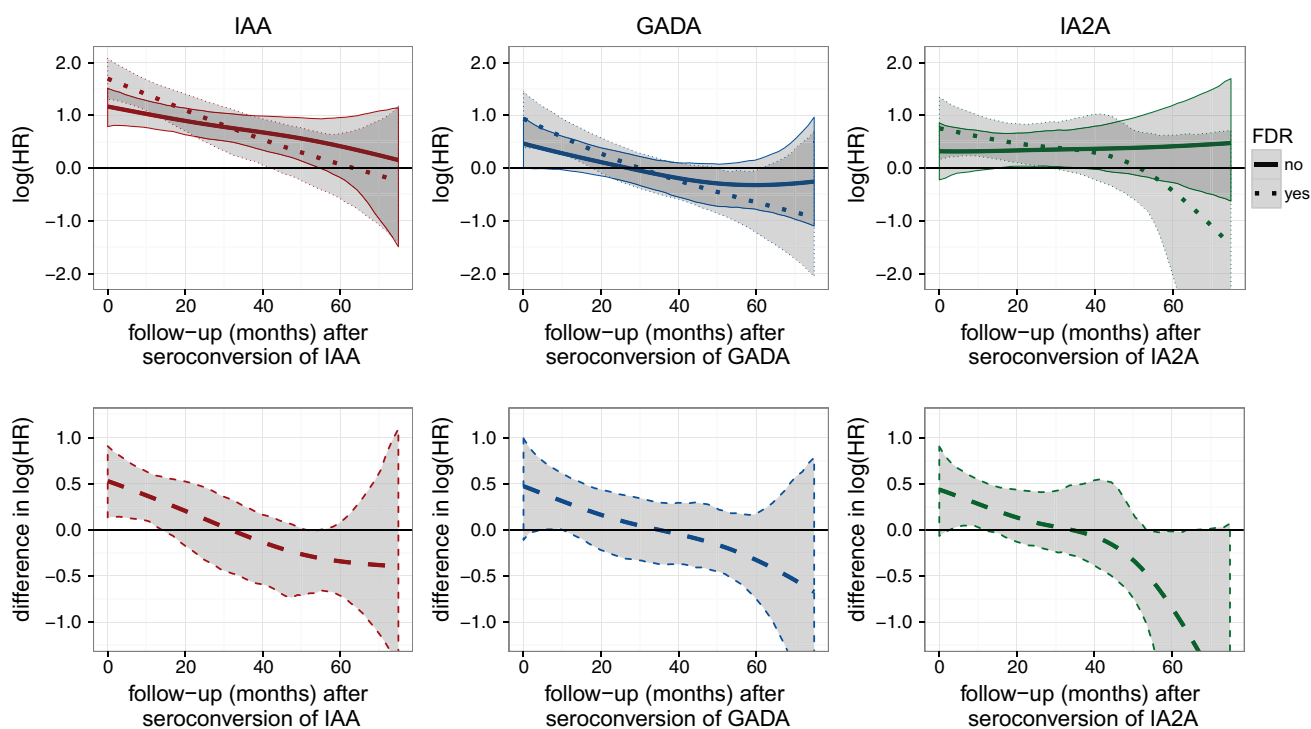


Fig. 2 Posterior mean estimates (lines/dots) and 95% credibility intervals of $\eta_\alpha(t, \text{FDR})$, the time-varying log hazard ratio (HR) of the association between longitudinal autoantibody trajectories and type 1 diabetes (T1D) progression stratified for children that had a first-degree relative (FDR) with T1D or not (upper panel) and of the difference of the association between the groups over time, $\eta_\alpha(t, \text{FDR} = 1) - \eta_\alpha(t, \text{FDR} = 0)$ (lower panel)

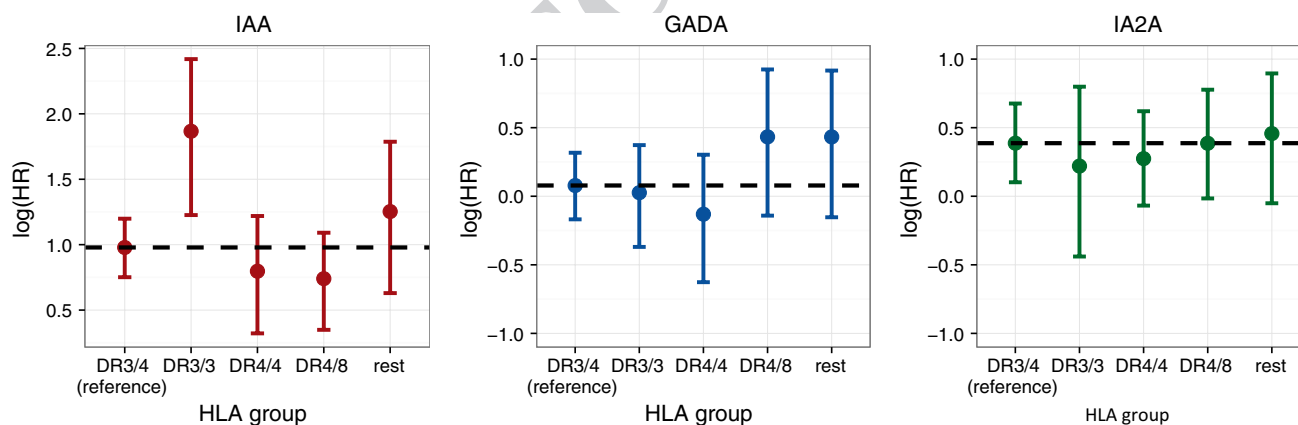


Fig. 3 Posterior mean estimates and 95% credibility intervals of $\eta_\alpha(\text{HLA})$, the time-constant log hazard ratio (HR) of the association between longitudinal autoantibody trajectories and type 1 diabetes progression, per HLA genotype. The dashed line represents the estimated log hazard ratio of the reference group

277 positive associations between the autoantibody titers of
 278 IAA and GADA, and the risk of T1D progression, indicat-
 279 ing that the T1D progression risk associated with autoan-
 280 tibody titers was highest shortly after seroconversion of
 281 the respective autoantibody. The hazard ratio was highest
 282 for IAA, especially at early time points. Additionally, we
 283 observed that the associations of the autoantibody titer and

the T1D risk early after seroconversion were more pro-
 nounced in children with first-degree relatives with T1D.

These results were in line with earlier results from other
 cohorts, where initial and mean IAA and IA2A titers were
 shown to be associated with the risk of progression [2, 6,
 20] as well as from a more recent and methodologically
 advanced study based on the TEDDY data. In this study

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291 the relationship between titers of the same autoantibodies
292 over time and the risk of progression to T1D was modeled
293 assuming a time-constant association [7]. By using mean
294 levels of the respective autoantibodies as time-varying pre-
295 dictors in a Cox model, the authors could show a positive
296 association between autoantibody titers and the time to
297 T1D progression for IAA and IA2A in their analyses.

298 Potential limitations of this previous approach are,
299 however, that (a) only subject's mean titers until a certain
300 time point are taken into account and not all observed val-
301 ues over time, (b) in a time-varying Cox model the time-
302 varying predictor is assumed constant between observa-
303 tions, and (c) the association between autoantibodies and
304 the risk of progression is assumed to be time-constant.
305 These limitations were addressed by our joint modeling
306 approach. Here, we flexibly modeled the trajectories of all
307 three autoantibodies in each subject as a smooth function of
308 time, i.e., obtaining predictions for the autoantibody titers
309 between the measurements at discrete time points, and
310 could use all this information as a time-varying covariate
311 in the survival model. Additionally, we allowed their asso-
312 ciation with the risk of T1D progression to vary over time
313 and between groups of subjects (children with and without
314 first-degree relatives with T1D as well as boys and girls).
315 In consequence, we were able to explore the association
316 between autoantibodies and the risk of T1D beyond the
317 previous results. For example, we observed that increased
318 GADA titers may predict T1D progression within the first
319 1.5 years after seroconversion, but not thereafter. As this
320 association averages to 0 over the whole time range this
321 association was potentially not captured in the simpler
322 modeling from the previous analysis. Furthermore, our
323 modeling approach revealed that the time-varying associa-
324 tions appear to be more pronounced in children with first-
325 degree relatives with T1D compared to children without.

326 The modeling of autoantibodies as longitudinal bio-
327 markers and the time to clinical T1D poses a challenge
328 due to the nature of the data beyond the aspects men-
329 tioned above. Longitudinal biomarkers usually contain
330 potential random variation both due to the laboratory
331 measurement process as well as short- and long-term
332 biological fluctuations and are only observed until an
333 event occurs. Whereas not accounting for the random
334 fluctuations in a time-varying Cox model might result in
335 an underestimation of the hazard ratio [12], ignoring the
336 latter might distort the estimation of covariate effects in
337 the longitudinal model. By jointly analyzing the longitu-
338 dinal and survival model we could address these issues
339 and gained further insights as to how covariates affected
340 both the autoantibody titers over time, and the risk of
341 T1D progression. We found that earlier seroconversion
342 for the respective autoantibody, if the respective sero-
343 conversion was the initial one, as well as the presence of

344 other autoantibodies was associated with higher autoan-
345 tibody titers. The age at the respective seroconversion
346 was also inversely related to the risk of T1D progression
347 for every autoantibody. While joint modeling approaches
348 allow for detailed and unbiased estimations, they demand
349 a high number of subjects, especially when complex
350 associations are modeled in the survival part. TEDDY
351 is the largest prospective study on the determinants of
352 T1D worldwide and thus offers a unique opportunity to
353 explore the application of joint modeling techniques on
354 these complex relationships due to the high number of
355 subjects and the detailed measurement schedule.

356 Currently, the presented flexible joint model only allows
357 the assessment of one longitudinal biomarker at a time.
358 In consequence, one limitation is that we were not able to
359 combine all three markers into one joint model. We partly
360 addressed this issue by including information on the pres-
361 ence of other autoantibodies and the order of their occur-
362 rence in our model. While they provide insights into the
363 mechanisms of disease progression, a drawback of our
364 results is that they cannot easily be translated from a cohort
365 setting with frequent measurements into clinical practice,
366 as the age at the respective seroconversion plays a crucial
367 role in the prediction of T1D progression risk, but is not
368 readily available in practice.

369 In conclusion, by using state of the art joint modeling
370 techniques we were able to give insights into the complex
371 relationship between longitudinal autoantibody titers and
372 the risk of progression to clinical T1D. Risk stratification
373 basing on autoantibody titers should focus on time points
374 early after seroconversion.

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510 **Compliance with ethical standards**

511 **Conflict of interest** The authors declare that they have no conflict
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513 **Human and animal rights** All procedures followed were in accord-
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