



Alcohol consumption is associated with widespread changes in blood DNA methylation: Analysis of cross-sectional and longitudinal data

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Funding information

British Heart Foundation, Grant/Award Number: SP/04/002; Bundesministerium für Bildung und Forschung; ERAB: The European Foundation for Alcohol Research; National Health and Medical Research Council; National Institute for Health Research; Wellcome Trust; Singapore Translational Research Investigator; Singapore Ministry of Health's National

Abstract

DNA methylation may be one of the mechanisms by which alcohol consumption is associated with the risk of disease. We conducted a large-scale, cross-sectional, genome-wide DNA methylation association study of alcohol consumption and a longitudinal analysis of repeated measurements taken several years apart. Using the Illumina HumanMethylation450 BeadChip, DNA methylation was measured in blood samples from 5606 Melbourne Collaborative Cohort Study (MCCS) participants. For 1088 of them, these measures were repeated using blood samples collected a median of 11 years later. Associations between alcohol intake and blood DNA methylation were assessed using linear mixed-effects regression models. Independent data from

Medical Research Council; H2020 programs, Grant/Award Number: iHealth-T2D, 643774; European Union FP7, Grant/Award Number: 279143; NIHR Official Development Assistance, Grant/Award Number: 16/136/68; NIHR, Grant/Award Number: RP-PG-0407-10371; Wellcome Trust, Grant/Award Numbers: 098381, 090532 and 084723/Z/08/Z; Medical Research Council, Grant/Award Numbers: G0700931 and G0601966; National Institute for Health Research (NIHR) Comprehensive Biomedical Research Centre Imperial College Healthcare NHS Trust; European Foundation for Alcohol Research, Grant/Award Numbers: EA1817 and ERAB 2018; NHMRC, Grant/Award Numbers: 1061177, 1074383, 1043616, 1050198, 1027505, 1026892 and 1011618; Australian NHMRC, Grant/Award Numbers: 504711, 251553 and 209057; Australian National Health and Medical Research Council (NHMRC)

the London Life Sciences Prospective Population (LOLIPOP) (N = 4042) and Cooperative Health Research in the Augsburg Region (KORA) (N = 1662) cohorts were used to replicate associations discovered in the MCCS. Cross-sectional analyses identified 1414 CpGs associated with alcohol intake at $P < 10^{-7}$, 1243 of which had not been reported previously. Of these novel associations, 1078 were replicated ($P < .05$) using LOLIPOP and KORA data. Using the MCCS data, we also replicated 403 of 518 previously reported associations. Interaction analyses suggested that associations were stronger for women, non-smokers, and participants genetically predisposed to consume less alcohol. Of the 1414 CpGs, 530 were differentially methylated ($P < .05$) in former compared with current drinkers. Longitudinal associations between the change in alcohol intake and the change in methylation were observed for 513 of the 1414 cross-sectional associations. Our study indicates that alcohol intake is associated with widespread changes in DNA methylation across the genome. Longitudinal analyses showed that the methylation status of alcohol-associated CpGs may change with alcohol consumption changes in adulthood.

KEYWORDS

alcohol consumption, cross-sectional data, DNA methylation, epigenome-wide association study, EWAS, HM450 assay, longitudinal data

1 | INTRODUCTION

Deoxyribonucleic acid (DNA) methylation is the addition of methyl groups to the 5' carbon of cytosine in cytosine-guanine dinucleotides (CpGs) and is thought to play a role in the development of disease through its influence on gene expression and cellular function.^{1,2} DNA methylation is strongly affected by the underlying genetic DNA sequence,³ sex, age, and ethnicity^{4,5} and is modified by lifestyle factors and environmental exposures such as smoking and adiposity.⁶⁻¹⁰

Alcohol consumption is a major lifestyle risk factor contributing to the worldwide burden of disease, responsible for an estimated 2.7 million deaths and 4% of the global burden of disease annually.¹¹ Even modest use of alcohol may increase disease risk, but greatest risks are observed with heavy and long-term drinking. Alcohol consumption is a potentially modifiable risk factor that can be targeted with preventive interventions at both the policy and the individual levels.¹² Molecular mechanisms such as DNA methylation are thought to underlie or enhance a predisposition to addictions and substance abuse, including alcohol drinking,¹³ but although there is a plausible relationship between alcohol intake and altered 1-carbon metabolism and DNA methylation,¹⁴ most evidence comes from studies that either had small sample size, were not specific to humans, or were carried out using tissues other than blood.¹⁵⁻¹⁸ Identifying associations between alcohol consumption and DNA methylation may be useful to derive biomarkers of alcohol-related conditions. To our knowledge, only one large genome-wide study of blood DNA methylation and alcohol consumption has been conducted.¹⁹

In the present study, we sought to (a) identify novel associations between alcohol consumption and blood DNA methylation and assess

whether these were modified by factors such as age, sex, or genetic prediction to drink alcohol, (b) replicate previously reported associations, (c) assess the reversibility of associations, (d) assess associations with changes in alcohol consumption using longitudinally collected data, and (e) assess overrepresentation of the discovered genes in biological pathways to investigate the potential functional implications of our findings. We used samples from the Melbourne Collaborative Cohort Study (MCCS) to discover potential associations and sought to replicate the findings using samples from the Cooperative Health Research in the Augsburg Region (KORA) and London Life Sciences Prospective Population (LOLIPOP) studies.

2 | MATERIALS AND METHODS

2.1 | Study participants

Between 1990 and 1994 (baseline), 41 513 participants were recruited to the MCCS. The majority (99%) were aged 40 to 69 years, and 41% were men. Southern European migrants were oversampled to extend the range of lifestyle-related exposures.²⁰ Participants were contacted again between 2003 and 2007. Blood samples were taken at baseline and follow-up from 99% and 64% of participants, respectively. Baseline samples were stored as dried blood spots on Guthrie cards for the majority (73%), as mononuclear cell samples for 25% and as buffy coat samples for 2% of the participants. Follow-up samples were stored as buffy coat aliquots and dried blood spots on Guthrie cards.

All participants provided written informed consent, and the study protocols were approved by the Cancer Council Victoria Human Research Ethics Committee.

The present study sample comprised MCCS participants selected for inclusion in one of seven previously conducted nested case-control studies of DNA methylation.²¹⁻²⁵ Controls were matched to incident cases of prostate, colorectal, gastric, lung or kidney cancer, urothelial cell carcinoma, or mature B-cell neoplasms on sex, year of birth, country of birth, baseline sample type, and smoking status (the latter for the lung cancer study only). Participants included in each nested case-control study were free of cancer at baseline. After quality control, methylation data for baseline blood samples (baseline study) were available for 5606 MCCS participants. Methylation

measures were repeated in DNA extracted from blood samples collected on Guthrie cards at follow-up (longitudinal study) and, after quality control, were available for a subset of 1088 of the controls who also had their baseline sample collected on a Guthrie card (Table 1).

2.2 | Alcohol and other variables

At both baseline and follow-up, participants completed questionnaires that included detailed questions on demographic characteristics, medical history, cigarette smoking, alcohol consumption, physical activity, and diet, the latter using food frequency questionnaires. On

TABLE 1 Characteristics of participants in the Melbourne Collaborative Cohort Study (MCCS) at baseline and follow-up visits

| | Cross-sectional Analysis | | Longitudinal Analysis | | | |
|--|---------------------------------|-----|--------------------------------|-----|------------------|-----|
| | | | Baseline data | | Follow-up data | |
| | (N = 5606) | | (N = 1088) | | (N = 1088) | |
| Age in years, median (IQR) | 61 (54-65) | | 59 (51-64) | | 70 (63-76) | |
| Sex, male | 3793 | 68% | 740 | 68% | 740 | 68% |
| Country of birth | | | | | | |
| AU/NZ/Other | 3744 | 67% | 831 | 76% | 831 | 76% |
| Greece | 434 | 8% | 44 | 4% | 44 | 4% |
| Italy | 819 | 15% | 89 | 8% | 89 | 8% |
| UK | 609 | 11% | 124 | 11% | 124 | 11% |
| BMI (kg/m ²), median (range) | 26.9 (24.5-29.5) | | 26.3 (24.1-29.0) | | 26.8 (24.2-29.4) | |
| Smoking status | | | | | | |
| Never | 2519 | 45% | 549 | 50% | 545 | 50% |
| Former, ≥15 y ago | 1152 | 21% | 230 | 21% | 485 | 45% |
| Former, <15 y ago | 1,100 | 20% | 196 | 18% | | |
| Current, <20 cig/d | 322 | 6% | 62 | 6% | 58 | 5% |
| Current, ≥20 cig/d | 513 | 9% | 51 | 5% | | |
| Drinking status | | | | | | |
| Lifetime abstainers | 1314 | 24% | 209 | 20% | 168 | 16% |
| Former drinkers | 569 | 10% | 94 | 9% | 111 | 11% |
| Current drinkers | 3626 | 66% | 759 | 71% | 744 | 73% |
| Missing | 97 | 2% | 26 | 2% | 65 | 6% |
| Median (IQR) for intake in last decade (g/day) | 4.6 (0.0-21.5) missing: N = 97 | | 7.5 (0.0-21.3) missing: N = 26 | | | |
| Median (IQR) for intake in last week (g/day) | 4.3 (0.0-18.7) missing: N = 41 | | 7.4 (0.0-19.0) missing: N = 23 | | | |
| Median (IQR) for lifetime intake (g/day) | 8.4 (0.3-23.5) missing: N = 115 | | 8.8 (1.3-22.3) missing: N = 42 | | | |
| Median (IQR) for intake in last year (g/day) | | | 8.9 (0.3-22.7) missing: N = 70 | | | |
| Median (IQR) for the difference between follow-up and baseline (last week) (g/day) | | | 0.0 (-2.5 to 6.4) | | | |
| Median (IQR) for the difference between follow-up and baseline (lifetime) (g/day) | | | 0.0 (-5.5 to 5.0) | | | |
| Alcohol uptake (non-drinkers at baseline who were drinkers at follow-up) (yes) | | | | | 107 | 10% |
| Alcohol cessation (drinkers at baseline who were non-drinkers at follow-up) (yes) | | | | | 88 | 8% |

both occasions, anthropometric measurements were obtained by trained personnel using standard procedures. Height was only measured at baseline.

Alcohol intake at baseline was recorded as frequency and quantity of intake per drinking occasion by type (beer, wine, spirits) and by decade of age starting from 20 years. Participants were also asked about their alcohol intake on each day during the previous week, in terms of the number, measure, and type of drink (eg, two glasses of wine). At follow-up, the frequency and quantity of intake, by type of drink, during the previous calendar year were assessed as described above. Grams per day (g/day) were calculated as reported previously.²⁶ Participants who reported a weekly or current intake higher than 200 g/day were excluded. Four alcohol consumption variables were considered: g/day in the last week (continuous), g/day in the current decade (continuous), g/day over the lifetime (continuous), and drinking status (never, former, or current; based on current decade and lifetime variables).

2.3 | DNA methylation and genetic data

Methods relating to DNA extraction and bisulfite conversion, DNA methylation data processing, normalisation and quality control, and genotyping are described in the Supporting Information, methods.

2.4 | Statistical analysis

2.4.1 | Methylome-wide association study

We assessed cross-sectional associations at each individual CpG by regressing DNA methylation M-values on alcohol consumption using linear mixed-effects regression models, using the function *lmer* from the R package *lme4*. Alcohol intake was represented using three continuous variables (consumption in the previous year, consumption in the previous week, and lifetime consumption, in g/day) that were modelled separately. Models were adjusted by fitting fixed effects for age (continuous), sex, smoking status (never; former, ≥ 15 years ago; former, < 15 years ago; current, < 20 cigarettes per day; current, ≥ 20 cigarettes per day), body mass index (BMI) (≤ 25 , > 25 to ≤ 30 , and > 30 kg/m²), country of birth (Australia/New Zealand, Italy, Greece, UK), sample type (peripheral blood mononuclear cells, dried blood spots, buffy coats), and white blood cell composition (percentage of CD4+ T cells, CD8+ T cells, B cells, NK cells, monocytes, and granulocytes, estimated using the Houseman algorithm²⁷), and random effects for study, plate, and chip. A significance threshold of P value $< 10^{-7}$ was used to account for multiple testing.⁷ The false discovery rate (FDR) was used to identify suggestive associations.

Interaction terms were tested between alcohol consumption (previous week) and each of age (continuous), sex, smoking status (pseudo-continuous variable: 1, never to 5, current smoker greater than 20 cigarettes/day), BMI (continuous), country of birth (as categorised above), future cancer case status, and a polygenic score for alcohol consumption, the latter derived as described below.

These analyses were restricted to CpGs associated at $P < 10^{-7}$ in the cross-sectional analysis with any of the three continuous alcohol variables ("last week," "current decade," and "lifetime" intake).

Associations for each type of alcohol (beer, wine, spirits) were assessed by including the three variables in a same model, so that intake of each type was adjusted for the two others. This was done for the current decade and lifetime alcohol intake variables because type of alcoholic beverage was not determined for the "alcohol last week" variable.

Sensitivity analyses were conducted to assess potential confounding by (a) fitting the same models without adjustment for smoking or BMI, (b) fitting the same models with additional adjustment for socioeconomic status, educational attainment, a physical activity score based on metabolic equivalents,²⁸ and a score of healthy dietary habits,²⁹ (c) examining whether methylation was associated with these covariates, and (d) using less stringent adjustment for white blood cell composition by omitting each one of the six cell-type variables.

2.4.2 | Differentially methylated regions

We also tested for regions of contiguous CpGs associated with alcohol using the DMRcate package,³⁰ with default options (bandwidth of $\lambda = 1000$ base pairs and a scaling factor $C = 2$ for the Gaussian kernel smoother). The models were adjusted using as fixed effects age (continuous), sex, smoking status, country of birth, sample type, white cell composition (all categorised as described above), study, and plate. Associations with region-wise Stouffer P value $< 10^{-7}$ were considered statistically significant.

2.4.3 | Replication of novel associations

CpGs associated with alcohol intake at $P < 10^{-7}$ in the MCCS were selected for replication using data from the KORA cohort ($N = 1662$, assessed in 1999-2001) including German participants and the LOLIPOP cohort ($N = 4042$, assessed in 2003-2008), including predominantly Asian participants, respectively.^{31,32} Alcohol intake in KORA was defined as the average of alcohol intake "in previous day" and "in previous week." Alcohol intake in LOLIPOP was defined over a week (Supporting Information, methods). Each cohort applied a normalisation method based on control probes³³ and adjusted models for the same covariates defined in a similar way to those used in the MCCS analyses (Table S1; Supporting Information, methods). Results from the two cohorts were pooled using fixed-effects meta-analysis with inverse-variance weights.³⁴ An association was considered replicated if $P < .05$ and the direction of association was the same as in the MCCS.⁷

2.4.4 | Polygenic score for alcohol consumption

A polygenic score for alcohol consumption was constructed using MCCS data based on the genome-wide association study by Clarke and colleagues, which identified 14 single nucleotide polymorphisms

(SNPs) associated with alcohol consumption using UK Biobank data.³⁵ Data for 13 out of the 14 SNPs were available and were combined using the formula: $Polygenic\ score[i] = b_1 * d_{1,i} + \dots + b_{13} * d_{13,i}$, where $d_{k,i}$ is the imputed allele dosage of variant k for person i , and b_k the per-allele regression coefficient reported for SNP k .

2.4.5 | Replication of previously reported associations with alcohol consumption

Using MCCS data, we assessed replication ($P < .05$) of associations with alcohol consumption (g/day) reported in a recent pooled, large-scale analysis of Europeans and African Americans.¹⁹ A total of 518 CpGs were considered, comprising 363 identified for participants of European ancestry and a further 155 CpGs identified for those of African ancestry.

2.4.6 | Reversibility of associations

We calculated regression coefficients for comparisons of “former” with “never,” “current” with “never,” and “current” with “former” drinkers using MCCS data. As there were too few never drinkers in the KORA data ($N = 22$), we only considered the comparison of former with current; we pooled the latter using fixed-effects meta-analysis.

In the MCCS, we calculated a “reversibility coefficient,” expressed as a percentage defined as coefficient (former compared with current)/coefficient (never compared with current). These analyses were undertaken for CpGs with $P < 10^{-7}$ in the MCCS methylome-wide association study (EWAS).

2.4.7 | Longitudinal associations

We further examined longitudinal associations with alcohol consumption for CpGs with $P < 10^{-7}$ in the MCCS EWAS, incorporating data from follow-up. Linear mixed-effects regression models were used to assess the association between changes in DNA methylation (outcome) and changes in alcohol consumption (exposure) at each CpG. The change in alcohol consumption was computed in g/day as the difference between follow-up (alcohol intake in previous year) and baseline (previous week intake); study was included as a random effect, and the following variables were included as fixed effects: baseline alcohol intake, baseline BMI and change in BMI (continuous), baseline age and change in age (continuous), sex, smoking status at baseline (as defined previously), smoking status at follow-up (yes/no), country of birth (as defined previously), baseline cell composition (as defined previously), change in each cell-type composition (continuous), and baseline methylation M value. The change in methylation was calculated as the difference between follow-up and baseline ComBat-normalised methylation M values. The same analyses were conducted for the KORA cohort, in which methylation measures taken approximately 7 years later (2006-2008) were available for 1332 participants (Supporting Information, methods). As adjustment for baseline methylation in analyses of change in methylation may lead to bias in some

circumstances,³⁶ we conducted a sensitivity analysis using models without adjustment for baseline methylation in the MCCS.

2.5 | Pathway analyses

We used the *gometh* function from the *missMethyl* package³⁷ for pathway analyses assessing over-representation relative to all Kyoto Encyclopedia of Genes and Genomes (KEGG) and Gene Ontology (GO) pathways.^{38,39} To investigate potentially different biological pathways underlying (a) acute compared with chronic alcohol consumption and (b) most dynamic compared with least dynamic methylation sites, *gometh* was applied, respectively, to (a) CpGs associated with last week and lifetime alcohol intake and (b) CpGs with a reversibility coefficient greater and lower than 50%. A nominal P value lower than .05 was considered to indicate a potentially relevant pathway.

All statistical analyses were performed using the software R (version 3.4.0).

3 | RESULTS

Altogether, 5606 MCCS participants were included in the cross-sectional analysis; their median age was 61 years (IQR: 54-65), 68% were males, and alcohol intakes were wide-ranging (Table 1). There were moderate-to-high correlations between the baseline alcohol variables and between the baseline and follow-up variables, with Spearman correlations ranging from 0.68 to 0.76 (Table S2). Participants in the longitudinal analysis were younger and generally had healthier lifestyle.

3.1 | Methylome-wide association study

Across the three cross-sectional continuous alcohol variables considered, we observed 1414 associations with $P < 10^{-7}$. The most statistically significant associations are presented in Table 2. There were 1318, 358, and 392 CpGs associated with alcohol intake over the last week, decade, and lifetime, respectively (Table S3, Figure S1). Associations were consistently stronger for alcohol intake in the previous week (Figure S2).

Of these 1414 associations, 1243 were novel (ie, not reported in Liu et al¹⁹), and the CpGs were located in 842 genes; 241 CpGs were intergenic. Compared with the rest of the HM450 assay, alcohol-associated CpGs were over-represented in gene bodies (binomial proportion test, $P = .001$), promoter regions TSS1500 ($P = 5 \times 10^{-6}$) and 5' UTR ($P = 3 \times 10^{-8}$); unannotated regions were underrepresented ($P = 2 \times 10^{-5}$). These 1414 associations corresponded to 1084 unique clusters when considering that CpGs separated by less than 50 kb of each other formed a single genomic region (Table S4). For the overwhelming majority (99%) of associations, greater alcohol intake was associated with lower methylation. The number of CpGs associated with alcohol intake over the last week, decade, and lifetime was 16 732,

TABLE 2 The 15 most statistically significant novel associations ($N = 1243$) between alcohol intake and blood DNA methylation discovered in the MCCS and replication in external cohorts KORA and LOLIPOP

| CpG | Chr. | Position | Gene | Location | Mean | Discovery Data Set (MCCS) | | | | | | Replication Data Sets (KORA and LOLIPOP) | | | | | | |
|------------|------|-----------|---------|----------|------|---------------------------|---------------------|----------------|---------------------|---------------|---------------------|--|--------------------|---------|---------------------|--------|---------------------|-----|
| | | | | | | LifETIME | | Current decade | | Previous week | | KORA | | LOLIPOP | | Pooled | | |
| | | | | | | Est. | P | Est. | P | Est. | P | Est. | P | Est. | P | Est. | P | |
| cg26856289 | 1 | 24307516 | SFRS13A | TSS1500 | 0.34 | -1.8 | 4×10^{-17} | -1.6 | 5×10^{-18} | -2.1 | 3×10^{-25} | -1.3 | 4×10^{-5} | -2.1 | 3×10^{-19} | -1.8 | 2×10^{-22} | Yes |
| cg00422488 | 9 | 100747767 | ANP32B | Body | 0.11 | -2.4 | 2×10^{-12} | -2.3 | 3×10^{-13} | -3.3 | 7×10^{-23} | -1.9 | 4×10^{-4} | -2.9 | 8×10^{-17} | -2.6 | 3×10^{-19} | Yes |
| cg11826008 | 20 | 34249284 | CPNE1 | 5'UTR | 0.30 | -1.7 | 3×10^{-13} | -1.3 | 8×10^{-11} | -2.2 | 1×10^{-22} | -0.9 | 3×10^{-2} | -1.4 | 5×10^{-10} | -1.3 | 6×10^{-11} | Yes |
| cg03607573 | 11 | 93471889 | TAF1D | Body | 0.39 | -1.5 | 1×10^{-11} | -1.4 | 8×10^{-12} | -2.1 | 2×10^{-21} | -1.7 | 2×10^{-5} | -2.1 | 1×10^{-21} | -2.0 | 8×10^{-26} | Yes |
| cg20625334 | 2 | 152991620 | STAM2 | Body | 0.58 | -1.9 | 9×10^{-13} | -1.6 | 2×10^{-12} | -2.4 | 6×10^{-21} | -1.7 | 3×10^{-4} | -1.9 | 3×10^{-14} | -1.9 | 3×10^{-17} | Yes |
| cg23684449 | 16 | 46919194 | GPT2 | Body | 0.53 | -1.0 | 5×10^{-9} | -0.9 | 3×10^{-9} | -1.5 | 8×10^{-21} | -0.8 | 2×10^{-4} | -0.9 | 5×10^{-10} | -0.9 | 4×10^{-13} | Yes |
| cg15114651 | 19 | 47289410 | SLC1A5 | TSS1500 | 0.56 | -1.0 | 3×10^{-15} | -0.8 | 6×10^{-12} | -1.1 | 2×10^{-20} | -1.0 | 8×10^{-8} | -1.0 | 2×10^{-11} | -1.0 | 7×10^{-18} | Yes |
| cg06829760 | 2 | 16845412 | FAM49A | 5'UTR | 0.48 | -1.3 | 5×10^{-13} | -1.2 | 2×10^{-14} | -1.5 | 6×10^{-20} | -0.7 | 4×10^{-3} | -1.1 | 3×10^{-10} | -0.9 | 1×10^{-11} | Yes |
| cg26841068 | 1 | 203456691 | PRELP | 3'UTR | 0.41 | -1.2 | 1×10^{-12} | -1.0 | 3×10^{-12} | -1.4 | 6×10^{-20} | -1.1 | 6×10^{-6} | -1.4 | 7×10^{-15} | -1.3 | 3×10^{-19} | Yes |
| cg05288253 | 16 | 89552259 | ANKRD11 | 5'UTR | 0.59 | -1.1 | 4×10^{-12} | -1.1 | 3×10^{-15} | -1.4 | 1×10^{-19} | -1.1 | 5×10^{-8} | -1.1 | 1×10^{-12} | -1.1 | 3×10^{-19} | Yes |
| cg13408712 | 11 | 129817591 | PRDM10 | TSS200 | 0.61 | -0.8 | 3×10^{-10} | -0.9 | 3×10^{-13} | -1.2 | 4×10^{-19} | -0.9 | 4×10^{-7} | -0.4 | 2×10^{-3} | -0.6 | 5×10^{-8} | Yes |
| cg04367503 | 19 | 13121571 | NFIX | Body | 0.52 | -3.1 | 4×10^{-13} | -2.8 | 2×10^{-13} | -3.7 | 6×10^{-19} | -2.6 | 4×10^{-5} | -3.7 | 2×10^{-14} | -3.3 | 9×10^{-18} | Yes |
| cg08899105 | 15 | 93450671 | CHD2 | Body | 0.47 | -0.9 | 4×10^{-8} | -0.9 | 3×10^{-9} | -1.4 | 1×10^{-18} | -0.2 | 4×10^{-1} | -0.8 | 3×10^{-9} | -0.7 | 1×10^{-8} | Yes |
| cg04261072 | 13 | 79977499 | RBM26 | Body | 0.49 | -1.4 | 2×10^{-9} | -1.3 | 7×10^{-11} | -1.9 | 2×10^{-18} | -0.4 | 4×10^{-1} | -1.1 | 6×10^{-8} | -1.0 | 2×10^{-7} | Yes |
| cg07577824 | 20 | 17943403 | SNX5 | Body | 0.38 | -1.1 | 2×10^{-8} | -1.1 | 1×10^{-9} | -1.7 | 3×10^{-18} | -0.7 | 2×10^{-2} | -1.3 | 1×10^{-7} | -1.0 | 2×10^{-8} | Yes |

Note. All models were adjusted by fitting fixed effects for age, sex, smoking status, BMI, country of birth, sample type and white blood cell composition (percentage of CD4+ T cells, CD8+ T cells, B cells, NK cells, monocytes, and granulocytes, estimated using the Houseman algorithm), and batch effects. Est. pooled and P pooled are results from the fixed-effects meta-analysis of results from KORA and LOLIPOP. Regression coefficients are given for intakes in grams per day and multiplied by 1000. The full results are presented in the Supporting Information.

Abbreviations: Chr., chromosome; Est., regression coefficients from linear mixed regression models (MCCS) or linear regression models (KORA and LOLIPOP); KORA, Cooperative Health Research in the Augsburg Region; LOLIPOP, London Life Sciences Prospective Population; MCCS, Melbourne Collaborative Cohort Study.

5751, and 6585, respectively when considering an FDR-adjusted $P < .05$ (Table S3).

3.1.1 | Genomic regions associated with alcohol intake

Applying DMRcate to the MCCS data, association with alcohol intake in last week was observed for methylation at 221 regions of contiguous CpGs ($P < 1.0 \times 10^{-7}$) (Table S19). Several of the most statistically significant regions contained a large number of CpGs annotated to genes that were not ranked among the highest in the EWAS; for example, a cluster was identified in each of *MAB21L1* (23 CpGs, $P = 2 \times 10^{-38}$); *HOXA4* (25 CpGs, $P = 3 \times 10^{-31}$); *GAS5* (9 CpGs, $P = 4 \times 10^{-30}$); *SNHG1* (9 CpGs, $P = 6 \times 10^{-30}$).

3.1.2 | Replication of novel associations using external data

Of the 1243 novel associations, 1078 (87%) were replicated using data from KORA and LOLIPOP. Replication rates were 87%, 89%, and 93% for CpGs associated with alcohol intake over the last week, decade, and lifetime, respectively ($P < .05$; Table 2 and Table S5).

3.1.3 | Interaction analyses

Using the Bonferroni correction for multiple testing ($P = .05/1414 = 3.5 \times 10^{-5}$) and the last week alcohol intake variable, we observed stronger associations for women at *cg13446906* (*MIR548F5*), and *cg22363327* (*SFRS13B*), and weaker associations for smokers at *cg05104080* (*ILKAP*), *cg17058475* (*CPT1A*), and *cg01395047* (*TLR9*). At $P < .05$, stronger associations were observed for women at 200 CpGs (test for binomial proportions, $P = 6 \times 10^{-29}$); weaker associations were observed for participants with a higher smoking score ($N = 159$, $P = 6 \times 10^{-20}$) and a higher BMI ($N = 165$, $P = .003$) (Table 3; Table S15).

For the MCCS sample with genetic and DNA methylation data ($N = 3859$), the polygenic score was positively associated with alcohol intake in previous week ($P = 9 \times 10^{-5}$) (Table S13). Weaker associations ($P < .05$) for participants with higher polygenic score were observed for 156 CpGs ($P = 2 \times 10^{-19}$), although none passed the Bonferroni correction (Table 3 and Tables S12 and S13). There was only weak evidence of association with methylation at CpGs located in regions including SNPs in the polygenic score (Table S14).

3.1.4 | Alcohol types

For alcohol intake in the previous decade, the regression coefficients were the greatest for beer intake for 917 (65%) CpGs, for spirit intake for 309 (22%) CpGs, and for wine intake for 188 (13%) CpGs, whereas these percentages were 27%, 32%, and 41%, respectively, for lifetime alcohol intake (Table S16).

3.1.5 | Sensitivity analyses

Age, sex, smoking, BMI, country of birth, and cell composition were associated with methylation at many alcohol-associated CpG sites. Adjustment for smoking, but not BMI, made a substantial difference to the estimated coefficients; for the "previous week" alcohol intake variable, 1985 associations were observed when no adjustment for smoking was made. Adjusting the models for additional health-related variables or not adjusting for all white blood cell types made virtually no difference to the results (Supporting Information).

3.2 | Replication of previously reported associations using MCCS data

We examined the replication of the associations between alcohol intake and whole-blood DNA methylation previously reported in Liu et al with $P < 10^{-7}$.¹⁹ Of the 518 associations, we replicated 403 (78%) at $P < .05$, using the MCCS "previous week" alcohol intake

TABLE 3 Interaction between alcohol intake and other factors in association with DNA methylation changes for 1414 alcohol-related CpG sites

| Variable | Positive Interaction ($P < .05$) | Negative Interaction ($P < .05$) | P^* | Genes Annotated to CpGs for which Strongest Evidence of Interaction Observed ($P < .001$) (full results in Table S13) |
|--------------------------------------|------------------------------------|------------------------------------|---------------------|--|
| Sex (F vs M) | 200 | 5 | 6×10^{-29} | <i>MIR548F5</i> , <i>SFRS13B</i> , <i>MIR548F5</i> , <i>MGAT5B</i> , <i>NSD1</i> , <i>ABI3</i> , <i>SARM1</i> , <i>NFIX</i> , <i>SC65</i> , <i>NTF3</i> , <i>ADRA2A</i> , <i>ESRP2</i> , <i>ZNF532</i> |
| Age | 53 | 11 | 0.07 | <i>SLC7A11</i> |
| Smoking score ^a | 7 | 159 | 6×10^{-20} | <i>TLR9</i> , <i>CPT1A</i> , <i>ILKAP</i> , <i>MYB</i> , <i>SLC7A11</i> , <i>HDAC1</i> , <i>GPR39</i> , <i>AKR1A1</i> , <i>RPP21</i> , <i>C6orf227</i> |
| Body mass index | 7 | 65 | 0.003 | <i>BCAN</i> , <i>JAK1</i> , <i>TOP1MT</i> |
| Case status (Case vs. control) | 35 | 16 | 1 | - |
| Polygenic score for alcohol drinking | 8 | 156 | 2×10^{-19} | <i>RPL6</i> , <i>DOPEY2</i> , <i>NPM1</i> , <i>MSI2</i> , <i>VPS54</i> |

* P value from test of binomial proportions that the greatest number (eg, 200 for sex) is greater than 35 (number of associations expected by chance). P values were calculated assuming independence between methylation values at the 1414 CpGs.

^aSmoking was coded as a pseudo-continuous score: 0: never; 1: former, ≥ 15 years ago; 2: former, < 15 years ago; 3: current, < 20 cig/day; 4: current, ≥ 20 cig/day.

variable; 169 reached genome-wide significance $P < 10^{-7}$ (Table 4, Table S5). Replication was substantially higher for associations identified in European ancestry individuals (335/363, 92%), compared with those of African ancestry (68/155, 44%).

3.3 | Reversibility of associations

The current decade alcohol consumption variable of the MCCS was classified into current, former, and never. Of the 1414 CpGs considered, 280 were differentially methylated ($P < .05$) between former and never drinkers and 282 were differentially methylated ($P < .05$) between former and current drinkers, with only two overlapping CpGs. The reversibility coefficients, assessing, for former drinkers, the degree of return to the average methylation levels of never drinkers, were wide-ranging (median: 47%, IQR: 19%-79%) (Table S7). In KORA, there was also substantial reversibility, as 332 CpGs were differentially methylated when comparing former with current drinkers. A total of 86 CpGs were differentially methylated in both data sets for the comparison of former with current drinkers, and 530 CpGs when pooling results from both cohorts ($P < .05$); the most significant associations are presented in Table 5.

3.4 | Longitudinal associations

We further tested the 1414 CpGs with cross-sectional associations for longitudinal associations using the MCCS and KORA. Repeated methylation measures and alcohol information were available a median of 11 and 7 years apart in the MCCS and KORA, respectively. Change in alcohol intake was associated with change in methylation ($P < .05$) for 267 CpGs in the MCCS and for 331 CpGs in KORA, with 92 overlapping associations. After pooling the results, we observed evidence of change over time for 513 CpG sites (Table S8). The most statistically significant longitudinal associations are shown in Table 6. Fewer associations were observed in the MCCS when no adjustment for baseline methylation levels was made ($N = 125$; Supporting Information, material). The analyses of alcohol cessation ($N = 88$ participants, 8%) and uptake ($N = 107$, 10%) revealed 147 and 40 associations, respectively, overlapping little with the 513 identified CpGs ($N = 65$ and $N = 19$, respectively). CpG sites that showed stronger evidence of association in the longitudinal analysis appeared somewhat more reversible in the cross-sectional analysis (Figure S5), and 245 CpGs appeared differentially methylated in both analyses (Table S10).

3.5 | Gene description and pathway analyses

The role of each of the 842 identified genes is summarised in Table S20. The *gometh* function was applied separately to CpG sites associated with lifetime ($N = 392$) and last week alcohol intake ($N = 1318$) and for associations with reversibility coefficients lower and greater than 50% ($N = 753$ and $N = 661$, respectively) (Table S11). The most significant KEGG pathways were for the last week variable: "chronic

myeloid leukaemia," "ribosome," and "glycosaminoglycan biosynthesis"; lifetime variable: "regulation of actin cytoskeleton," "biosynthesis of amino acids," and "platelet activation"; persistent associations: ribosome and "human cytomegalovirus infection"; reversible associations: "cellular senescence" and "MAPK signalling pathway." There was strong evidence for enrichment of many GO pathways, and the strongest associations were observed for the last week variable: "negative regulation of gene expression"; persistent associations: "homophilic cell adhesion via plasma membrane adhesion molecules"; reversible associations: "translation" (Table S12).

4 | DISCUSSION

Our study identified 1481 methylation sites associated with alcohol consumption, including 1078 discovered in the MCCS and replicated in independent cohorts and 403 replicated from a previous large EWAS. An additional 513 CpGs discovered in the cross-sectional analysis were indirectly replicated in the longitudinal analysis. The findings using a less conservative significance threshold than $P < 10^{-7}$ (ie, 16 732 associations observed using FDR adjusted $P < .05$ for previous week variable) indicate that many more alcohol-associated CpGs are likely to exist across the genome. The majority of CpG sites we identified were hypomethylated with increased alcohol intake. The overrepresentation of gene body probes in our analysis could be explained by an overall hypomethylation resulting from the potential disruption of 1-carbon metabolism induced by increased alcohol consumption and might not have specific functional implications. In contrast, aberrant hypomethylation of gene promoter CpGs, which were also overrepresented, might affect specific gene functions, but the functional relevance of our findings could not be tested in the present study.

Alcohol-related hypomethylation appears to be largely reversible upon alcohol cessation; this was inferred from three analyses. First, we observed substantially more and stronger associations with alcohol consumed in the last week than in the last decade or lifetime, indicating that the alcohol intake most relevant to DNA methylation was that closest to blood draw. Similar data were not available from other cohorts to replicate this finding. Second, the cross-sectional comparison of current and former with never drinkers revealed that the difference in terms of DNA methylation between former and current drinkers was on average half that between never and current drinkers, with wide-ranging estimates; we also identified CpG sites that were consistently differentially methylated in former compared with current drinkers in MCCS and KORA. Third, using longitudinal data taken several years apart (11 years in the MCCS and 7 years in KORA), we identified a set of 513 CpG sites that varied with change in alcohol consumption, and 245 of these corresponded to differentially methylated sites in the comparison of former with current drinkers (cross-sectional). The longitudinal analysis had less power because of a lower number of included participants and relatively small variation in drinking status over the periods considered and because the variable reflecting changes in alcohol drinking may have been measured with error. In the MCCS, the assessment of alcohol consumption was not

TABLE 4 Replication of the strongest associations ($P < 10^{-15}$) reported in Liu et al, Mol. Psychiatry, 2016 using MCCS data

| CpG | Chr. | Position | Gene | Location | Discovery Set Ancestry | Est. Liu | P Liu | P for Alcohol Intake in Previous Week (MCCS) | Replicated |
|------------|------|-----------|----------|----------|------------------------|----------|---------------------|--|------------|
| cg11610002 | 3 | 126080368 | | | African | -0.10 | 1×10^{-28} | 0.06 | |
| cg15061231 | 3 | 186353575 | | | African | -0.09 | 9×10^{-28} | 0.6 | |
| cg22524735 | 8 | 144779094 | BREA2 | TSS200 | African | -0.03 | 1×10^{-24} | 0.8 | |
| cg25729907 | 3 | 146263345 | PLSCR1 | TSS1500 | African | -0.05 | 4×10^{-24} | 0.7 | |
| cg25120484 | 3 | 197186156 | | | African | -0.05 | 2×10^{-22} | 1 | |
| cg07710247 | 18 | 55268669 | NARS | 3'UTR | African | -0.04 | 1×10^{-20} | 0.8 | |
| cg05538701 | 6 | 116866658 | FAM26D | 5'UTR | African | -0.04 | 1×10^{-19} | 0.5 | |
| cg02583484 | 12 | 54677008 | HNRNPA1 | Body | white | -0.39 | 2×10^{-19} | 2×10^{-29} | Yes |
| cg13729116 | 4 | 1859262 | LETM1 | TSS1500 | white | -0.18 | 7×10^{-18} | 1×10^{-6} | Yes |
| cg13057576 | 14 | 64692154 | SYNE2 | Body | African | -0.04 | 9×10^{-18} | 0.1 | |
| cg07470207 | 8 | 132828746 | | | African | 0.04 | 1×10^{-17} | 0.5 | |
| cg09935388 | 1 | 92947588 | GFI1 | Body | African | -0.17 | 2×10^{-17} | 2×10^{-6} | Yes |
| cg05593667 | 6 | 35490744 | | | white | -0.25 | 4×10^{-16} | 3×10^{-7} | Yes |
| cg02003183 | 14 | 103415882 | CDC42BPB | Body | African | 0.10 | 7×10^{-16} | 0.04 | Yes |
| cg13620705 | 17 | 73559492 | LLGL2 | Body | African | -0.09 | 9×10^{-16} | 0.2 | |

Note. The full results are presented in the Supporting Information.

Abbreviations: Chr., chromosome; Est. Liu, regression coefficient reported in the study by Liu et al, Mol. Psychiatry, 2016; MCCS, Melbourne Collaborative Cohort Study.

TABLE 5 Reversibility of associations (cross-sectional current vs former vs never) for the fifteen most significant associations in the MCCS EWAS for the "current decade" alcohol intake variable

| CpG | Chr. | Position | Current Decade Intake, g/day | | Former vs Never | | Current vs Never | | Reversibility Coefficient ^a |
|------------|------|-----------|------------------------------|---------------------|-----------------|-------|------------------|---------------------|--|
| | | | Est. | P | Est. | P | Est. | P | |
| cg06690548 | 4 | 139162808 | -6.6 | 7×10^{-70} | -0.01 | 0.8 | -0.12 | 2×10^{-8} | 95% |
| cg14476101 | 1 | 120255992 | -3.4 | 6×10^{-34} | -0.02 | 0.4 | -0.10 | 2×10^{-10} | 81% |
| cg12825509 | 3 | 185648568 | -2.4 | 2×10^{-26} | -0.05 | 0.009 | -0.07 | 2×10^{-8} | 31% |
| cg02711608 | 19 | 47287964 | -2.1 | 2×10^{-24} | -0.01 | 0.5 | -0.04 | 8×10^{-5} | 76% |
| cg18120259 | 6 | 43894639 | -1.9 | 1×10^{-20} | -0.04 | 0.009 | -0.05 | 6×10^{-6} | 14% |
| cg18336453 | 6 | 43082296 | -1.3 | 2×10^{-19} | -0.03 | 0.01 | -0.04 | 5×10^{-8} | 32% |
| cg19693031 | 1 | 145441552 | -2.5 | 5×10^{-19} | -0.02 | 0.3 | -0.06 | 3×10^{-4} | 61% |
| cg11376147 | 11 | 57261198 | -1.5 | 1×10^{-18} | -0.04 | 0.001 | -0.05 | 3×10^{-7} | 7% |
| cg26856289 | 1 | 24307516 | -1.6 | 5×10^{-18} | -0.01 | 0.5 | -0.05 | 4×10^{-6} | 78% |
| cg17058475 | 11 | 68607737 | -3.6 | 9×10^{-18} | -0.07 | 0.03 | -0.16 | 5×10^{-13} | 56% |
| cg16246545 | 1 | 120255941 | -1.7 | 1×10^{-17} | 0.00 | 0.9 | -0.04 | 1×10^{-4} | 104% |
| cg06644515 | 1 | 173834831 | -1.6 | 6×10^{-17} | 0.00 | 0.9 | -0.04 | 3×10^{-4} | 106% |
| cg02583484 | 12 | 54677008 | -1.4 | 1×10^{-16} | -0.02 | 0.2 | -0.05 | 3×10^{-7} | 62% |
| cg15804598 | 17 | 43224418 | -1.1 | 1×10^{-15} | -0.02 | 0.1 | -0.04 | 2×10^{-7} | 52% |
| cg00252472 | 6 | 150739173 | -2.7 | 1×10^{-15} | -0.03 | 0.3 | -0.08 | 5×10^{-6} | 63% |

Note. All models were adjusted by fitting fixed effects for age, sex, smoking status, BMI, country of birth, sample type and white blood cell composition (percentage of CD4+ T cells, CD8+ T cells, B cells, NK cells, monocytes, and granulocytes, estimated using the Houseman algorithm), and batch effects. Regression coefficients are given for intakes in grams per day and multiplied by 1000. The full results are presented in the Supporting Information.

Abbreviations: Chr., chromosome; Est., regression coefficients from linear mixed regression models (MCCS); EWAS, methylome-wide association study; MCCS, Melbourne Collaborative Cohort Study.

^aReversibility coefficient, expressed as a percentage defined as coefficient ("former" compared with "current")/coefficient ("never" compared with "current").

TABLE 6 Longitudinal associations ($P < 10^{-5}$) assessed in the MCCS and in KORA (data collected 11 and 7 years apart, respectively)

| CpG | Chr | Position | Gene | Location | Cross-sectional Data | | Longitudinal Data | | | | | |
|------------|-----|-----------|--------------|----------|----------------------|---------------------|-------------------|--------------------|---------------------|---------------------|------|---------------------|
| | | | | | MCCS (previous week) | | MCCS | KORA | Pooled ^a | | | |
| | | | | | Est. | <i>P</i> | Est. | <i>P</i> | Est. | <i>P</i> | Est. | <i>P</i> |
| cg06690548 | 4 | 139162808 | SLC7A11 | Body | -7.9 | 1×10^{-86} | -4.7 | 2×10^{-5} | -5.4 | 7×10^{-11} | -5.2 | 5×10^{-15} |
| cg18120259 | 6 | 43894639 | LOC100132354 | Body | -2.4 | 4×10^{-27} | -1.5 | 5×10^{-3} | -1.5 | 5×10^{-8} | -1.5 | 8×10^{-10} |
| cg02711608 | 19 | 47287964 | SLC1A5 | 1stExon | -2.6 | 6×10^{-30} | -2.0 | 4×10^{-4} | -1.4 | 1×10^{-6} | -1.5 | 3×10^{-9} |
| cg14476101 | 1 | 120255992 | PHGDH | Body | -4.0 | 2×10^{-39} | -2.9 | 1×10^{-4} | -1.8 | 3×10^{-6} | -2.1 | 3×10^{-9} |
| cg16246545 | 1 | 120255941 | PHGDH | Body | -2.2 | 4×10^{-23} | -1.7 | 2×10^{-3} | -1.4 | 7×10^{-7} | -1.4 | 4×10^{-9} |
| cg11376147 | 11 | 57261198 | SLC43A1 | Body | -2.1 | 2×10^{-29} | -1.0 | 8×10^{-3} | -1.4 | 1×10^{-6} | -1.3 | 3×10^{-8} |
| cg20732160 | 3 | 48590040 | PFKFB4 | Body | -1.7 | 3×10^{-16} | -1.6 | 3×10^{-3} | -1.2 | 1×10^{-5} | -1.3 | 1×10^{-7} |
| cg03068497 | 7 | 30635838 | GARS | Body | -3.0 | 3×10^{-14} | -2.0 | 2×10^{-2} | -2.4 | 4×10^{-6} | -2.3 | 2×10^{-7} |
| cg07626482 | 19 | 47289503 | SLC1A5 | TSS1500 | -1.4 | 4×10^{-20} | -1.2 | 2×10^{-3} | -1.0 | 4×10^{-5} | -1.1 | 3×10^{-7} |
| cg13526915 | 14 | 24164078 | | | -2.1 | 8×10^{-13} | -1.2 | 8×10^{-2} | -1.7 | 1×10^{-5} | -1.6 | 3×10^{-6} |
| cg14756878 | 2 | 12568736 | | | -1.2 | 3×10^{-9} | -1.4 | 2×10^{-3} | -0.9 | 4×10^{-4} | -1.0 | 3×10^{-6} |
| cg04460609 | 4 | 16532808 | LDB2 | Body | -2.2 | 3×10^{-19} | -1.2 | 2×10^{-2} | -1.3 | 1×10^{-4} | -1.3 | 5×10^{-6} |
| cg21626848 | 17 | 39969267 | SC65 | TSS1500 | -1.6 | 1×10^{-16} | -1.8 | 2×10^{-4} | -0.8 | 2×10^{-3} | -1.1 | 8×10^{-6} |
| cg03533472 | 16 | 46919112 | GPT2 | Body | -2.2 | 1×10^{-15} | -2.5 | 9×10^{-4} | -1.7 | 2×10^{-3} | -2.0 | 9×10^{-6} |

Note. All models were adjusted by fitting fixed effects for age, sex, smoking status, BMI, country of birth, sample type and white blood cell composition (percentage of CD4+ T cells, CD8+ T cells, B cells, NK cells, monocytes, and granulocytes, estimated using the Houseman algorithm), and batch effects. Regression coefficients are given for intakes in grams per day and multiplied by 1000. The full results are presented in the Supporting Information.

Abbreviations: Chr., chromosome; Est., regression coefficients from linear mixed regression models (MCCS) or linear regression models (KORA); KORA, Cooperative Health Research in the Augsburg Region; MCCS, Melbourne Collaborative Cohort Study.

^aEst. and *P* refer to the pooling of results from the MCCS and KORA using fixed-effects meta-analysis.

identical at baseline (alcohol in the previous week) and follow-up (alcohol in the previous year). These findings taken together indicate a substantial degree of reversibility in the associations, which was not assessed by previous studies.

Another potential limitation of our study is residual confounding, most notably by smoking or white blood cell-type composition, which are both strongly associated with alcohol drinking and DNA methylation.⁸ We observed that many CpG sites associated with alcohol drinking were also associated with other factors such as smoking, white blood cell composition, BMI, and other factors, which may indicate that these loci are very sensitive to the environment. Cell composition was estimated with the widely used Houseman algorithm modified by Jaffe and Irizarry,^{27,40} and we did not assess sensitivity to the method used for deriving cell composition.⁴¹ Although our adjustment for smoking was relatively comprehensive, our sensitivity analyses demonstrate that alcohol and smoking may exert joint influences on many CpGs across the genome.

We observed less substantial replication for CpGs discovered in individuals of African ancestry, which may indicate that alcohol-associated methylation changes are not generalizable to all human populations. Associations in individuals of African ancestry in the study by Liu et al were discovered by pooling data from a smaller number of studies so might have been less replicable by nature.

The associations between alcohol consumption and methylation were stronger for people with genetic predisposition to consume less alcohol, as defined by a 13-SNP polygenic score. Several genes

included in the polygenic score are involved in alcohol metabolism, for example, those of the alcohol dehydrogenase family (*ADH1B*, *ADH1C*, and *ADH5*) and *GCKR* (glucokinase regulatory protein, involved in glucose metabolism), which could provide a biological explanation for the interaction between the polygenic score and alcohol intake, given links between alcohol metabolism pathway and epigenetic mechanisms.⁴²

Associations appeared weaker for smokers and men, consistent with the observation that these population subgroups tend to drink more alcohol in most cultures, perhaps because of being less susceptible to the harmful effects of alcohol.⁴³ Women have previously been reported to have slower alcohol metabolism than men.⁴⁴ These findings should be confirmed by further studies. We included in the analysis participants who later developed cancer, which could give rise to collider bias when both DNA methylation and alcohol are associated with cancer risk.⁴⁵ We found no evidence of differences in associations by case/control status in our study. Further, that most discovered associations were replicated in independent cohorts of healthy participants with distinct ethnic origin is a strong testament that our findings were not driven by the inclusion of future cancer cases.

Our study is difficult to compare with studies other than that by Liu et al, because of differences in terms of study design, measurement of DNA methylation, and tissue analysed. Auta et al⁴⁶ measured global methylation levels in liver and cerebellum tissue of rats and concluded that each tissue was characterised by a distinct methylation signature of alcohol consumption, but no results at specific genetic

loci were reported. In a US study including over 600 samples, the association between alcohol dependence and epigenetic aging in various sets of blood, liver, and prefrontal cortex tissue was investigated.⁴⁷ The authors found some evidence of accelerated epigenetic aging although results were only partly consistent across tissues; significant associations were found in one set of blood samples and in liver. In a previous study using MCCS data, we also found accelerated epigenetic aging in the blood of heavy drinkers.⁴⁸ Lohoff and colleagues analysed data from multiple cohorts and tissues and found that alcohol use disorder was associated with differential methylation at the promoter region of *PCSK9*.⁴⁹ Interestingly, the association was observed in blood, brain, and liver, and *PCSK9* is a gene primarily expressed in the liver. However, *PCSK9* was not differentially methylated with respect to alcohol consumption in our study or that of Liu et al.

Another example of general lack of replication is a cross-sectional analysis of DNA methylation measures (Illumina 27 K assay) in the blood of 81 drinkers and 81 age- and sex-matched controls, which reported 906 differentially methylated CpG sites.⁵⁰ However, these findings were not independently replicated, and as reported by the authors, only two of these were identified at $P < 10^{-7}$ in the study by Liu et al¹⁹; these were also identified in our study. Tay and colleagues studied 1287 adolescents for whom blood DNA methylation was measured using the HM450K assay⁵¹ to identify associations with psychosocial stress and substance use, including alcohol consumption. There was weak evidence of an association between alcohol use and methylation of the *SPDEF* gene, the latter being associated with number of stressful life events. The *SPDEF* gene was not associated with alcohol consumption in our study.

The review by Zhang and Gelernter identified a number of small studies ($N < 200$ participants) that examined associations of DNA methylation at candidate genes with various alcohol use or disorder outcomes.¹⁸ Even for the largest of these studies, associations were generally weak and not replicated in our study, for example, reported associations for the *MAOA* gene,⁵² for *OPRM1*,⁵³ for *POMC*⁵⁴ at which an association in the opposite direction was observed in our study (cg09916783 promoter of *POMC*: $P = 1.9 \times 10^{-6}$ for last week intake). Several studies investigated the *SLC6A3/4* genes for which there was very weak evidence of association in our study ($P > 4 \times 10^{-4}$). A study of 383 heavy drinkers whose DNA methylation was measured in saliva samples using a custom assay targeting six CpG sites in the promoter region of the dopamine D2 receptor (*DRD2*) gene reported associations of DNA methylation with activation of neural reward regions, as well as severity alcohol use disorder symptoms.⁵⁵ In our setting, we did not observe associations between *DRD2* DNA methylation and alcohol consumption, and only marginally significant associations with other *DRD* genes. Finally, Bruckmann et al observed hypomethylation of *GDAP1* in alcohol-dependent patients, with a dose-response relationship by clinical severity⁵⁶; interestingly, we also observed hypomethylation of *GDAP1* promoter at cg23779890 (last week intake, $P = 3.1 \times 10^{-9}$).

The newly discovered CpG sites with strongest evidence of association with alcohol consumption were all located in genes, including in

regulatory regions. The five most strongly differentially methylated newly identified CpGs located in regulatory regions were in *SFRS13A* (Serine and Arginine Rich Splicing Factor 10, a splicing factor that in its dephosphorylated form acts as a general repressor of pre-mRNA splicing), *CPNE1* (encoding Copine 1, a calcium-dependent phospholipid-binding protein that plays a role in calcium-mediated intracellular processes), *SLC1A5* (Solute Carrier Family 1 Member 5, encoding a sodium-dependent neutral amino acid transporter), *FAM49A* (Family with sequence similarity 49, member A), and *PRELP* (Proline And Arginine Rich End Leucine Rich Repeat Protein, which encodes a leucine-rich repeat protein present in connective tissue extracellular matrix). We searched the literature for these genes and found little evidence of obvious links to alcohol metabolism, consumption, or alcohol-related diseases. We did not examine causality in our study; we hypothesise that if DNA methylation were the cause of alcohol drinking, it would likely be at a restricted number of loci involved in addiction mechanisms and alcohol metabolism. Our investigation of biological pathways using two enrichment databases did not clearly identify pathways or mechanisms known to play a key role in alcohol metabolism or alcohol-related diseases such as cancer, cardiovascular disease, and mental health and addiction pathologies. Further analyses at the gene expression level, in particular for the most significant pathways we identified, are required to investigate to what extent the methylation changes associated with alcohol drinking identified in our study are involved in epigenetic regulation of genes or networks of genes. Our study identified methylation patterns in blood whose usefulness as biomarkers to predict or monitor alcohol-related disease should be validated in further studies.

Our study shows that alcohol consumption is associated with widespread changes in blood DNA methylation. These changes appear more pronounced in women, non-smokers, and individuals with lower genetic predisposition to drink alcohol and are at least partially reversible.

ACKNOWLEDGMENTS

This work (MCCS) was supported by the Australian National Health and Medical Research Council (NHMRC) (Grant 1088405). MCCS cohort recruitment was funded by VicHealth and Cancer Council Victoria. The MCCS was further supported by Australian NHMRC Grants 209057, 251553, and 504711 and by infrastructure provided by Cancer Council Victoria. Cases were ascertained through the Victorian Cancer Registry (VCR) and the Australian Cancer Database (Australian Institute of Health and Welfare). The nested case-control methylation studies were supported by the NHMRC Grants 1011618, 1026892, 1027505, 1050198, 1043616, and 1074383. M.C.S. is an NHMRC Senior Research Fellow (1061177). The KORA study was initiated and financed by the Helmholtz Zentrum München – German Research Center for Environmental Health, which is funded by the German Federal Ministry of Education and Research (BMBF) and by the State of Bavaria. Furthermore, KORA research has been supported within the Munich Center of Health Sciences (MC-Health), Ludwig-Maximilians-Universität, as part of LMUinnovativ. This work has received funding from the European Foundation for Alcohol Research (ERAB 2018 –

EA1817). We thank all members of field staffs who were involved in the planning and conduct of the MONICA/KORA Augsburg studies. The LOLIPOP study is supported by the National Institute for Health Research (NIHR) Comprehensive Biomedical Research Centre Imperial College Healthcare NHS Trust, the British Heart Foundation (SP/04/002), the Medical Research Council (G0601966, G0700931), the Wellcome Trust (084723/Z/08/Z, 090532, and 098381), the NIHR (RP-PG-0407-10371), the NIHR Official Development Assistance (ODA, award 16/136/68), the European Union FP7 (EpiMigrant, 279143), and H2020 programs (iHealth-T2D, 643774). We acknowledge support of the MRC-PHE Centre for Environment and Health and the NIHR Health Protection Research Unit on Health Impact of Environmental Hazards. The work was carried out in part at the NIHR/Wellcome Trust Imperial Clinical Research Facility. The views expressed are those of the author(s) and not necessarily those of the Imperial College Healthcare NHS Trust, the NHS, the NIHR, or the Department of Health. We thank the participants and research staff who made the study possible. JC is supported by the Singapore Ministry of Health's National Medical Research Council under its Singapore Translational Research Investigator (STaR) Award (NMRC/STaR/0028/2017).

AUTHOR CONTRIBUTIONS

PAD, JEJ, EM, DRE, GGG, and RLM were responsible for the study concept and design. CHJ, JEJ, EM, DFS, LB, GS, CG, KHL, AP, JSK, MCS, DRE, MW, JCC, GGG, and RLM contributed to the acquisition of data. HJ, XW, CHJ, and JEJ contributed to the preparation of the data. PAD, RW, and BL performed the statistical analyses. PAD and RLM drafted the manuscript. All authors provided critical revision of the manuscript for important intellectual content. All authors critically reviewed content and approved final version for publication.

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How to cite this article: Dugué P-A, Wilson R, Lehne B, et al. Alcohol consumption is associated with widespread changes in blood DNA methylation: Analysis of cross-sectional and longitudinal data. *Addiction Biology*. 2019;e12855. <https://doi.org/10.1111/adb.12855>