

Anthropometric risk factors for ovarian cancer in the NIH-AARP Diet and Health Study

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ABSTRACT

Background: Identifying potentially modifiable risk factors for ovarian cancer is essential for prevention because this cancer is predominantly detected at a late stage. Here we estimated the relations of general adiposity and measures reflecting body fat distribution to the risk of epithelial ovarian cancer.

Methods: We ascertained 683 ovarian epithelial cancers (344 high-grade serous) among 145,575 women, aged 50 to 72 years (median follow-up 12.6 years), from the National Institutes of Health - American Association of Retired Persons (NIH-AARP) Diet and Health Study. Using Cox models, we estimated confounder-adjusted hazard ratios (HRs) and 95% confidence intervals (CI) for associations of overall ovarian cancer and high-grade serous carcinoma with body mass index, waist circumference, hip circumference, waist-hip ratio, waist-height ratio, body adiposity index, body shape index, and abdominal volume index.

Results: Anthropometric measures were unrelated to overall ovarian cancer and high-grade serous ovarian cancer. For example, the HR for overall ovarian cancer per standard deviation increment of body mass index at baseline was 0.97 (95% CI: 0.87 to 1.09). Similar associations were observed with measurements of body fat distribution.

Conclusions: These results do not indicate that adult adiposity is associated with ovarian cancer risk in post-menopausal women.

Impact: These findings provide little evidence that lowering obesity prevents ovarian cancer.

Word count: abstract: 208, main text: 1,766

Introduction

Ovarian cancer is the fifth most common cause of cancer death in women in North America and western Europe (1). Meta-analyses and pooled observational studies suggest that obesity may be positively related to ovarian cancer risk (2-5). The World Cancer Research Fund/American Institute for Cancer Research concluded that the evidence for the link between obesity and increased ovarian cancer risk is probable (6), and an umbrella review graded the evidence as suggestive (3). A growing body of research further suggests that body composition plays an important role in site-specific cancer development (7-9). Previous studies indicate that anthropometric measures of general obesity and body fat distribution may be differentially related to ovarian cancer risk; however, results have been conflicting (4,5,10-12). We conducted a cohort study among post-menopausal women using data from the National Institutes of Health - American Association of Retired Persons (NIH-AARP) Diet and Health Study to provide further insights into the association between body fatness and subtype-specific epithelial ovarian cancer risk by comparing indicators of general obesity and body fat distribution.

Methods

Study Population

The NIH-AARP Diet and Health Study is a prospective cohort study of persons in the U.S. (13). At baseline (1995-1996), 3.5 million AARP members aged 50 to 71 years who resided in six states (California, Florida, Louisiana, New Jersey, North Carolina, and Pennsylvania) and two metropolitan areas (Atlanta, Georgia; and Detroit, Michigan) were invited to complete a questionnaire on demographic, diet, and lifestyle characteristics. Questionnaires were completed satisfactorily by 566,398 participants. Six months after completing the baseline questionnaire, a second questionnaire was mailed to living participants who did not have a self-reported colon, breast, or prostate cancer at baseline to collect additional information. Self-reported weight and height were collected on the baseline questionnaire. Self-reported waist circumference and hip circumference were assessed on the second questionnaire. We

excluded male participants (n= 339,666); those with unknown cancer or previous diagnosis of cancer other than non-melanoma skin cancer at baseline before completion of the second questionnaire (n=16,300); those with bilateral oophorectomy and unknown oophorectomy status (n=57,047); those with no information on height, weight, waist circumference, or hip circumference (n=95,764); and subjects with body mass index less than 18.5 kg/m² or more than 65 kg/m² (n=3,528). Our final analytical datasets included 145,575 women for the analysis on body mass index, 60,999 for waist circumference, 60,826 for hip circumference, 60,597 for waist-to-hip ratio, 60,999 for waist-to-height ratio, 60,826 for body adiposity index, 60,999 for body shape index, and 60,597 for abdominal adiposity index. The Special Studies Institutional Review Board of the U.S. National Cancer Institute approved the study (13). All participants gave informed consent by virtue of completing and returning the baseline questionnaire.

Assessment of anthropometric measures

Self-reported height and weight were obtained from the baseline questionnaire. Body mass index was calculated as weight (kg) divided by the square of height (in meters). In the second questionnaire, participants were instructed to measure their waist circumference and hip circumference using a tape measure to the nearest 0.25 inch while standing. Waist circumference was to be measured 1 inch above the navel if this was not the waistline. Hip circumference was defined as the largest circumference between the upper edge of the pelvis and the femur. Waist-hip ratio was calculated by dividing waist circumference (cm) by hip circumference (cm), and waist-height ratio was calculated as waist circumference (cm) / height (cm). The body adiposity index was calculated as hip circumference (cm)/height^{1.5}(m)-18 (14). The body shape index was based on waist circumference (cm)/BMI^{2/3}×height²(m) (15). The abdominal volume index was also quantified:

$$2\text{cm} \times \text{waist circumference}^2(\text{cm}) + 0.7\text{cm} \times \left(\text{waist circumference}(\text{cm}) - \text{hip circumference}(\text{cm}) \right)^2 / 1000$$

(16).

Definition of cancer outcomes

Diagnoses of ovarian cancer were ascertained through September 31, 2011, via linkage to state cancer registries of the eight recruitment areas where the study participants were most likely to relocate: Arizona, Nevada, and Texas. This approach has been estimated to yield a sensitivity of 90% and a specificity of nearly 100% (17). Newly diagnosed ovarian cancer cases were identified using the International Classification of Diseases for Oncology (ICD-O-3) topography (C56.9) and morphology codes (8441, 8460, 8461, 8470, 8471, 8480, 8481, 8380, 8381, 8560, 8570, 8310, 8313, 8010, 8020, 8021, 8050, 8070, 8120, 8140, 8255, 8260, 8323, 8440, 8450, 8562, 9000) (18). The high-grade serous group included all invasive serous cancers except low-grade (8461/3) (19).

Baseline confounders

We controlled for several baseline participant characteristics that were assumed to cause adiposity or ovarian cancer (3,5,20-22). We assumed that direct causes of the exposure or outcome, excluding possible instrumental variables, would identify a sufficient set of confounders (23). Potential confounding variables included age, education (no school degree/unknown, primary school, technical school/secondary, university), participants' self-reported information on race/ethnicity (none-Hispanic white, other), cigarette smoking (never, current <15 cigarettes per day, current \geq 15 cigarettes/day, former < 10 years, former \geq 10 years), alcohol consumption (in grams of pure alcohol per day), parity (0, 1, \geq 2 children), age at menarche (\leq 12 years, > 12 years), family history of ovarian cancer, oral contraceptive use, and menopausal hormone therapy.

Statistical Analysis

Cox proportional hazards regression was used for estimating adjusted hazard ratios (HRs) and 95% confidence intervals (CIs) for overall ovarian cancer and ovarian cancer subtypes. Time of study entry was age at baseline (or second questionnaire) and exit time was age at cancer diagnosis or the last date at which follow-up was considered complete. After confirming that the linearity assumption was met by testing cubic spline transformations (24), HRs were estimated

per standard deviation (1-SD) increase in anthropometric measures. Models were stratified by five-year age groups to minimize departure from proportionality and adjusted for education, race, smoking, alcohol consumption, parity, age at menarche, family history of ovarian cancer, oral contraceptive use, and menopausal hormone therapy.

The improvement in predictive accuracy after adding anthropometric measures to a null model (including race, age, education, smoking, alcohol consumption, parity, family history of ovarian cancer, oral contraceptive use and menopausal hormone therapy) was evaluated in terms of explained variation (R^2) (25), the Bayes Information Criterion (BIC), and model discrimination using Harrel's C index (24) derived from flexible parametric models (26). *P* values for the difference between Harrel's C indices of models with and without anthropometric indicators were computed using the method proposed by Antolini et al. (27). We used 1,000 bootstrap replications to perform internal validations and to correct R^2 , BIC, and Harrel's C indices for optimism (24).

In a sensitivity analysis, we used regression calibration for self-reported body mass index, waist circumference, and hip circumference to assess possible regression dilution bias (28). Because replicate measurements were not available, we applied published reliability coefficients (29-32), ranging from 0.5 to 0.9. A further threat to the validity of our estimates is potential unobserved confounding by undiagnosed ovarian cancer (often referred to as "reverse causation" (33)) if these conditions are symptomatic enough to induce a change in body weight. We, therefore, assumed a three year minimum latent period required for weight change due to unobserved disease to affect the outcome and excluded events that occurred during this time (34). The statistical analysis was performed using Stata 15.1.

Results

In the analytical sample of 145,575 women, the mean (SD) age at baseline was 61.8 (5.4) years. During a median follow-up time of 12.6 years, participants contributed 1,897,323 person-years and 683 ovarian cancer (343 high-grade serous) cases occurred. The baseline characteristics of the analytical are provided in Table 1.

Neither body mass index nor other anthropometric measurements were associated with the risk of ovarian cancer (Table 2). For example, HRs for overall ovarian cancer per 1-SD increment in body mass index, waist circumference, hip circumference, waist-hip ratio, waist-height-ratio were 0.97 (95% CI: 0.87-1.09), 1.07 (95% CI: 0.91-1.25), 1.05 (95% CI: 0.90-1.24), 1.03 (95% CI: 0.88-1.20), 1.07 (95% CI: 0.91-1.26), respectively. No associations were observed for high-grade serous carcinomas. The accuracy of models predicting ovarian cancer risk was not improved after adding anthropometric measures (Table 3).

Sensitivity analysis indicated that HRs could have been attenuated towards the null because of measurement error in self-reported anthropometric measurements. For example, the unadjusted HR for waist circumference and overall ovarian cancer was 1.07 (95% CI: 0.91-1.25), but after accounting for potential regression dilution bias, assuming an attenuation factor of 0.7, the HR was 1.09 (95% CI: 0.90-1.29) (Supplementary Table 1). The associations were virtually unchanged when events occurring during the first three years of follow-up were excluded (Supplementary Table 2).

Discussion

This study examined the association of indicators of general obesity and body fat distribution with ovarian cancer risk using a large U.S. prospective cohort study of post-menopausal women. We found no association of anthropometric measures with the risk of overall ovarian cancer. Overall, our analysis does not support the hypothesis that central adiposity or measures of body fat distribution improve the prediction of ovarian cancer risk.

A larger body of research examined the association between obesity and ovarian cancer risk. However, the findings of more than 30 epidemiologic studies have been weak and mixed (35). Several meta-analyses and pooled analyses reported weak positive associations between adult body mass index and ovarian cancer risk, noting substantial between-study heterogeneity with weaker associations in prospective than case-control studies (2,6,36). A systematic review from the World Cancer Research Fund and the American Institute for Cancer Research included 28 prospective studies on ovarian cancer and calculated a summary relative risk for a 5 unit increment in body mass index of 1.07 (95% CI: 1.03-1.11) (2,6). The effect size was similar in

the post-menopausal group but it was less precise (relative risk per 5 kg m⁻² = 1.07; 95% CI: 1.00-1.14). Results from the 2013 Ovarian Cancer Association Consortium (4) pooled analysis of case-control studies found that the positive association with body mass index was stronger in pre-menopausal women. The heterogeneity of findings reported here and previously could be explained by menopausal status and higher statistical efficacy of meta-analysis. Few studies have examined how different measures of body fat distribution are related to ovarian cancer and its subtypes (10-12). Existing cohort studies found no association of waist circumference and waist-hip ratio and ovarian cancer risk (2,6,10). Similar to previous studies (4,5), the present study found no notable differences between histotypes.

The present study has several limitations. It relied on self-reported anthropometric data and potential measurement error could have attenuated the observed associations. Our study also lacked updated information on anthropometric measurements during follow-up. Another drawback is the low number of cases by ovarian cancer subtype and a lack of statistical power to test for effect modification.

In summary, results from this prospective study of post-menopausal women does not support associations between measures of central obesity and body fat distribution and risk of ovarian cancer.

Disclosure of Potential Conflicts of Interest

No potential conflicts of interest were disclosed.

Author's Contribution

Conception and design: Sebastian E. Baumeister, Inga Schlecht, Michael F. Leitzmann

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References

1. Wild CP, Weiderpass E, Stewart BW. World cancer report 2020. Lyon, France: International Agency for Research on Cancer; 2020.
2. Aune D, Navarro Rosenblatt DA, Chan DS, Abar L, Vingeliene S, Vieira AR, *et al.* Anthropometric factors and ovarian cancer risk: a systematic review and nonlinear dose-response meta-analysis of prospective studies. *International journal of cancer* **2015**;136(8):1888-98 doi 10.1002/ijc.29207.
3. Kalliala I, Markozannes G, Gunter MJ, Paraskevidis E, Gabra H, Mitra A, *et al.* Obesity and gynaecological and obstetric conditions: umbrella review of the literature. *bmj* **2017**;359:j4511.
4. Olsen CM, Nagle CM, Whiteman DC, Ness R, Pearce CL, Pike MC, *et al.* Obesity and risk of ovarian cancer subtypes: evidence from the Ovarian Cancer Association Consortium. *Endocrine-related cancer* **2013**;20(2):251-62 doi 10.1530/erc-12-0395.
5. Wentzensen N, Poole EM, Trabert B, White E, Arslan AA, Patel AV, *et al.* Ovarian Cancer Risk Factors by Histologic Subtype: An Analysis From the Ovarian Cancer Cohort Consortium. *Journal of clinical oncology : official journal of the American Society of Clinical Oncology* **2016**;34(24):2888-98 doi 10.1200/jco.2016.66.8178.
6. World Cancer Research Fund International, American Institute for Cancer Research. Diet, nutrition, physical activity and cancer: a global perspective. third expert report. 2018.
7. Harding JL, Shaw JE, Anstey KJ, Adams R, Balkau B, Brennan-Olsen SL, *et al.* Comparison of anthropometric measures as predictors of cancer incidence: A pooled collaborative analysis of 11 Australian cohorts. *International journal of cancer* **2015**;137(7):1699-708 doi 10.1002/ijc.29529.
8. Freisling H, Arnold M, Soerjomataram I, O'Doherty MG, Ordonez-Mena JM, Bamia C, *et al.* Comparison of general obesity and measures of body fat distribution in older adults in relation to cancer risk: meta-analysis of individual participant data of seven prospective cohorts in Europe. *Br J Cancer* **2017**;116(11):1486-97 doi 10.1038/bjc.2017.106.
9. Arnold M, Leitzmann M, Freisling H, Bray F, Romieu I, Renehan A, *et al.* Obesity and cancer: An update of the global impact. *Cancer epidemiology* **2016**;41:8-15 doi 10.1016/j.canep.2016.01.003.
10. Lahmann PH, Cust AE, Friedenreich CM, Schulz M, Lukanova A, Kaaks R, *et al.* Anthropometric measures and epithelial ovarian cancer risk in the European Prospective Investigation into Cancer and Nutrition. *International journal of cancer* **2010**;126(10):2404-15 doi 10.1002/ijc.24952.
11. Ma X, Beeghly-Fadiel A, Shu XO, Li H, Yang G, Gao YT, *et al.* Anthropometric measures and epithelial ovarian cancer risk among Chinese women: results from the Shanghai Women's Health Study. *Br J Cancer* **2013**;109(3):751-5 doi 10.1038/bjc.2013.384.
12. Minlikeeva AN, Moysich KB, Mayor PC, Etter JL, Cannioto RA, Ness RB, *et al.* Anthropometric characteristics and ovarian cancer risk and survival. *Cancer causes & control : CCC* **2018**;29(2):201-12 doi 10.1007/s10552-017-0997-5.

13. Schatzkin A, Subar AF, Thompson FE, Harlan LC, Tangrea J, Hollenbeck AR, *et al.* Design and serendipity in establishing a large cohort with wide dietary intake distributions : the National Institutes of Health-American Association of Retired Persons Diet and Health Study. *American journal of epidemiology* **2001**;154(12):1119-25 doi 10.1093/aje/154.12.1119.
14. Bergman RN, Stefanovski D, Buchanan TA, Sumner AE, Reynolds JC, Sebring NG, *et al.* A better index of body adiposity. *Obesity* **2011**;19(5):1083-9.
15. Krakauer NY, Krakauer JC. A new body shape index predicts mortality hazard independently of body mass index. *PLoS one* **2012**;7(7):e39504 doi 10.1371/journal.pone.0039504.
16. Wu F, Ho V, Fraser BJ, Schmidt MD, Dwyer T, Venn AJ, *et al.* Predictive utility of childhood anthropometric measures on adult glucose homeostasis measures: a 20-year cohort study. *International journal of obesity (2005)* **2018**;42(10):1762-70 doi 10.1038/s41366-018-0177-z.
17. Michaud D, Midthune D, Hermansen S, Leitzmann M, Harlan L, Kipnis V, *et al.* Comparison of cancer registry case ascertainment with SEER estimates and self-reporting in a subset of the NIH-AARP Diet and Health Study. *J Registry Manag* **2005**;32(2):70-5.
18. Fritz A, Percy C, Jack A, Shanmugaratnam K, Sobin L, Parkin DM, *et al.* International classification of diseases for oncology: ICD-O, 3rd ed. Geneva: World Health Organization **2000**.
19. Carcangiu M, Kurman RJ, Carcangiu ML, Herrington CS. WHO classification of tumours of female reproductive organs. International Agency for Research on Cancer; 2014.
20. Kossai M, Leary A, Scoazec J-Y, Genestie C. Ovarian Cancer: A Heterogeneous Disease. *Pathobiology* **2018**;85(1-2):41-9.
21. Matulonis UA, Sood AK, Fallowfield L, Howitt BE, Sehoul J, Karlan BY. Ovarian cancer. *Nature reviews Disease primers* **2016**;2:16061 doi 10.1038/nrdp.2016.61.
22. Reid BM, Permuth JB, Sellers TA. Epidemiology of ovarian cancer: a review. *Cancer biology & medicine* **2017**;14(1):9-32 doi 10.20892/j.issn.2095-3941.2016.0084.
23. VanderWeele TJ. Principles of confounder selection. *European journal of epidemiology* **2019**;34(3):211-9 doi 10.1007/s10654-019-00494-6.
24. Harrell F. Regression modeling strategies: with applications to linear models, logistic and ordinal regression, and survival analysis. Springer; 2015.
25. O'Quigley J, Xu R, Stare J. Explained randomness in proportional hazards models. *Statistics in medicine* **2005**;24(3):479-89 doi 10.1002/sim.1946.
26. Royston P, Parmar MK. Flexible parametric proportional-hazards and proportional-odds models for censored survival data, with application to prognostic modelling and estimation of treatment effects. *Statistics in medicine* **2002**;21(15):2175-97 doi 10.1002/sim.1203.
27. Antolini L, Nam B-H, D'Agostino RB. Inference on correlated discrimination measures in survival analysis: a nonparametric approach. *Communications in statistics-Theory and Methods* **2004**;33(9):2117-35.
28. Keogh RH, White IR. A toolkit for measurement error correction, with a focus on nutritional epidemiology. *Statistics in medicine* **2014**;33(12):2137-55 doi 10.1002/sim.6095.
29. Maukonen M, Männistö S, Tolonen H. A comparison of measured versus self-reported anthropometrics for assessing obesity in adults: a literature review. *Scandinavian journal of public health* **2018**;46(5):565-79.
30. McAdams MA, Van Dam RM, Hu FB. Comparison of self-reported and measured BMI as correlates of disease markers in US adults. *Obesity (Silver Spring, Md)* **2007**;15(1):188-96 doi 10.1038/oby.2007.504.
31. Rimm EB, Stampfer MJ, Colditz GA, Chute CG, Litin LB, Willett WC. Validity of self-reported waist and hip circumferences in men and women. *Epidemiology (Cambridge, Mass)* **1990**;1(6):466-73 doi 10.1097/00001648-199011000-00009.
32. McAdams MA, Van Dam RM, Hu FB. Comparison of self-reported and measured BMI as correlates of disease markers in US adults. *Obesity* **2007**;15(1):188-

33. Danaei G, Robins JM, Young JG, Hu FB, Manson JE, Hernan MA. Weight Loss and Coronary Heart Disease: Sensitivity Analysis for Unmeasured Confounding by Undiagnosed Disease. *Epidemiology (Cambridge, Mass)* **2016**;27(2):302-10 doi 10.1097/ede.0000000000000428.
34. Flegal KM, Graubard BI, Yi SW. Comparative effects of the restriction method in two large observational studies of body mass index and mortality among adults. *European journal of clinical investigation* **2017**;47(6):415-21.
35. Jochem C, Schlecht I, Leitzmann M. Epidemiologic Relationship Between Obesity and Ovarian Cancer. *Focus on Gynecologic Malignancies*: Springer; 2018. p 21-30.
36. Kyrgiou M, Kalliala I, Markozannes G, Gunter MJ, Paraskeva E, Gabra H, *et al.* Adiposity and cancer at major anatomical sites: umbrella review of the literature. *Bmj* **2017**;356:j477 doi 10.1136/bmj.j477.

Table 1 Baseline characteristics of ovarian cancer cases among 145,575 women in the NIH-AARP Study

	<i>All study subjects</i>	No ovarian cancer	Ovarian cancer	High-grade serous
Age (years)	62.2 (57.4, 66.4)	62.2 (57.4, 66.4)	63.8 (58.1, 67.4)	63.0 (57.6, 67.4)
Education (%)				
<12 yrs	5.8	8.6	5.3	3.1
12 yrs	25.7	25.7	24.9	24.3
>12 yrs	68.5	68.5	69.8	72.6
Race (%)				
Non-hispanic white	91.2	91.2	95.0	96.2
Non-hispanic black	8.1	5.3	2.6	1.1
Hispanic/other	5.3	3.5	2.4	2.7
Body mass index	25.8 (22.9, 29.3)	25.7 (22.9, 29.3)	25.6 (22.5, 29.7)	25.1 (22.5, 29.0)
Waist circumference (cm)	82.6 (74.9, 91.4)	82.6 (74.9, 91.4)	83.3 (74.3, 93.9)	82.6 (74.3, 91.4)
Hip circumference (cm)	101.6 (96.5, 109.2)	101.6 (96.5, 109.2)	101.6 (96.5, 112.4)	101.6 (96.5, 109.2)
Waist-hip ratio	0.81 (0.75, 0.86)	0.81 (0.76, 0.86)	0.81 (0.76, 0.87)	0.81 (0.76, 0.86)
Waist -height ratio	0.51 (0.46, 0.57)	0.51 (0.46, 0.57)	0.51 (0.46, 0.57)	0.51 (0.46, 0.56)
Body adiposity index	31.0 (28.0, 34.7)	31.0 (28.0, 35.0)	31.1 (28.3, 34.9)	31.4 (28.2, 34.5)
Body shape index	0.01 (0.01, 0.01)	0.01 (0.01, 0.01)	0.01 (0.01, 0.01)	0.01 (0.01, 0.01)
Abdominal volume index	13.9 (11.6, 17.2)	14.0 (11.6, 17.2)	14.5 (11.4, 17.8)	13.9 (11.4, 17.2)
Smoking status (%)				
Never smoked	45.3	45.3	49.4	54.6
Former smoker, ≤20 cigarettes per day	27.6	27.6	27.8	23.7
Former smoker, >20 cigarettes per day	13.7	13.7	11.3	10.5
Current smoker, ≤20 cigarettes per day	10.5	10.5	8.8	8.9
Alcohol consumption (grams/day)	1.0 (0.3, 4.9)	1.0 (0.3, 4.9)	1.0 (0.4, 5.3)	0.9 (0.1, 5.3)
Parity				
Never had a child	15.3	15.3	18.7	17.2

1 child	10.3	10.3	11.9	11.8
2 and more children	74.4	74.4	69.4	70.1
Age at menarche				
≤12 yrs	48.6	48.6	49.5	52.2
>12 yrs	51.4	51.4	50.5	47.9
Family history of ovarian cancer (%)	6.1	6.1	7.3	6.2
Ever oral contraceptive use (%)	40.6	40.6	33.2	31.6
Ever hormone replacement therapy (%)	46.2	46.2	50.2	54.3

NIH-AARP, NIH-AARP Diet and Health Study. Entries are percent values for categorical variables and medians (25th percentile, 75th percentile) for continuous variables.

Table 2 Association of general obesity and indicators of body fat distribution with ovarian cancer in NIH-AARP

	Ovarian Cancer Risk	High-grade serous
Body mass index (<i>n</i> =145,575),		
number of cases	683	343
HR per SD (95% CI)	0.97 (0.87; 1.09)	0.90 (0.76; 1.06)
Waist circumference		
(<i>n</i> =60,999), number of cases	295	151
HR per SD (95% CI)	1.07 (0.91; 1.25)	0.98 (0.77; 1.23)
Hip circumference (<i>n</i> =60,826),		
number of cases	295	150
HR per SD (95% CI)	1.05 (0.90; 1.24)	0.95 (0.75; 1.21)
Waist-hip ratio (<i>n</i> =60,597),		
number of cases	294	150
HR per SD (95% CI)	1.03 (0.88; 1.20)	1.01 (0.81; 1.27)
Waist-height ratio (<i>n</i> =60,999),		
number of cases	295	151
HR per SD (95% CI)	1.07 (0.91; 1.26)	0.97 (0.77; 1.23)
Body adiposity index		
(<i>n</i> =60,826), number of cases	294	150
HR per SD (95% CI)	1.06 (0.90; 1.25)	0.95 (0.75; 1.21)
Body shape index		
(<i>n</i> =60,999), number of cases	294	151
HR per SD (95% CI)	0.99 (0.84; 1.17)	1.10 (0.87; 1.39)
Abdominal volume index		
(<i>n</i> =60,597), number of cases	295	150
HR per SD (95% CI)	1.06 (0.90; 1.24)	0.95 (0.75; 1.21)

NIH-AARP, NIH-AARP Diet and Health Study. HR (hazard ratio) from age-group stratified multivariable Cox model adjusted for education, race, smoking, alcohol consumption, parity, age at menarche, family history of ovarian cancer, oral contraceptive use, and menopausal hormone therapy.

Table 3 General obesity and indicators of body fat distribution for prediction of ovarian cancer

	Baseline model	Body mass index	Waist circumference	Hip circumference	Waist-to-hip ratio	Waist-to-height ratio	Body adiposity index	Body shape index	Abdominal volume index
Ovarian cancer									
Adjusted R ²	0.060	0.059	0.071	0.067	0.066	0.071	0.068	0.069	0.067
BIC	16396.29	16408.11	6681.254	6659.244	6657.74	6681.033	6658.871	6682.416	6656.822
Harrel's C-index (P)	0.600	0.601 (0.601)	0.616 (0.517)	0.616 (0.423)	0.615 (0.572)	0.617 (0.465)	0.617 (0.320)	0.617 (0.419)	0.615 (0.519)
High-grade serous									
Adjusted R ²	0.090	0.092	0.109	0.103	0.103	0.109	0.103	0.112	0.103
BIC	8303.135	8313.223	3492.481	3471.319	3470.144	3492.480	3471.347	3491.962	3470.177
Harrel's C-index (P)	0.624	0.627 (0.146)	0.648 (0.255)	0.645 (0.287)	0.645 (0.296)	0.648 (0.255)	0.645 (0.281)	0.648 (0.256)	0.645 (0.276)

Null model included the predictors age, education, race, smoking, alcohol consumption, physical activity, parity, family history of ovarian cancer, family history of breast cancer, hormone therapy. Adjusted R²: explained variation. BIC: Bayes Information Criterion. P: P Value for difference of Harrel's C vs null model.