

Causal relationships between anthropometric traits, bone mineral density, osteoarthritis and spinal stenosis: A Mendelian randomisation investigation

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SUMMARY

Objective: Spinal stenosis is a common condition among older individuals, with significant morbidity attached. Little is known about its risk factors but degenerative conditions, such as osteoarthritis (OA) have been identified for their mechanistic role. This study aims to explore causal relationships between anthropometric risk factors, OA, and spinal stenosis using Mendelian randomisation (MR) techniques.

Design: We applied two-sample MR to investigate the causal relationships between genetic liability for select risk factors and spinal stenosis. Next, we examined the genetic relationship between OA and spinal stenosis with linkage disequilibrium score regression and Causal Analysis Using Summary Effect estimates MR method. Finally, we used multivariable MR (MVMR) to explore whether OA and body mass index (BMI) mediate the causal pathways identified.

Results: Our analysis revealed strong evidence for the effect of higher BMI (odds ratio [OR] = 1.54, 95%CI: 1.41–1.69, p-value = 2.7×10^{-21}), waist (OR = 1.43, 95%CI: 1.15–1.79, p-value = 1.5×10^{-3}) and hip (OR = 1.50, 95%CI: 1.27–1.78, p-value = 3.3×10^{-6}) circumference on spinal stenosis. Strong evidence of causality was also observed for higher bone mineral density (BMD): total body (OR = 1.21, 95%CI: 1.12–1.29, p-value = 1.6×10^{-7}), femoral neck (OR = 1.35, 95%CI: 1.09–1.37, p-value = 7.5×10^{-7}), and lumbar spine (OR = 1.38, 95%CI: 1.25–1.52, p-value = 4.4×10^{-11}). We detected high genetic correlations between spinal stenosis and OA (rg range: 0.47–0.66), with Causal Analysis Using Summary Effect estimates results supporting a causal effect of OA on spinal stenosis (OR_{alloA} = 1.6, 95%CI: 1.41–1.79). Direct effects of BMI, BMD on spinal stenosis remained after adjusting for OA in the MVMR.

Conclusions: Genetic susceptibility to anthropometric risk factors, particularly higher BMI and BMD can increase the risk of spinal stenosis, independent of OA status. These results may inform preventative strategies and treatments.

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Abbreviations: BMD, bone mineral density; BMI, body mass index; CI, confidence interval; OA, osteoarthritis; OR, odds ratio.

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Introduction

Spinal stenosis is a potentially debilitating condition with symptomatic spinal stenosis affecting about 10% of Western populations and prevalence only increasing with age.^{1–3} It is characterised by narrowing of the spinal canal that results in compression of the spinal cord and/or nerves, leading to symptoms such as back pain, sciatica and spinal claudication.⁴ Consequently, spinal stenosis often has a significant adverse impact on affected

individuals' quality of life.⁵ Although spinal stenosis can occur at any level of the spine, it most commonly affects the lumbar and cervical regions.⁶ Treatment can be conservative but ever increasing rates of surgery in the USA^{7,8} mean it now accounts for > \$15 billion⁹ per year in healthcare spending in the USA.

The two main, mutually non-exclusive causes of spinal stenosis are degenerative (acquired), and less commonly developmental (congenital).¹ Degenerative spinal stenosis is thought to be caused by changes associated with aging and spinal osteoarthritis (OA).¹⁰ For example, in UK Biobank (UKBB), a large prospective cohort in the UK, 50% of individuals with spinal stenosis diagnosis had a concurrent OA diagnosis.

Despite its increasing prevalence and the associated increasing healthcare cost, little is known about the epidemiology of spinal stenosis, in particular its modifiable risk factors.¹¹ It has been hypothesised that increased risk factor burden in the population could be responsible for this rise.¹² For example, there is observational evidence that high body mass index (BMI) predisposes to degenerative spinal disease, including spinal stenosis.^{13–15} Knutsson et al.¹² showed that obese construction workers had a twofold increased risk of lumbar spinal stenosis compared with normal weight workers. However, observational studies are liable to confounding and reverse causation making causal conclusions difficult. In addition, these studies are unable to assess whether raised BMI causes spinal stenosis through degenerative changes or other pathways. Previously, it has been shown that BMI is positively correlated with bone mineral density (BMD)^{16,17} and increased BMD has been associated with OA.¹⁸ It is therefore feasible that BMD and OA are mediating the relationship between BMI and spinal stenosis. An alternative explanation is that BMI is confounding the relationship between BMD and spinal stenosis.

Mendelian randomisation (MR) is an increasingly popular method for causal inference in epidemiology due to biobank-driven expansion in genome-wide association studies (GWAS) on a variety of phenotypes.¹⁹ MR utilises genetic variants that are randomly assigned at conception to explore causal relationships between exposures and outcomes. The technique capitalises on the Mendelian principles of inheritance, where segregation of genetic variation is independent of confounding factors and reverse causation, so MR is particularly useful when investigating risk factors that may be challenging to examine with conventional epidemiological methods. While not applied to spinal stenosis so far to the best of our knowledge, MR has previously confirmed a causal effect of BMI²⁰ and BMD^{20,21} on site-specific OA.

In this study, we employ two-sample univariable MR techniques to firstly explore the total causal relationships between genetic susceptibility to anthropometric risk factors and spinal stenosis. Among individual risk factors, we focus on measures of adiposity (BMI, waist circumference, hip circumference, waist-to-hip ratio) as overall BMI may not reflect body fat distribution and its effect on spine degeneration via mechanical and inflammatory pathways.²² We also look at height due to potentially increased mechanical stress in tall individuals,²³ bone mineral-related traits (BMD – total and lumbar, circulating calcium and phosphorus) due to importance in maintaining bone and joint health.^{24,25} Next, we employ a multivariable MR (MVMR) approach to elucidate the underlying independent mechanisms contributing to aetiology of spinal stenosis adjusting for effects of OA and BMD.

Methods

Genetic association studies

We used two (Table 1) publicly available spinal stenosis GWAS studies in European populations available from FinnGen release 8

(https://r8.risteys.finngen.fi/phenocode/M13_SPINSTENOSIS)²⁶ and UK Biobank (UKBB) available via PheWeb (<https://pheweb.org/UKB-TOPMed/pheno/720>)²⁷ with the spinal stenosis diagnosis defined as having been assigned the International Classification of Disease revision 10 (ICD-10) M48.0 (spinal stenosis) code. The FinnGen study was used in our main results due to increased power offered by its sample size (16,698 cases in FinnGen and 3713 in UKBB) whereas the UKBB GWAS is used as a sensitivity analysis. The reasons for reduced prevalence seen in UKBB can be potentially attributed to misclassification due to requirement for hospital inpatient admission for ICD-10 code assignment and lower MRI diagnosis rate relative to Finland (Ville Mattila, personal communication).

Since we were interested to study the effect of genetic liability for OA on spinal stenosis, we used the Genetics of Osteoarthritis (GO) European OA GWAS across 3 body sites and 2 composite phenotypes (hip, knee, knee/hip – i.e. knee and/or hip, spine, all – i.e. hip, knee, hand, finger, thumb and spine) in our main analyses (Table 1). To avoid bias induced by sample overlap between exposure and outcome in the analysis involving UKBB spinal stenosis GWAS we used custom GO GWAS with no UKBB individuals included.

We aimed to study the direct genetic effect of a number of anthropometric risk factors on spinal stenosis (Fig. 1): adiposity (BMI,^{28,29} hip circumference,³⁰ waist circumference,³⁰ waist-to-hip ratio³⁰), height,³¹ bone mineral density (BMD: total,³² lumbar spine³³ and femoral neck – this study) as well as circulating albumin-adjusted calcium (this study, Supplementary Methods and Supplementary Table 1) and circulating phosphate (Neale Lab GWAS available via OpenGWAS^{34,35}). Again, to prevent sample overlap in a subset of MR analyses we included additional BMI,²⁸ femoral neck BMD³³ GWAS with low number/no UKBB participants (but adjusted for weight).

Power calculations

We used the mRnd calculator (<https://shiny.cns.genomics.com/mRnd/>) to calculate the minimum detectable odds ratio (OR) at 80% power in our main two-sample MR analyses involving spinal stenosis as the outcome.

Linkage disequilibrium score regression

We utilised the LD Score (LDSC) ver 1.0.1 software³⁶ to estimate the genetic correlation (r_g) between OA and spinal stenosis, using the standard procedures described in the LDSC tutorial, using HapMap 3 single-nucleotide polymorphisms (SNPs) and 1000 Genomes European ancestry reference panel to calculate linkage disequilibrium (LD) scores.

Selection of genetic instruments

To identify genetic instruments for each exposure, we selected SNPs that showed strong association at a genome-wide significance level (p -value $< 5 \times 10^{-8}$). We further clumped the SNPs to ensure that LD as measured by $r^2 < 0.001$ between any pair of significant SNPs in a 10 Mbp window in the 1000 Genomes European panel³⁷ to avoid multiple instruments capturing the same causal effect. This was done using plink ver 1.9³⁸ as called by `ld_clump` function in the `ieugwasr` R package (<https://mrcieu.github.io/ieugwasr>). In each case, genetic variant associations for the outcome trait were extracted and harmonised using default settings in the `TwoSampleMR`³⁴ package. We next calculated the F-statistics and R^2 to check for weak instrument bias.

Two-sample MR analyses

We applied the two-sample MR approach, which utilises summary-level data from two non-overlapping GWAS, to estimate the

Exposure	GWAS source	Includes UKBB?	Sample size (cases/controls)	Number of SNP instruments*	R ² %	Mean F-statistic	Pubmed ID
<i>Anthropometric risk factors</i>							
Body mass index (BMI)	GIANT ³	N	339,224	78	1.56%	65.7	25673413
Body mass index (BMI)	GIANT ³	Y	681,275	496	4.91%	72.7	30124842
Hip circumference	GIANT ³	N	213,038	52	1.37%	55.0	25673413
Waist circumference	GIANT ³	N	232,101	42	1.09%	59.3	25673413
Waist-to-hip ratio	GIANT ³	N	212,244	29	0.68%	48.3	25673413
Height	GIANT ³	N	253,288	381	11.96%	78.2	25282103
Total body BMD ¹	GEFOS ⁴	Y	56,284	84	9.71%	65.2	30598549
Femoral neck BMD ¹	GEFOS ⁴	N	32,735	18	1.96%	54.6	26367794
Femoral neck BMD ¹	UKBB	Y	38,645	45	6.32%	54.4	This study
Lumbar spine BMD ¹	GEFOS ⁴	N	28,498	23	2.48%	48.2	26367794
Circulating calcium	UKBB	Y	361,194	233	11.85%	101.4	This study
Circulating phosphate	UKBB	Y	361,194	148	4.81%	93.2	Neale lab**
<i>Osteoarthritis</i>							
Hip OA ² with UKBB	GO Consortium	Y	36,445/316,943	40	0.53%	46.2	34450027
Hip OA ² no UKBB	GO Consortium	N	25,237/272,284	14	0.32%	37.7	34450027
KneeHip OA ² with UKBB	GO Consortium	Y	89,625/399,222	39	0.32%	40.1	34450027
KneeHip OA ² no UKBB	GO Consortium	N	60,683/282,999	8	0.09%	37.7	34450027
Knee OA ² with UKBB	GO Consortium	Y	62,603/332,423	31	0.31%	39.2	34450027
Knee OA ² no UKBB	GO Consortium	N	43,102/254,144	4	0.05%	40.2	34450027
All OA ² with UKBB	GO Consortium	Y	177,591/647,127	31	0.14%	37.4	34450027
All OA ² no UKBB	GO Consortium	N	108,970/399,281	4	0.05%	35.5	34450027
Spine OA ² with UKBB	GO Consortium	Y	27,916/303,489	1	0.01%	30.3	34450027
Spine OA ² no UKBB	GO Consortium	N	16,777/258,933	1	0.01%	30.3	34450027
<i>Spinal stenosis</i>							
Spinal stenosis	FinnGen	N	16,698/248,831	21	0.27%	40.8	36653562
Spinal stenosis	PheWeb UKBB	Y	3713/390,237	NA ⁵	NA ⁵	NA ⁵	32504056

*MR analyses using FinnGen spinal stenosis GWAS dataset as outcome.

**Available via OpenGWAS id: ukb-d-30810_irnt.

¹BMD - bone mineral density.

²OA - osteoarthritis.

³GIANT - The Genetic Investigation of ANthropometric Traits consortium.

⁴GEFOS - Genetic Factors for Osteoporosis.

⁵NA - not available.

Table 1

Osteoarthritis and Cartilage

GWAS used as sources for instrumental variables in the study.

causal effect of anthropometric risk factors on spinal stenosis using the TwoSampleMR³⁴ R package. We used the inverse-variance weighted (IVW) method as the primary analysis, where the causal estimate is obtained by combining the SNP-specific Wald ratios using a random-effects IVW meta-analysis. To combine the causal estimates obtained using FinnGen and UKBB spinal stenosis outcomes, we meta-analysed them with fixed-effects inverse variance method used for pooling in the R meta package. Effect estimates are interpretable as change in outcome per 1 standard deviation increase in continuous exposure or per doubling in the risk of binary exposure.

Sensitivity analyses

To assess the robustness of our findings and potential violation of MR assumptions, we performed several sensitivity analyses, including:

Weighted median and mode estimator

These approaches estimate the causal effect by calculating the median and mode of the individual Wald ratios, respectively, providing a consistent estimate if at least 50% of the weight comes from valid instruments (for median estimator) and the largest subset of variants identifies the same causal effect (for mode estimator).

MR-Egger regression

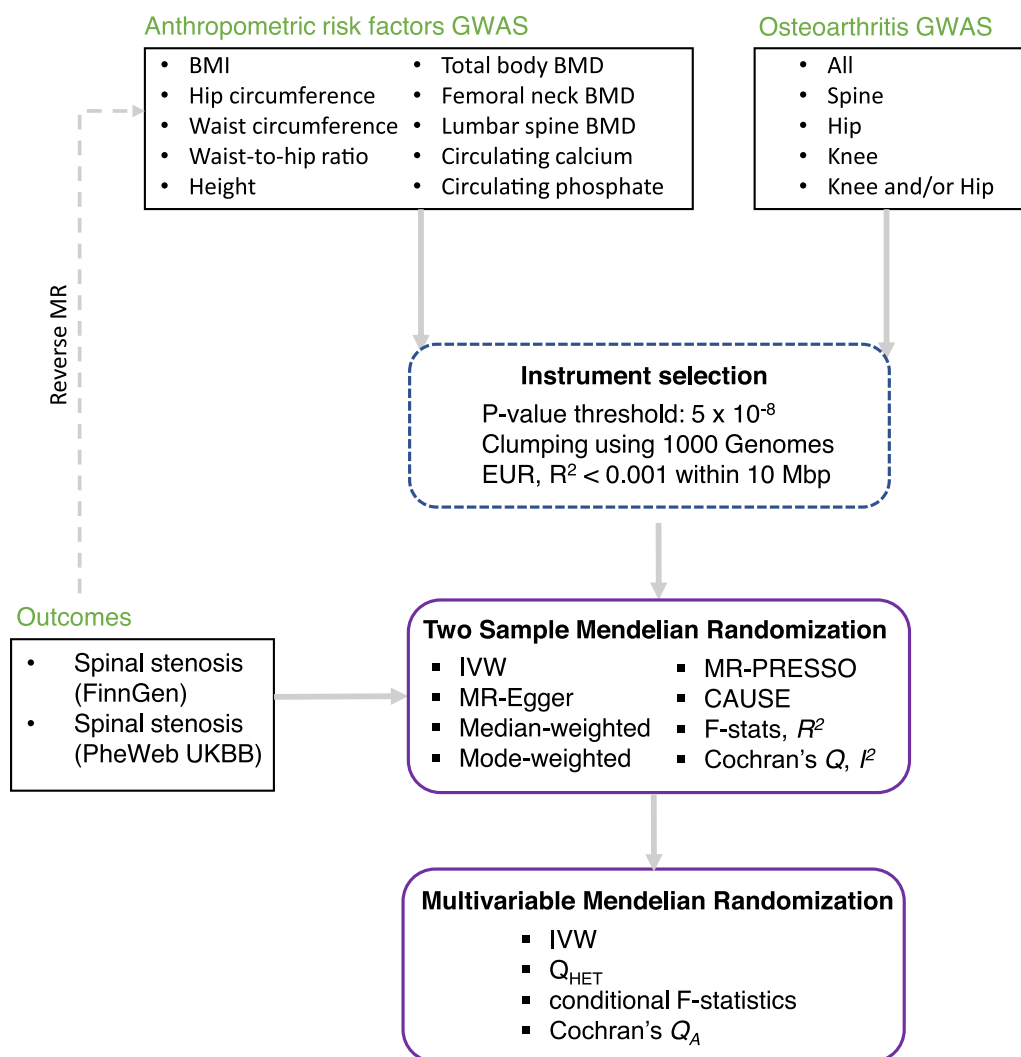
This method is robust to balanced pleiotropic effects, i.e. positive and negative effects of the instrument acting through alternative pathway cancelling each other out. It provides an estimate of the causal effect by regressing the SNP-outcome associations on the SNP-exposure associations, while allowing for an intercept term that captures the average pleiotropy across instruments.

MR-Pleiotropy RESidual Sum and Outlier³⁹

The MR-Pleiotropy RESidual Sum and Outlier (MR-PRESSO) test was used to detect and correct for horizontal pleiotropy by identifying and removing outlier SNPs that could bias the causal effect, using the default settings.

Causal Analysis Using Summary Effect estimates⁴⁰

The Causal Analysis Using Summary Effect estimates (CAUSE) is a Bayesian MR method which harnesses the full genome-wide set of variant summary statistics (as opposed to only genome-wide significant SNPs in a traditional MR) to distinguish the causal effect from correlated pleiotropy (when a variant affects the exposure and outcome through a shared heritable factor) and uncorrelated horizontal pleiotropy (when a variant affects the exposure and outcome through separate mechanisms). We used it to help discern if the effect of OA on spinal stenosis seen in standard MR analysis was driven more by shared genetic heritability of the two traits or causal effect.

**Fig. 1**

Flowchart providing overview of datasets and methods used in the current MR study.

Reverse MR

We also carried out reverse MR, i.e. we used the FinnGen spinal stenosis GWAS (Table 1) as the exposure to detect any potential causal effect of genetic liability for spinal stenosis on any of the tested risk factors.

Heterogeneity

We used the standard statistics of Cochran's Q and I^2 to assess heterogeneity in our MR IVW analyses.

MVMR analyses

In order to test if the effect of risk factors with significant effect on spinal stenosis, as identified in the univariable analysis, is mediated by OA, a MVMR model was used combining both exposure variables in a single regression test and meta-analysed using IVW method. We also carried out MVMR analyses adjusting simultaneously for BMI and body fat distribution traits,³⁰ as well as BMI and BMD⁴¹ as these are strongly positively genetically correlated.

The instrument strength (conditional F-statistics, F_{TS}) and effect heterogeneity (Cochran's Q_A) in MVMR context were calculated using the MVMR⁴² package with the covariance between genetic associations with each exposure fixed at zero in the primary analysis, but a range of values was also tested. Since we detected presence of weak instrument bias towards the (likely) confounded observational association, the Q-minimisation approach (Q_{HET}) from the MVMR package⁴² was run as a sensitivity analysis to complement the MVMR-IVW results.

Results

Investigation of total effect of risk factors on spinal stenosis using two sample MR

Our power analysis showed that we had at least 80% power for detecting small-to-moderate effects (odds ratio: 1.07–1.27) for a range of anthropometric risk factors using the FinnGen spinal stenosis GWAS (Supplementary Table II). Unless otherwise stated, all the main results presented are derived using the IVW estimator and

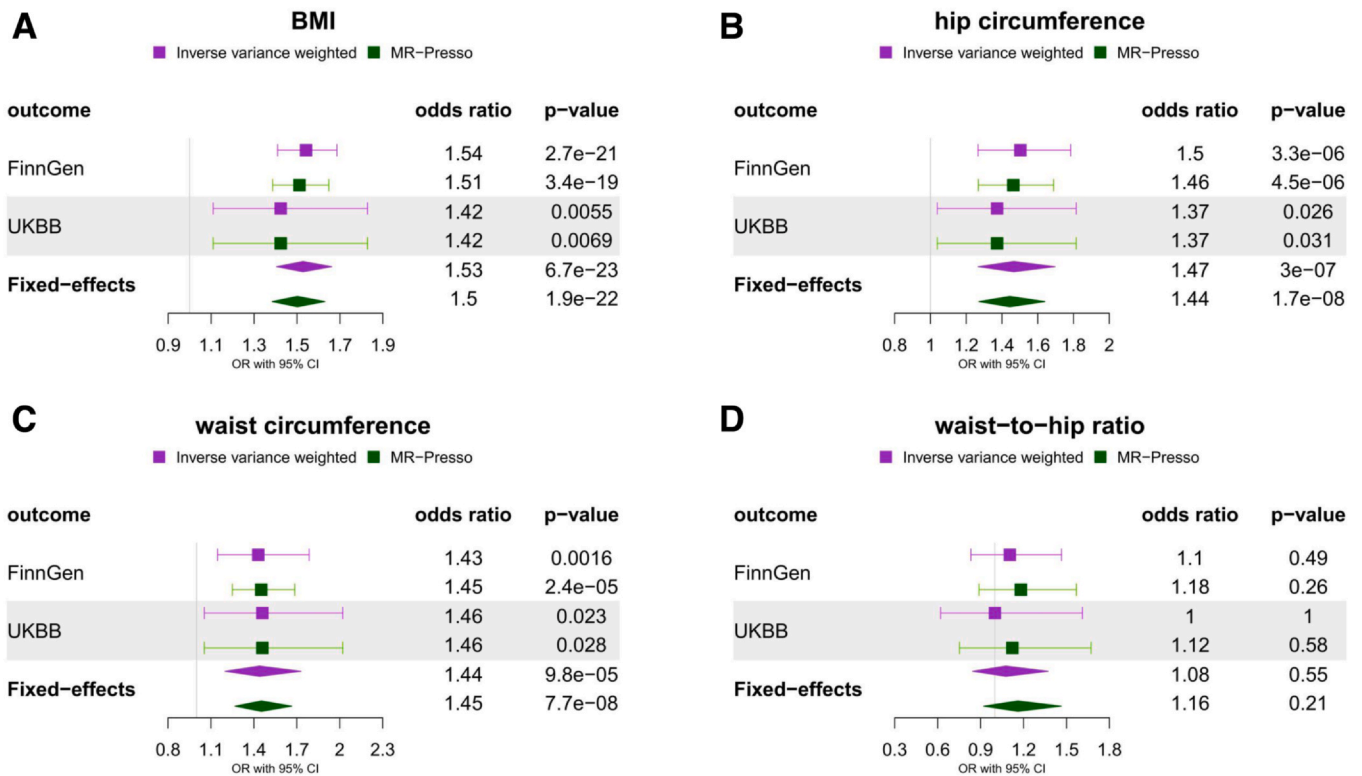


Fig. 2

Osteoarthritis and Cartilage

Two sample Mendelian randomisation results for the effect of genetic susceptibility for adiposity traits (A – BMI, B – hip circumference, C – waist circumference, D – waist-to-hip ratio) on spinal stenosis (FinnGen and UK BioBank). Plots compare results obtained using IVW and outlier-robust MR-PRESSO method and display fixed-effects meta-analysis results of the odds ratio per SD increase in exposure obtained using FinnGen and UK Biobank outcomes.

FinnGen spinal stenosis outcome. Among the adiposity traits, we found strong evidence for the effect of higher BMI (OR = 1.54, 95% CI: 1.41–1.69, p -value = 2.7×10^{-21} , Fig. 2A), hip circumference (OR = 1.50, 95% CI: 1.27–1.78, p -value = 3.3×10^{-6} , Fig. 2B) and waist-circumference (OR = 1.43, 95% CI: 1.15–1.79, p -value = 1.5×10^{-3} , Fig. 2C) but not waist-to-hip ratio (OR = 1.10, 95% CI: 0.83–1.47, p -value = 0.49, Fig. 2D) on spinal stenosis. Among the skeletal traits, we found weak evidence for a causal effect of increased height (FinnGen: OR = 1.06, 95% CI: 0.99–1.14, p -value = 0.10; UKBB: OR = 1.15, 95% CI: 1.04–1.28, p -value = 6.6×10^{-3} ; Fig. 3A) but strong evidence for a causal effect of higher total BMD (OR = 1.21, 95% CI: 1.12–1.29, p -value = 1.6×10^{-7} , Fig. 3B), femoral neck BMD (OR = 1.22, 95% CI: 1.09–1.37, p -value = 5.9×10^{-4} , Fig. 3C) and lumbar spine BMD (OR = 1.38, 95% CI: 1.25–1.52, p -value = 4.4×10^{-11} , Fig. 3D) on spinal stenosis. On the other hand, little evidence of an effect was found for circulating calcium (OR = 1.02, 95% CI: 0.93–1.11, p -value = 0.69, Supplementary Fig. 1A) and phosphate (OR = 0.94, 95% CI: 0.85–1.03, p -value = 0.19, Supplementary Fig. 1B).

Sensitivity analyses – two sample MR

Fixed-effects meta-analysis of MR results based on both FinnGen and UK Biobank spinal stenosis outcome GWAS resulted in similar estimates to those obtained using solely FinnGen. SNP outliers apparent in the scatter plots (Supplementary Figs. 2–3) along with

significant Cochran's Q values for all exposures (except for lumbar spine BMD, Supplementary Table III) suggested presence of effect heterogeneity. However, outlier-robust sensitivity method MR-PRESSO reproduced the same magnitude of associations, while other methods (MR Egger, weighted median and mode) were consistent with IVW/MR-PRESSO estimates overall (Supplementary Table IV). Non-significant Egger's intercept (Supplementary Table V) suggested limited presence of horizontal pleiotropy. Reverse MR analysis with the FinnGen spinal stenosis found little evidence of effect on all risk factor traits, apart from lumbar spine BMD (OR = 1.15, 95% CI: 1.06–1.26, p -value = 1.5×10^{-3}).

Shared genetic liability for spinal stenosis and OA

We then investigated the magnitude of LD score-derived genetic correlation between spinal stenosis and OA across various sites (Fig. 4). As expected, the highest correlation was found between the two spinal stenosis GWAS (r_g = 0.77, p -value = 1.1×10^{-23}), however high genetic correlation was also revealed between spine OA ($r_{gFinnGen}$ = 0.66, p -value = 4×10^{-22} ; r_{gUKBB} = 0.73, p -value = 1.3×10^{-11}) and spinal stenosis. Genetic correlation across other OA sites was high in the FinnGen spinal stenosis GWAS (from r_g = 0.47, p -value = 2.5×10^{-23} and p -value = 2.8×10^{-17} for knee and hip OA, respectively, to r_g = 0.52, p -value = 2.4×10^{-25} for all OA) and moderate in the UKBB spinal stenosis GWAS (r_g ranging 0.3–0.38).

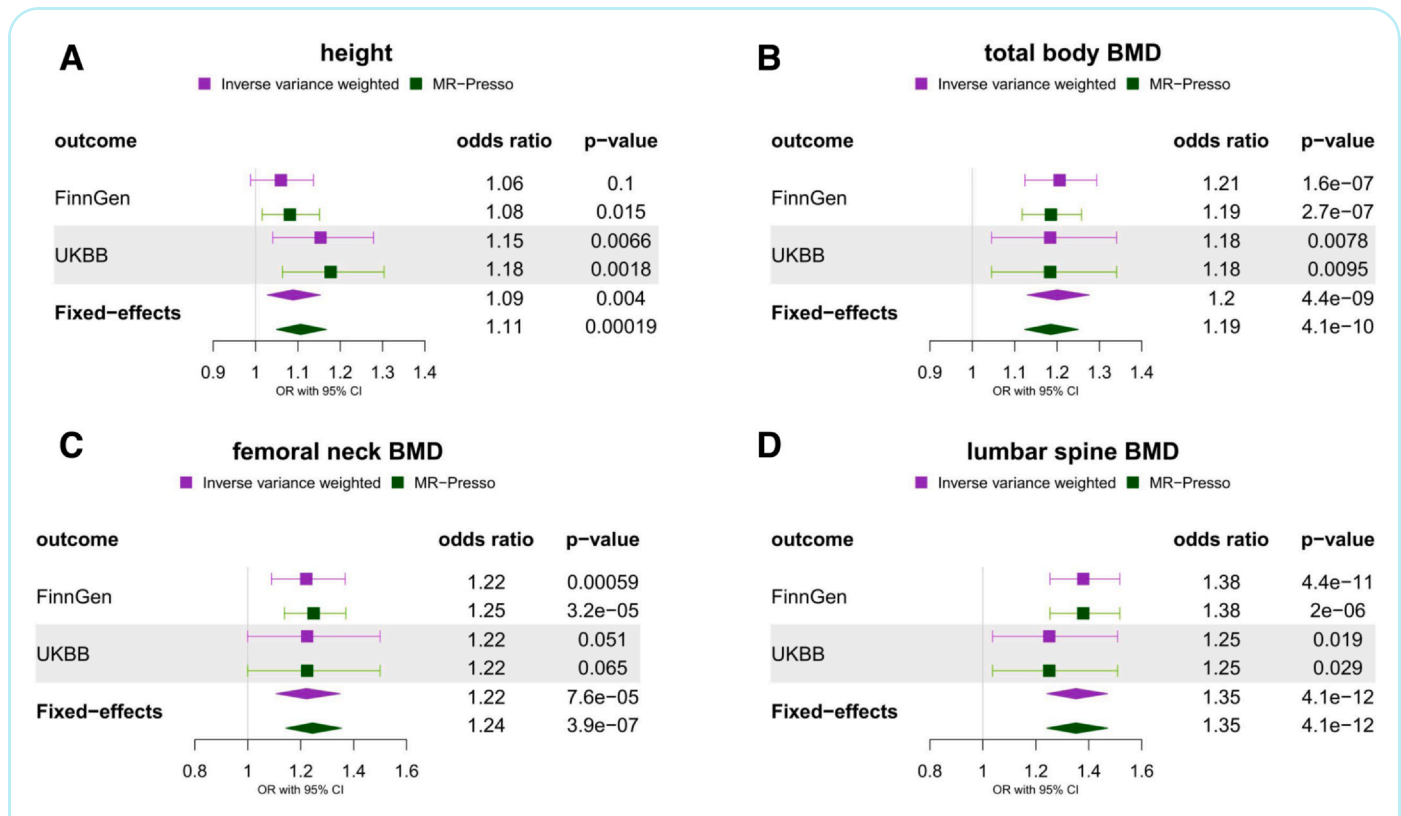


Fig. 3

Osteoarthritis and Cartilage

Two sample Mendelian randomisation results for the effect of genetic susceptibility for skeletal traits (A – height, B – total body BMD, C – femoral neck BMD, D – lumbar spine BMD) on spinal stenosis (FinnGen and UK Biobank). Plots compare results obtained using IVW and outlier-robust MR-PRESSO method and display fixed-effects meta-analysis results of the odds ratio per SD increase in exposure obtained using FinnGen and UK Biobank outcomes.

Shared risk factors for OA and spinal stenosis

Given the substantial genetic correlation of spinal stenosis with OA and evidence of causal effects of adiposity traits as well as BMD on OA in previous MR studies,^{20,21,43,44} we hypothesised that OA may be a major mediator of the effects of these risk factors on spinal stenosis. First, to investigate this hypothesis in the two-step MR framework,⁴⁵ we replicated the evidence for causal effects of anthropometric risk factors on OA (Supplementary Tables VI–VIII): BMI, hip circumference, waist circumference (Supplementary Fig. 4), height, total BMD, femoral neck BMD and lumbar spine BMD (Supplementary Fig. 5). As the next step, we assessed the bidirectional relationship between OA and spinal stenosis (Supplementary Figs. 6–7, Supplementary Tables IV, VI). The main IVW result confirmed the causal effect of all site OA on spinal stenosis with OR of 1.44 (95% CI: 1.13–1.84, p-value = 3.1×10^{-3}). IVW result for knee OA (OR = 1.16, 95% CI: 0.98–1.38, p-value = 0.09) was markedly increased after outlier correction using MR-PRESSO (OR = 1.34, 95% CI: 1.18–1.52, p-value = 1.3×10^{-4}) and a very uncertain estimate was available for spine OA (OR = 1.13, 95% CI: 0.75–1.71, p-value = 0.56) as calculated using a single instrument (F-statistic = 30).

Spinal stenosis is causally downstream of OA

To help overcome this power limitation and establish the true causal path between OA and spinal stenosis given their shared genetic heritability, we applied the Bayesian CAUSE method (Supplementary Table

IX). When evaluating the bidirectional relationship between OA and FinnGen spinal stenosis GWAS, the causal model was always picked over the sharing model (p-value from 1.4×10^{-6} to 4.9×10^{-3}). In each case, effect size for the OA to spinal stenosis direction dominated (OR_{all OA} = 1.6, 95% CI: 1.41–1.79; OR_{spinal OA} = 1.4, 95% CI: 1.21–1.62) over the reverse direction (OR_{all OA} = 1.07, 95% CI: 1.05–1.09; OR_{spinal OA} = 1.13, 95% CI: 1.09–1.17).

Investigation of direct effect of risk factors on spinal stenosis independent of OA using MVMR

In light of the predicted strong causal effect of OA on spinal stenosis and both OA and spinal stenosis sharing the same set of anthropometric risk factors in our two-sample MR analyses, we employed MVMR to estimate the direct effect of a given risk factor on spinal stenosis accounting for OA (Supplementary Table X). The direct effect of higher BMI on spinal stenosis (Fig. 5A) ranged from OR = 1.29 for all OA mediator (95% CI: 1.16–1.45, p-value = 7.2×10^{-6}) to OR = 1.37 for spine OA mediator (95% CI: 1.24–1.51, p-value = 4.7×10^{-10}) which corresponded to all OA mediating 16.2% (95% CI: 14.2%–17.8%) of the total effect of BMI on spinal stenosis. For height, adjusting for OA resulted in the direct effect being consistent with the null hypothesis (Fig. 5B) for all OA (OR = 1.01, 95% CI: 0.94–1.08, p-value = 0.79) and spine OA (OR = 1.01, 95% CI: 0.94–1.08, p-value = 0.85), albeit a weak direct effect remained in the UKBB analysis. Next, total body BMD direct effect adjusted for OA (Fig. 5C) resulted in OR = 1.19 (95% CI: 1.11–1.29, p-value = 6.6×10^{-6}) for all OA and in

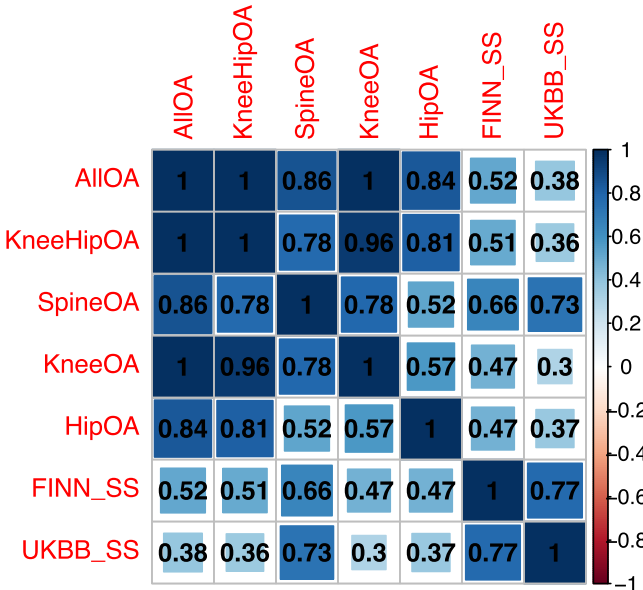


Fig. 4 Osteoarthritis and Cartilage

Genetic correlation of osteoarthritis and spinal stenosis phenotypes estimated by LD score regression. Correlation coefficients are displayed within cells and the colour/area of the cells are proportionally scaled. All p-values are significant after FDR correction. FDR, false discovery rate, FINN_SS, FinnGen spinal stenosis, UKBB_SS, UK Biobank spinal stenosis.

OR = 1.2 (95% CI: 1.11–1.29, p-value = 1×10^{-6}) for spine OA. Interestingly, unadjusted odds-ratio for total body BMD did not meaningfully differ (OR = 1.21, 95% CI: 1.12–1.29, p-value = 1.6×10^{-7}) suggesting total body BMD affects OA through an independent pathway. This was not unlike femoral neck BMD (Fig. 5D), where the direct effect accounting for all OA (OR = 1.19, 95% CI: 1.06–1.33, p-value = 3.2×10^{-3}) and spine OA (OR = 1.15, 95% CI: 1.01–1.30, p-value = 0.03) equated to all OA mediating 2.5% of the total effect of femoral neck BMD on spinal stenosis. Similarly, relatively low (5.8%) degree of mediation was found for the lumbar spine BMD outcome (Supplementary Fig. 8). Summary of the main findings from OA mediation analysis is provided in Fig. 6.

Direct effect of waist/hip circumference on spinal stenosis independent of BMI

Since the two non-BMI adiposity risk factors which we identified (waist and hip circumference) are phenotypically and genetically correlated with BMI, we used the MVMR approach to arrive at direct estimates adjusted for BMI (Supplementary Fig. 9). We found that the corrected estimates shifted towards the null for both waist (OR = 1.13, 95% CI = 0.82–1.55, p-value = 0.45) and hip circumference (OR = 1.12, 95% CI = 0.85–1.46, p-value = 0.42).

Direct effect of BMD on spinal stenosis independent of both OA and BMI

Lastly, since previous research hypothesised that BMI can be a confounder of a relationship between BMD and OA,²¹ we were interested in studying the mutually adjusted effect of the three variables on spinal stenosis (Supplementary Fig. 10, Supplementary Table XI). In the model including total body BMD and all OA exposures, the estimated effect of BMI on spinal stenosis remained consistent (OR = 1.32, 95% CI: 1.16–1.47, p-value = 3.2×10^{-6}) with the

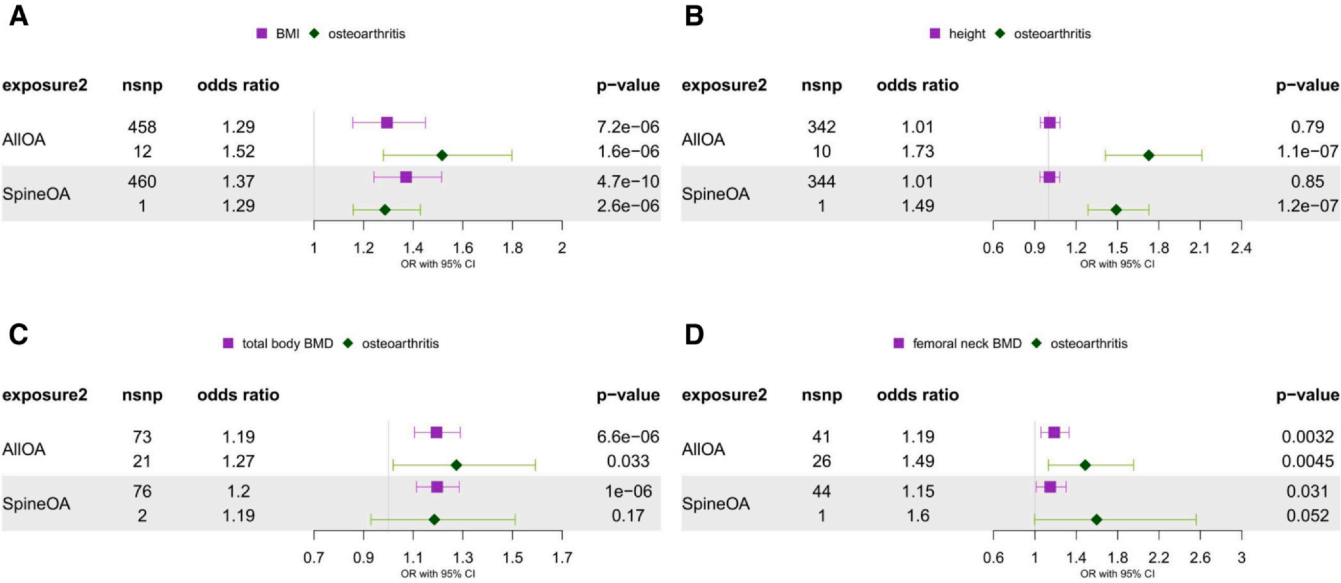
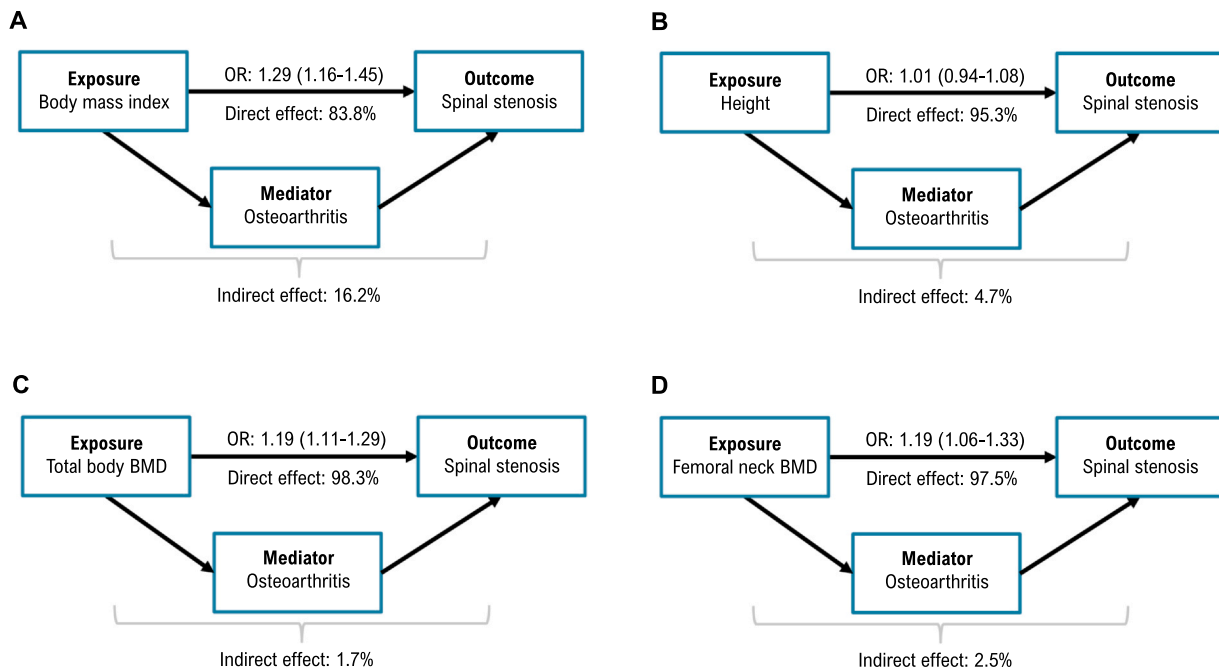


Fig. 5 Osteoarthritis and Cartilage

Multivariable Mendelian randomisation results for the jointly modelled effect of genetic susceptibility for risk factors (A – BMI, B – height, C – total body BMD, D – femoral neck BMD) and liability for osteoarthritis (all or spine) on spinal stenosis (FinnGen). The odds ratios are scaled per SD increase of risk factors and doubling in the odds of osteoarthritis.

**Fig. 6**

Osteoarthritis and Cartilage

Summary diagrams illustrating direct and indirect, osteoarthritis (all sites)-mediated effects of risk factors (A – BMI, B – height, C – total body BMD, D – femoral neck BMD) on spinal stenosis (FinnGen). OR indicates odds ratio (with 95% CI).

model including only OA covariate, while the total BMD estimate was slightly attenuated (OR = 1.13, 95% CI: 1.03–1.24, p -value = 8.3×10^{-3}) but there is a large amount of uncertainty in the estimate. The results for the model involving spine OA rather than all OA, and femoral neck BMD rather than total body BMD were analogous. There is some evidence that BMI is common cause of both lumbar spine BMD and OA shown by the significant reduction in the effect of lumbar spine BMD on spinal stenosis (adjusted for both BMI and all OA: OR = 1.01, 95% CI: 0.90–1.14, p -value = 0.83; adjusted for all OA only: OR = 1.30, 95% CI: 1.15–1.47, p -value = 4.4×10^{-5}).

Sensitivity analyses – MVMR

As we detected presence of potential pleiotropy due to high heterogeneity as measured by Cochran's Q_A and weak instrument bias evidenced by conditional F-statistics < 10 in our MVMR analyses (Supplementary Tables XII–XIII), we applied the robust estimator Q_{HET} in a sensitivity analyses. The method produced results generally consistent with the IVW MVMR results, albeit with a much higher degree of uncertainty around the true causal value (Supplementary Tables XIV–XV).

Discussion

Our understanding of spinal stenosis epidemiology remains quite limited despite the condition's relatively high prevalence among older adults and its association with substantial pain and mobility impairment. In this study, we applied a genetic epidemiology method (MR) to investigate the causal relationships between anthropometric risk factors, OA and spinal stenosis.

When analysed independently, BMI was found to act as a strong risk factor for spinal stenosis (meta-analysed OR = 1.53 per 1 standard deviation (SD) increase in exposure), similar to hip

circumference (OR = 1.47) and waist circumference (OR = 1.44) but these attenuated to the null after adjusting for BMI in multivariable analysis. BMD across different sites also showed a substantial effect on spinal stenosis: total (OR = 1.2), hip (femoral neck, OR = 1.22) and lumbar spine (OR = 1.35). As lumbar spine BMD measurement is liable to falsely increase with degenerative change^{46,47} and spinal stenosis liability affects lumbar spine BMD in our reverse MR analysis, we subsequently focussed on total and hip BMD. Interestingly, in a previous case-control study higher BMD was found in lumbar spinal stenosis cases across not only the lumbar spine, but also femoral neck and total hip.⁴⁸ In addition, we found that circulating calcium and phosphate exhibited little to no evidence for an effect on spinal stenosis.

OA, in particular facet joint OA of the spine, can contribute to the narrowing of the spinal canal thanks to joint hypertrophy and formation of synovial cysts.¹¹ In agreement with this biological mechanism, our MR analysis found a positive effect of a genetic predisposition to OA (when measured at all sites) on the development of spinal stenosis. These results were further supported by the Bayesian CAUSE model which found our results were more likely to be driven by a causal effect of a genetic predisposition to OA than by correlated and horizontal pleiotropy. We also identified a reverse causal effect, hypothesised to be indicative mostly of a shared genetic aetiology, as supported by LD score regression estimating inter-trait genetic correlation. It is worth noting that while our spinal OA signals showed consistent results, in terms of direction of effect, the estimates were less precise likely due to the reduced number of genetic instruments as compared with OA at all sites.

MVMR, which models the joint effects of multiple risk factors on an outcome to assess their individual contributions, identified a largely OA-independent causal pathway between BMI, BMD and spinal stenosis, with OA mediating < 20% of the effect of BMI and < 6% of BMD. However, weak evidence for the causal effect of height on spinal stenosis (OR = 1.09) was diminished to the null in the

MVMR analyses suggesting that the univariable effect was driven by the causal association with OA. Moreover, we did not find compelling evidence for BMI to be acting as a confounder for the association of BMD, OA and spinal stenosis.

MR can only provide reliable causal estimates subject to meeting three key assumptions which were tested in multiple ways in our analysis. The first criterion (“relevance”), that the genetic variants are robustly associated with the risk factor of interest was met by using variants with genome-wide significant associations with exposure and using variants with F-statistics > 30 that should minimise weak instrument bias, which can arise when the genetic variant explains only a small proportion of the variance in the risk factor. Weak instrument bias can move the MR estimate towards the observational confounded association and increase type 1 error rate. The second criterion, that the genetic variant shares no unmeasured confounder with the outcome (‘independence’/‘exchangeability’) is usually concerned with confounding by population stratification which is addressed during the initial GWAS analysis. In addition, bidirectional MR analysis confirmed that associations between risk factors and spinal stenosis were not confounded by reverse causation in all but one case.

Perhaps the most pervasive problem plaguing MR analysis is the violation of the third assumption, that the genetic variant affects the outcome only through its association with the risk factor, and not through any other independent pathways (‘exclusion restriction’, i.e. no horizontal pleiotropy). We evaluated this assumption with the MR Egger intercept test and MR-PRESSO analysis. Also included were a range of MR sensitivity methods (MR-Egger, weighted median, weighted mode) whose results are consistent in magnitude with the main IVW results and so indicate that the independence and exclusion restriction assumptions were not violated.

Our IVW MVMR analysis typically suffered from low strength of the genetic instrument for 1–2 exposures. We tried to rectify that by applying the Q-minimisation approach which is more robust to these violations of MR assumptions but there remains a possibility that our MVMR direct estimates are incompletely adjusted.

Since there was no gold standard diagnostic tool for spinal stenosis at the time of data collection with diagnosis based on clinical history, physical examination, and imaging,^{49,50} varying case definition will introduce an additional layer of heterogeneity into GWAS and reduce its power. Using a severe end of the phenotype spectrum can lead to reduced power in GWAS, and so fewer genomewide-significant hits. This is demonstrated by 0 versus 21 genome-wide significant loci in the UKBB (3713 cases) and FinnGen (16,698 cases) spinal stenosis GWAS, respectively. Likewise, the OA outcomes from the GO consortium included a range of definitions, including hospital diagnosis, radiographic evidence and self-reporting, which can inflate estimate heterogeneity, and so increase the risk of a weak instrument bias. Furthermore, this MR study could benefit from inclusion of more ancestrally diverse populations to compare the estimated effects of identified risk factors but currently no suitable spinal stenosis outcome GWAS in non-Europeans is available.

Our study has public health implications, as efforts to minimise prevalence of high adiposity in the population should lead to reduction in spinal stenosis incidence and associated benefits regarding quality of life and healthcare costs. Previously identified obese individuals with elevated BMD measurement could be especially targeted for weight loss intervention due to higher compounded risk of spinal stenosis. Moreover, while the current MR study uses condition prevalence as the outcome, it is quite likely that the risk factors identified could contribute to progression of symptoms.

In conclusion, we examined a variety of potential anthropometric risk factors for spinal stenosis, both independently and in conjunction with potential mediators. Our findings, confirmed by two-sample IVW MR, MR-PRESSO, and CAUSE analyses, demonstrate that a genetic predisposition to OA causally contributes to the development of spinal

stenosis. Overall, we have found evidence for OA-independent causal effect of BMI on spinal stenosis, in addition to BMI- and OA-independent causal effect of BMD. Further investigation is necessary to elucidate the mechanisms through which elevated BMD and BMI contribute to spinal stenosis, as well as to explore the functional genomics of spinal stenosis, including potential drug targets.

Ethics approval

UK Biobank received ethical approval from the Research Ethics Committee (REC reference: 11/NW/0382). As this study did not involve human subjects or individual-level data, no ethics approval was required.

Role of the funding source

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Contributions

MKS: Conception and design, Collection and assembly of data, Statistical expertise, Analysis and interpretation of the data, Drafting of the article, Final approval of the article. BGF, AH: Analysis and interpretation of the data, Critical revision of the article for important intellectual content, Final approval of the article. MF, GO consortium, LS, EZ, HT: Collection and assembly of data, Analysis and interpretation of the data, Final approval of the article. TRG: Obtaining of funding, Critical revision of the article for important intellectual content, Final approval of the article. MKS takes responsibility for the integrity of the work as a whole, from inception to finished article.

Competing interests

TRG receives funding from Biogen for unrelated research.

Data availability

GWAS summary statistics availability is provided by relevant publications. Custom summary statistics were obtained from the GO Consortium on application.

Appendix A. Supporting information

Supplementary data associated with this article can be found in the online version at [doi:10.1016/j.joca.2023.12.003](https://doi.org/10.1016/j.joca.2023.12.003).

References

1. Kalichman L, Cole R, Kim DH, et al. Spinal stenosis prevalence and association with symptoms: the Framingham Study. *Spine J* 2009;9(7):545–50. <https://doi.org/10.1016/j.spinee.2009.03.005>
2. Ishimoto Y, Yoshimura N, Muraki S, et al. Prevalence of symptomatic lumbar spinal stenosis and its association with physical performance in a population-based cohort in Japan: the Wakayama Spine Study. *Osteoarthritis Cartil* 2012;20(10):1103–8. <https://doi.org/10.1016/j.joca.2012.06.018>
3. Jensen RK, Jensen TS, Koes B, Hartvigsen J. Prevalence of lumbar spinal stenosis in general and clinical populations: a systematic review and meta-analysis. *Eur Spine J* 2020;29(9):2143–63. <https://doi.org/10.1007/s00586-020-06339-1>
4. Binder DK, Schmidt MH, Weinstein PR. Lumbar spinal stenosis. *Semin Neurol* 2002;22(02):157–66. <https://doi.org/10.1055/s-2002-36539>
5. Otani K, Kikuchi S, Yabuki S, et al. Lumbar spinal stenosis has a negative impact on quality of life compared with other comorbidities: an epidemiological cross-sectional study of 1862 community-dwelling individuals. *Sci World J* 2013;2013(590652):1–9.
6. Melancia JL, Francisco AF, Antunes JL. Chapter 35 - Spinal stenosis. In: Biller J, Ferro JM, editors. *Neurologic Aspects of Systemic Disease Part I*, Vol 119. Amsterdam, The Netherlands: Elsevier; 2014. p. 541–9. <https://doi.org/10.1016/B978-0-7020-4086-3.00035-7>
7. Ciol MA, Deyo RA, Howell E, Kreif S. An assessment of surgery for spinal stenosis: time trends, geographic variations, complications, and reoperations. *J Am Geriatr Soc* 1996;44(3):285–90. <https://doi.org/10.1111/j.1532-5415.1996.tb00915.x>
8. Deyo RA, Mirza SK, Martin BI, Kreuter W, Goodman DC, Jarvik JG. Trends, major medical complications, and charges associated with surgery for lumbar spinal stenosis in older adults. *JAMA* 2010;303(13):1259–65. <https://doi.org/10.1001/jama.2010.338>
9. Katz JN, Zimmerman ZE, Mass H, Makhni MC. Diagnosis and management of lumbar spinal stenosis: a review. *JAMA* 2022;327(17):1688–99. <https://doi.org/10.1001/jama.2022.5921>
10. Szpalski M, Gunzburg R. Lumbar spinal stenosis in the elderly: an overview. *Eur Spine J* 2003;12(2):S170–5. <https://doi.org/10.1007/s00586-003-0612-1>
11. Byvaltsev VA, Kalinin AA, Hernandez PA, et al. Molecular and genetic mechanisms of spinal stenosis formation: systematic review. *Int J Mol Sci* 2022;23(21):1–17. <https://doi.org/10.3390/ijms232113479>
12. Knutsson B, Sandén B, Sjöden G, Järvholm B, Michaëlsson K. Body mass index and risk for clinical lumbar spinal stenosis: a cohort study. *Spine* 2015;40(18):1451–6. https://journals.lww.com/spinejournal/Fulltext/2015/09150/Body_Mass_Index_and_Risk_for_Clinical_Lumbar.12.aspx
13. Liuke M, Solovieva S, Lamminen A, et al. Disc degeneration of the lumbar spine in relation to overweight. *Int J Obes* 2005;29(8):903–8. <https://doi.org/10.1038/sj.ijo.0802974>
14. Kalichman L, Guermazi A, Li L, Hunter DJ. Association between age, sex, BMI and CT-evaluated spinal degeneration features. *J Back Musculoskelet Rehabil* 2009;22:189–95. <https://doi.org/10.3233/BMR-2009-0232>
15. Ou C-Y, Lee T-C, Lee T-H, Huang Y-H. Impact of body mass index on adjacent segment disease after lumbar fusion for degenerative spine disease. *Neurosurgery* 2015;76(4):396–402. https://journals.lww.com/neurosurgery/Fulltext/2015/04000/Impact_of_Body_Mass_Index_on_Adjacent_Segment.16.aspx
16. Felson DT, Zhang Y, Hannan MT, Anderson JJ. Effects of weight and body mass index on bone mineral density in men and women: The framingham study. *J Bone Miner Res* 1993;8(5):567–73. <https://doi.org/10.1002/jbmr.5650080507>
17. Qiao D, Li Y, Liu X, et al. Association of obesity with bone mineral density and osteoporosis in adults: a systematic review and meta-analysis. *Public Health* 2020;180:22–8. <https://doi.org/10.1016/j.puhe.2019.11.001>
18. Hardcastle SA, Dieppe P, Gregson CL, Davey Smith G, Tobias JH. Osteoarthritis and bone mineral density: are strong bones bad for joints? *Bonekey Rep* 2015;4:624. <https://doi.org/10.1038/bonekey.2014.119>
19. Davies NM, Holmes MV, Davey Smith G. Reading Mendelian randomisation studies: A guide, glossary, and checklist for clinicians. *BMJ* 2018;362:1–11. <https://doi.org/10.1136/bmj.k601>
20. Funck-Brentano T, Nethander M, Movérare-Skrtic S, Richette P, Ohlsson C. Causal factors for knee, hip, and hand osteoarthritis: a mendelian randomization study in the UK Biobank. *Arthritis Rheumatol* 2019;71(10):1634–41. <https://doi.org/10.1002/art.40928>
21. Hartley A, Sanderson E, Granel R, et al. Using multivariable Mendelian randomization to estimate the causal effect of bone mineral density on osteoarthritis risk, independently of body mass index. *Int J Epidemiol* 2022;51(4):1254–67. <https://doi.org/10.1093/ije/dyab251>
22. Berikol G, Ekşi MŞ, Aydın L, Börekci A, Özcan-Ekşi EE. Subcutaneous fat index: a reliable tool for lumbar spine studies. *Eur Radiol* 2022;32(9):6504–13. <https://doi.org/10.1007/s00330-022-08775-7>
23. Heuch I, Heuch I, Hagen K, Zwart J-A. Association between body height and chronic low back pain: a follow-up in the Nord-Trøndelag Health Study. *BMJ Open* 2015;5(6):1–6. <https://doi.org/10.1136/bmjopen-2014-006983>
24. Ciosek Z, Kot K, Kosik-Bogacka D, Łanocha-Arendarczyk N, Rotter I. The effects of calcium, magnesium, phosphorus, fluoride, and lead on bone tissue. *Biomolecules* 2021;11(4):1–26. <https://doi.org/10.3390/biom11040506>
25. Aljuraibah F, Bacchetta J, Brandi ML, et al. An expert perspective on phosphate dysregulation with a focus on chronic hypophosphatemia. *J Bone Miner Res* 2022;37(1):12–20. <https://doi.org/10.1002/jbmr.4486>
26. Kurki MI, Karjalainen J, Palta P, et al. FinnGen provides genetic insights from a well-phenotyped isolated population. *Nature* 2023;613(7944):508–18. <https://doi.org/10.1038/s41586-022-05473-8>
27. Gagliano Taliun SA, VandeHaar P, Boughton AP, et al. Exploring and visualizing large-scale genetic associations by using PheWeb. *Nat Genet* 2020;52(6):550–2. <https://doi.org/10.1038/s41588-020-0622-5>
28. Locke AE, Kahali B, Berndt SI, et al. Genetic studies of body mass index yield new insights for obesity biology. *Nature* 2015;518(7538):197–206. <https://doi.org/10.1038/nature14177>
29. Yengo L, Sidorenko J, Kempner KE, et al. Meta-analysis of genome-wide association studies for height and body mass index in ~700,000 individuals of European ancestry. *Hum Mol Genet* 2018;27(20):3641–9. <https://doi.org/10.1093/hmg/ddy271>
30. Shungin D, Winkler T, Croteau-Chonka DC, et al. New genetic loci link adipose and insulin biology to body fat distribution. *Nature* 2015;518(7538):187–96. <https://doi.org/10.1038/nature14132>
31. Wood AR, Esko T, Yang J, et al. Defining the role of common variation in the genomic and biological architecture of adult human height. *Nat Genet* 2014;46(11):1173–86. <https://doi.org/10.1038/ng.3097>
32. Morris JA, Kemp JP, Youten SE, et al. An atlas of genetic influences on osteoporosis in humans and mice. *Nat Genet* 2019;51(2):258–66. <https://doi.org/10.1038/s41588-018-0302-X>
33. Zheng H-F, Forgetta V, Hsu Y-H, et al. Whole-genome sequencing identifies EN1 as a determinant of bone density and fracture. *Nature* 2015;526(7571):112–7. <https://doi.org/10.1038/nature14878>

34. Hemani G, Zheng J, Elsworth B, *et al.* The MR-Base platform supports systematic causal inference across the human phenome. *Elife* 2018;7, e34408. <https://doi.org/10.7554/eLife.34408>
35. Elsworth B, Lyon M, Alexander T, *et al.* The MRC IEU OpenGWAS data infrastructure. (Published online). *bioRxiv*2020:1–22. <https://doi.org/10.1101/2020.08.10.244293>. (Published online).
36. Bulik-Sullivan B, Finucane HK, Anttila V, *et al.* An atlas of genetic correlations across human diseases and traits. *Nat Genet* 2015;47(11):1236–41. <https://doi.org/10.1038/ng.3406>
37. Auton A, Abecasis GR, Altshuler DM, *et al.* A global reference for human genetic variation. *Nature* 2015;526(7571):68–74. <https://doi.org/10.1038/nature15393>
38. Zheng X, Levine D, Shen J, Gogarten SM, Laurie C, Weir BS. A high-performance computing toolset for relatedness and principal component analysis of SNP data. *Bioinformatics* 2012;28(24):3326–8. <https://doi.org/10.1093/bioinformatics/bts606>
39. Verbanck M, Chen C-Y, Neale B, Do R. Detection of widespread horizontal pleiotropy in causal relationships inferred from Mendelian randomization between complex traits and diseases. *Nat Genet* 2018;50(5):693–8. <https://doi.org/10.1038/s41588-018-0099-7>
40. Morrison J, Knoblauch N, Marcus JH, Stephens M, He X. Mendelian randomization accounting for correlated and uncorrelated pleiotropic effects using genome-wide summary statistics. *Nat Genet* 2020;52(7):740–7. <https://doi.org/10.1038/s41588-020-0631-4>
41. Kemp JP, Sayers A, Smith GD, Tobias JH, Evans DM. Using Mendelian randomization to investigate a possible causal relationship between adiposity and increased bone mineral density at different skeletal sites in children. *Int J Epidemiol* 2016;45(5):1560–72. <https://doi.org/10.1093/ije/dyw079>
42. Sanderson E, Spiller W, Bowden J. Testing and correcting for weak and pleiotropic instruments in two-sample multivariable Mendelian randomization. *Stat Med* 2021;40(25):5434–52. <https://doi.org/10.1002/sim.9133>
43. Lyu L, Cai Y, Xiao M, *et al.* Causal relationships of general and abdominal adiposity on osteoarthritis: a two-sample mendelian randomization study. *J Clin Med* 2022;12(1):1–13. <https://doi.org/10.3390/JCM12010320>
44. Ho J, Mak CCH, Sharma V, To K, Khan W. Mendelian randomization studies of lifestyle-related risk factors for osteoarthritis: a PRISMA review and meta-analysis. *Int J Mol Sci* 2022;23(19):1–24. <https://doi.org/10.3390/IJMS231911906>
45. Evans DM, Davey Smith G. Mendelian randomization: new applications in the coming age of hypothesis-free causality. *Annu Rev Genomics Hum Genet* 2015;16(1):327–50. <https://doi.org/10.1146/annurev-genom-090314-050016>
46. Dalle Carbonare L, Giannini S, Sartori L, *et al.* Lumbar osteoarthritis, bone mineral density, and quantitative ultrasound. *Aging* 2000;12(5):360–5. <https://doi.org/10.1007/BF03339861>
47. Tenne M, McGuigan F, Besjakov J, Gerdhem P, Åkesson K. Degenerative changes at the lumbar spine—implications for bone mineral density measurement in elderly women. *Osteoporos Int* 2013;24(4):1419–28. <https://doi.org/10.1007/s00198-012-2048-0>
48. Kim H-J, Lee H-M, Kim H-S, *et al.* Bone metabolism in postmenopausal women with lumbar spinal stenosis: analysis of bone mineral density and bone turnover markers. *Spine* 2008;33(22):2435–9. https://journals.lww.com/spinejournal/Fulltext/2008/10150/Bone_Metabolism_in_Postmenopausal_Women_With.12.aspx
49. Wu A-M, Zou F, Cao Y, *et al.* Lumbar spinal stenosis: an update on the epidemiology, diagnosis and treatment. *AME Med J* 2017;2(63):1–14. (Published online).
50. Tomkins-Lane C, Melloh M, Wong A. Diagnostic tests in the clinical diagnosis of lumbar spinal stenosis: Consensus and Results of an International Delphi Study. *Eur Spine* 2020;29(9):2188–97. <https://doi.org/10.1007/s00586-020-06481-w>